

# *Time Series Analysis*

The historian is a prophet in reverse.  
— Friedrich von Schlegel (1772–1829)

Even in my tape reading something enters  
that is more than mere arithmetic.  
— Edwin Lefèvre (1971–1943),  
*Reminiscences of a Stock Operator* (1923)

## GARCH Option Pricing

- Options can be priced when the underlying asset's return follows a GARCH (generalized autoregressive conditional heteroskedastic) process.<sup>a</sup>
- Let  $S_t$  denote the asset price at date  $t$ .
- Let  $h_t^2$  be the *conditional* variance of the return over the period  $[t, t + 1)$  given the information at date  $t$ .
  - “One day” is merely a convenient term for any elapsed time  $\Delta t$ .

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<sup>a</sup>Bollerslev (1986) and Taylor (1986). They are the “most popular models for time-varying volatility” (Alexander, 2001). A Bloomberg quant said to me on Feb 29, 2008, that GARCH is seldom used in trading.

## GARCH Option Pricing (continued)

- Adopt the following risk-neutral process for price:<sup>a</sup>

$$\ln \frac{S_{t+1}}{S_t} = r - \frac{h_t^2}{2} + h_t \epsilon_{t+1}, \quad (127)$$

where

$$h_{t+1}^2 = \beta_0 + \beta_1 h_t^2 + \beta_2 h_t^2 (\epsilon_{t+1} - c)^2, \quad (128)$$

$$\epsilon_{t+1} \sim N(0, 1) \text{ given information at date } t,$$

$$r = \text{daily riskless return,}$$

$$c \geq 0.$$

- This is called the nonlinear asymmetric GARCH (or NGARCH) model.

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<sup>a</sup>Duan (1995).

## GARCH Option Pricing (continued)

- The five unknown parameters of the model are  $c$ ,  $h_0$ ,  $\beta_0$ ,  $\beta_1$ , and  $\beta_2$ .
- It is postulated that  $\beta_0, \beta_1, \beta_2 \geq 0$  to make the conditional variance positive.
- There are other inequalities to satisfy such as  $\beta_1 + \beta_2 < 1$  (see text).
- It can be shown that  $h_t^2 \geq \min [h_0^2, \beta_0 / (1 - \beta_1)]$ .<sup>a</sup>

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<sup>a</sup>Lyuu & C. Wu (R90723065) (2005).

## GARCH Option Pricing (continued)

- It captures the volatility clustering in asset returns first noted by Mandelbrot (1963).<sup>a</sup>
  - When  $c = 0$ , a large  $\epsilon_{t+1}$  results in a large  $h_{t+1}$ , which in turns tends to yield a large  $h_{t+2}$ , and so on.
- It also captures the negative correlation between the asset return and changes in its (conditional) volatility.<sup>b</sup>
  - For  $c > 0$ , a positive  $\epsilon_{t+1}$  (good news) tends to decrease  $h_{t+1}$ , whereas a negative  $\epsilon_{t+1}$  (bad news) tends to do the opposite.

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<sup>a</sup>“... large changes tend to be followed by large changes—of either sign—and small changes tend to be followed by small changes ...”

<sup>b</sup>Noted by Black (1976): Volatility tends to rise in response to “bad news” and fall in response to “good news.”

## GARCH Option Pricing (continued)

- This is called the leverage effect.
  - A falling stock price raises the fixed costs, relatively speaking.<sup>a</sup>
  - Thus  $c$  is called the leverage effect parameter.
- With  $y_t \triangleq \ln S_t$  denoting the logarithmic price, the model becomes

$$y_{t+1} = y_t + r - \frac{h_t^2}{2} + h_t \epsilon_{t+1}. \quad (129)$$

- The pair  $(y_t, h_t^2)$  completely describes the current state.

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<sup>a</sup>Black (1992).

## GARCH Option Pricing (concluded)

- The conditional mean and variance of  $y_{t+1}$  are clearly

$$E[y_{t+1} | y_t, h_t^2] = y_t + r - \frac{h_t^2}{2}, \quad (130)$$

$$\text{Var}[y_{t+1} | y_t, h_t^2] = h_t^2. \quad (131)$$

- Finally, given  $(y_t, h_t^2)$ , the correlation between  $y_{t+1}$  and  $h_{t+1}$  equals

$$-\frac{2c}{\sqrt{2 + 4c^2}},$$

which is negative for  $c > 0$ .



## GARCH Model: Inferences

- Suppose the parameters  $c$ ,  $h_0$ ,  $\beta_0$ ,  $\beta_1$ , and  $\beta_2$  are given.
- Then we can recover  $h_1, h_2, \dots, h_n$  and  $\epsilon_1, \epsilon_2, \dots, \epsilon_n$  from the prices

$$S_0, S_1, \dots, S_n$$

under the GARCH model (127) on p. 956.

- This is useful in statistical inferences.

## The Ritchken-Trevor (RT) Algorithm<sup>a</sup>

- The GARCH model is a continuous-state model.
- To approximate it, we turn to trees with *discrete* states.
- Path dependence in GARCH makes the tree for asset prices explode exponentially.<sup>b</sup>
- We need to mitigate this combinatorial explosion.

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<sup>a</sup>Ritchken & Trevor (1999).

<sup>b</sup>Why?

## The RT Algorithm (continued)

- Partition a day into  $n$  periods.
- Three states follow each state  $(y_t, h_t^2)$  after a period.
- As the trinomial model combines, each state at date  $t$  is followed by  $2n + 1$  states at date  $t + 1$ .<sup>a</sup>
- These  $2n + 1$  values must approximate the distribution of  $(y_{t+1}, h_{t+1}^2)$  to guarantee convergence.
- So the conditional moments (130)–(131) at date  $t + 1$  on p. 960 must be matched by the trinomial model.

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<sup>a</sup>Recall p. 741.

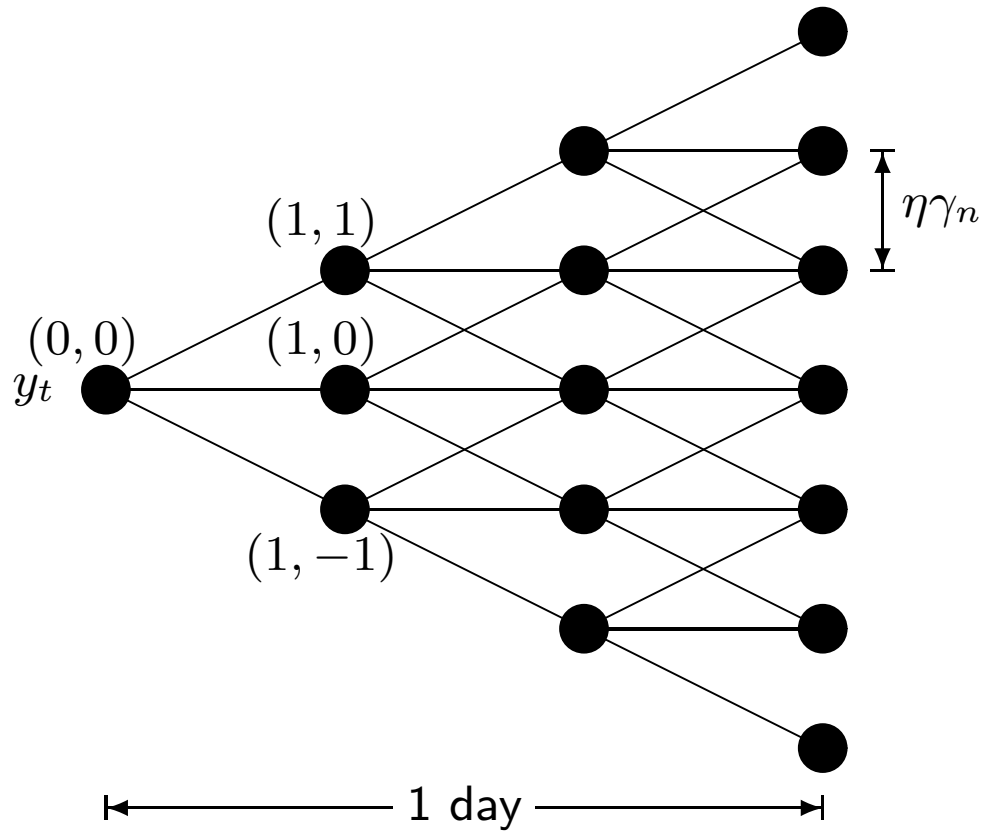
## The RT Algorithm (continued)

- It remains to pick the jump size and the three branching probabilities.
- The role of  $\sigma$  in the Black-Scholes option pricing model is played by  $h_t$  in the GARCH model.
- As a jump size proportional to  $\sigma/\sqrt{n}$  is picked in the BOPM, a comparable magnitude will be chosen here.
- Define  $\gamma \triangleq h_0$ , though other multiples of  $h_0$  are possible.
- Let

$$\gamma_n \triangleq \frac{\gamma}{\sqrt{n}}.$$

## The RT Algorithm (continued)

- The jump size will be some integer multiple  $\eta$  of  $\gamma_n$ .
- We call  $\eta$  the jump parameter (see next page).
- Clearly, the magnitude of  $\eta$  tends to grow with  $h_t$ .
- The middle branch does not change the underlying asset's price.



The seven values on the right approximate the distribution of logarithmic price  $y_{t+1}$ .

## The RT Algorithm (continued)

- The probabilities for the up, middle, and down branches are

$$p_u = \frac{h_t^2}{2\eta^2\gamma^2} + \frac{r - (h_t^2/2)}{2\eta\gamma\sqrt{n}}, \quad (132)$$

$$p_m = 1 - \frac{h_t^2}{\eta^2\gamma^2}, \quad (133)$$

$$p_d = \frac{h_t^2}{2\eta^2\gamma^2} - \frac{r - (h_t^2/2)}{2\eta\gamma\sqrt{n}}. \quad (134)$$

## The RT Algorithm (continued)

- It can be shown that:
  - The trinomial model takes on  $2n + 1$  values at date  $t + 1$  for  $y_{t+1}$ .
  - These values match  $y_{t+1}$ 's mean.
  - These values match  $y_{t+1}$ 's variance asymptotically.
- The central limit theorem guarantees convergence to the continuous-space model as  $n$  increases.<sup>a</sup>

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<sup>a</sup>Assume the probabilities are valid.



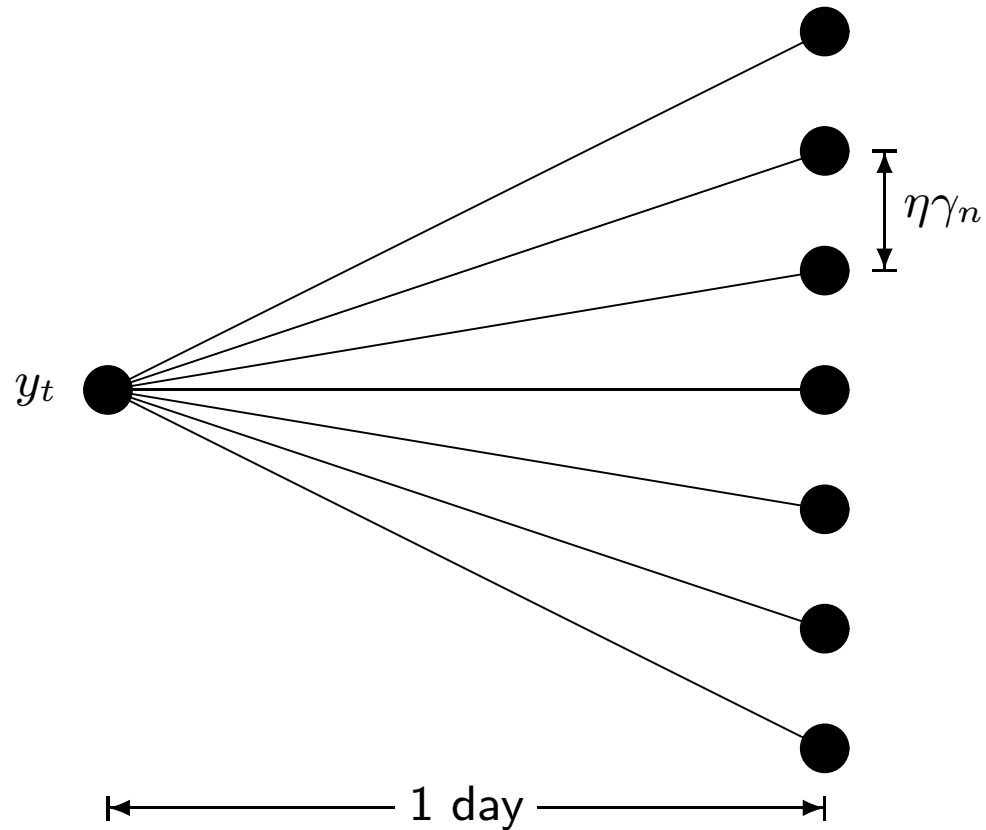
## The RT Algorithm (continued)

- We can dispense with the intermediate nodes *between* dates to create a  $(2n + 1)$ -nomial tree.<sup>a</sup>
- The resulting model is multinomial with  $2n + 1$  branches from any state  $(y_t, h_t^2)$ .
- There are two reasons behind this manipulation.
  - Interdate nodes are created merely to approximate the continuous-state model after one day.
  - Keeping the interdate nodes results in a tree that is  $n$  times larger.<sup>b</sup>

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<sup>a</sup>See p. 970.

<sup>b</sup>Contrast it with the case on p. 412.



This heptanomial model is the outcome of the trinomial tree on p. 966 after the intermediate nodes are removed.

## The RT Algorithm (continued)

- A node with logarithmic price  $y_t + \ell\eta\gamma_n$  at date  $t + 1$  follows the current node at date  $t$  with price  $y_t$ , where

$$-n \leq \ell \leq n.$$

- To reach that price in  $n$  periods, the number of up moves must exceed that of down moves by exactly  $\ell$ .
- The probability this happens is

$$P(\ell) \triangleq \sum_{j_u, j_m, j_d} \frac{n!}{j_u! j_m! j_d!} p_u^{j_u} p_m^{j_m} p_d^{j_d},$$

with  $j_u, j_m, j_d \geq 0$ ,  $n = j_u + j_m + j_d$ , and  $\ell = j_u - j_d$ .

## The RT Algorithm (continued)

- A simple way to calculate the  $P(\ell)$ s starts by noting<sup>a</sup>

$$(p_u x + p_m + p_d x^{-1})^n = \sum_{\ell=-n}^n P(\ell) x^\ell. \quad (135)$$

- Convince yourself that the “accounting” is done correctly.
- So we expand  $(p_u x + p_m + p_d x^{-1})^n$  and retrieve the probabilities by reading off the coefficients.
- It can be computed in  $O(n^2)$  time, if not less.

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<sup>a</sup>C. Wu (R90723065) (2003); Lyuu & C. Wu (R90723065) (2003, 2005).

## The RT Algorithm (continued)

- The updating rule (128) on p. 956 must be modified to account for the adoption of the discrete-state model.
- The logarithmic price  $y_t + \ell\eta\gamma_n$  at date  $t + 1$  following state  $(y_t, h_t^2)$  is associated with this variance:

$$h_{t+1}^2 = \beta_0 + \beta_1 h_t^2 + \beta_2 h_t^2 (\epsilon'_{t+1} - c)^2, \quad (136)$$

– Above, the z-score<sup>a</sup>

$$\epsilon'_{t+1} = \frac{\ell\eta\gamma_n - (r - h_t^2/2)}{h_t}, \quad \ell = 0, \pm 1, \pm 2, \dots, \pm n,$$

is a discrete random variable with  $2n + 1$  values.

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<sup>a</sup>Note that the mean of  $\epsilon'_{t+1}$  is  $r - (h_t^2/2)$ .

## The RT Algorithm (continued)

- Different  $h_t^2$  may require different  $\eta$  so that the probabilities (132)–(134) on p. 967 lie between 0 and 1.
- This implies varying jump sizes  $\eta\gamma_n$ .
- The necessary requirement  $p_m \geq 0$  implies  $\eta \geq h_t/\gamma$ .
- Hence we try

$$\eta = \lceil h_t/\gamma \rceil, \lceil h_t/\gamma \rceil + 1, \lceil h_t/\gamma \rceil + 2, \dots$$

until valid probabilities are obtained or until their nonexistence is confirmed.

## The RT Algorithm (continued)

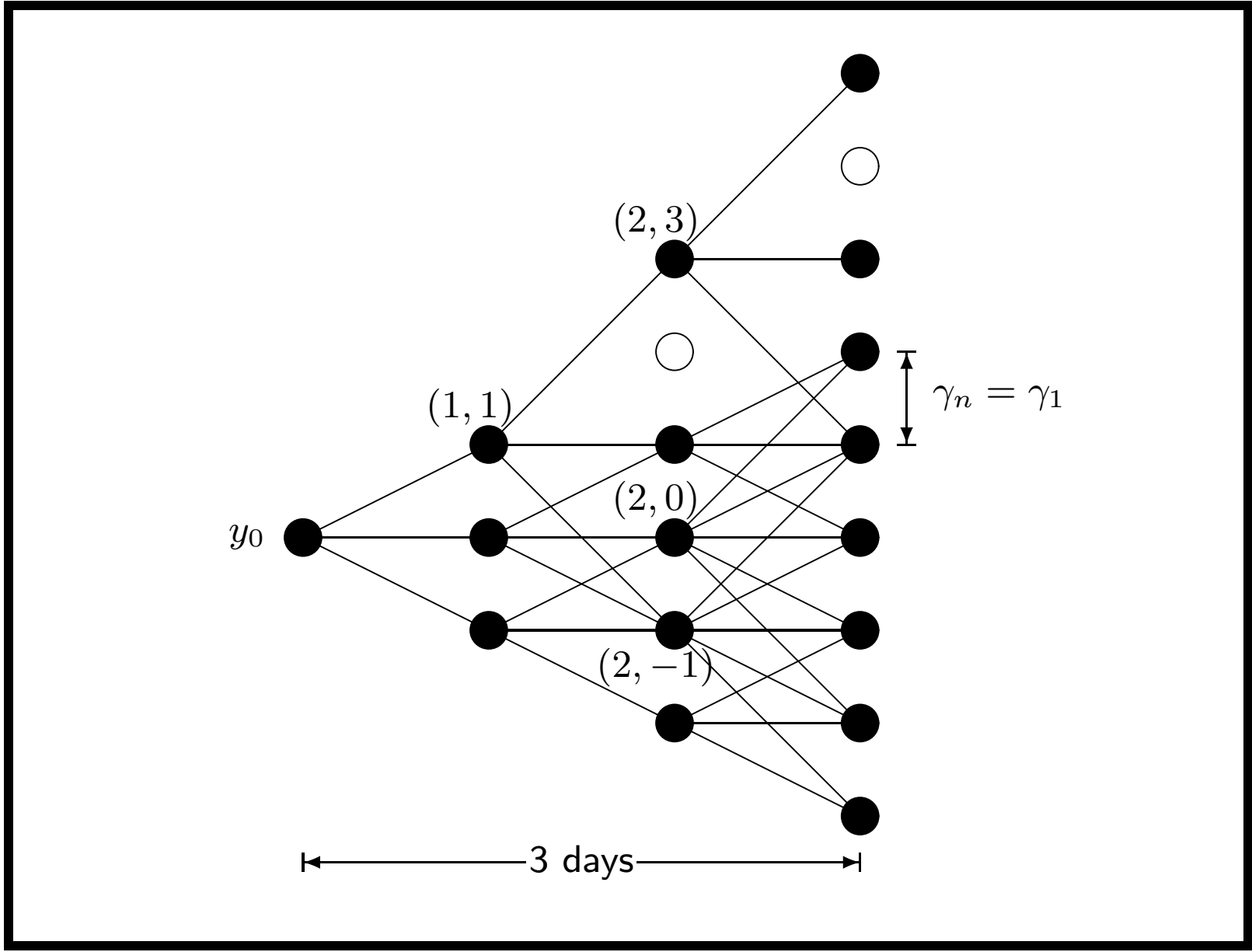
- The sufficient and necessary condition for valid probabilities to exist is<sup>a</sup>

$$\frac{|r - (h_t^2/2)|}{2\eta\gamma\sqrt{n}} \leq \frac{h_t^2}{2\eta^2\gamma^2} \leq \min\left(1 - \frac{|r - (h_t^2/2)|}{2\eta\gamma\sqrt{n}}, \frac{1}{2}\right).$$

- The plot on p. 976 uses  $n = 1$  to illustrate our points for a 3-day model.
- For example, node (1, 1) of date 1 and node (2, 3) of date 2 pick  $\eta = 2$ .

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<sup>a</sup>C. Wu (R90723065) (2003); Lyuu & C. Wu (R90723065) (2003, 2005).





## The RT Algorithm (continued)

- The topology of the tree is not a standard combining multinomial tree.
- For example, a few nodes on p. 976 such as nodes  $(2, 0)$  and  $(2, -1)$  have *multiple* jump sizes.
- The reason is path dependency of the model.
  - Two paths can reach node  $(2, 0)$  from the root node, each with a different variance  $h_t^2$  for the node.
  - One variance results in  $\eta = 1$ .
  - The other results in  $\eta = 2$ .

## The RT Algorithm (concluded)

- The number of possible values of  $h_t^2$  at a node can be exponential.
  - Because each path may result in a different  $h_t^2$ .
- To address this problem, we record only the maximum and minimum  $h_t^2$  at each node.<sup>a</sup>
- Therefore, each node on the tree contains only two states  $(y_t, h_{\max}^2)$  and  $(y_t, h_{\min}^2)$ .
- Each of  $(y_t, h_{\max}^2)$  and  $(y_t, h_{\min}^2)$  carries its own  $\eta$  and set of  $2n + 1$  branching probabilities.

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<sup>a</sup>Cakici & Topyan (2000). But see p. 1013 for a potential problem.

## Negative Aspects of the Ritchken-Trevor Algorithm<sup>a</sup>

- A small  $n$  may yield inaccurate option prices.
- But the tree will grow exponentially if  $n$  is large enough.
  - Specifically,  $n > (1 - \beta_1)/\beta_2$  when  $r = c = 0$ .
- A large  $n$  has another serious problem: The tree cannot grow beyond a certain date.
- Thus the choice of  $n$  may be quite limited in practice.
- The RT algorithm can be modified to be free of shortened maturity and exponential complexity.<sup>b</sup>

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<sup>a</sup>Lyuu & C. Wu (R90723065) (2003, 2005).

<sup>b</sup>Its size is only  $O(T^2)$  if  $n \leq (\sqrt{(1 - \beta_1)/\beta_2} - c)^2$ , where  $T$  is the number of days to maturity!

## Numerical Examples

- Assume
  - $S_0 = 100$ ,  $y_0 = \ln S_0 = 4.60517$ .
  - $r = 0$ .
  - $n = 1$ .
  - $h_0^2 = 0.0001096$ ,  $\gamma = h_0 = 0.010469$ .
  - $\gamma_n = \gamma/\sqrt{n} = 0.010469$ .
  - $\beta_0 = 0.000006575$ ,  $\beta_1 = 0.9$ ,  $\beta_2 = 0.04$ , and  $c = 0$ .

## Numerical Examples (continued)

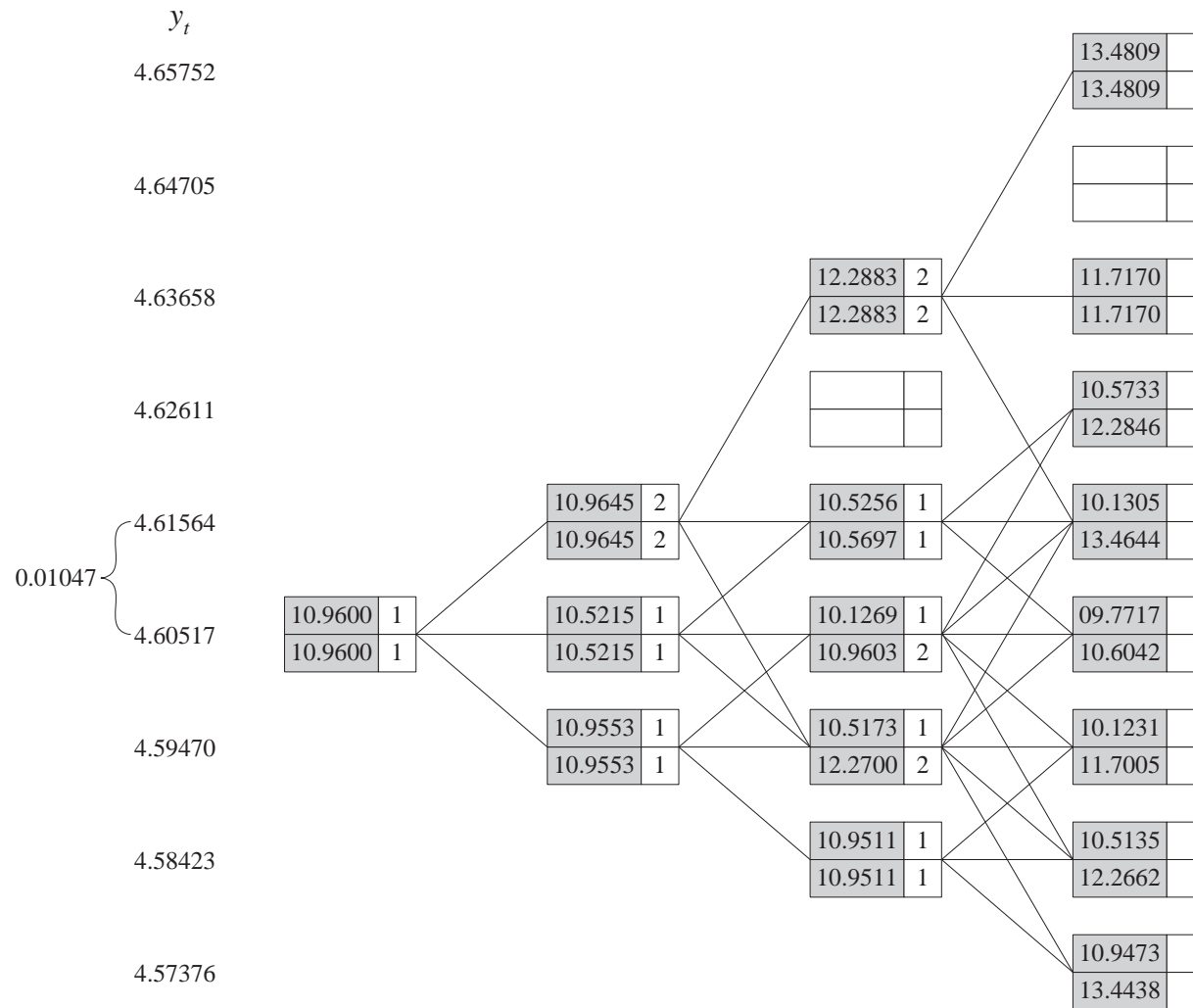
- A daily variance of 0.0001096 corresponds to an annual volatility of

$$\sqrt{365 \times 0.0001096} \approx 20\%.$$

- Let  $h^2(i, j)$  denote the variance at node  $(i, j)$ .
- Initially,  $h^2(0, 0) = h_0^2 = 0.0001096$ .

## Numerical Examples (continued)

- Let  $h_{\max}^2(i, j)$  denote the maximum variance at node  $(i, j)$ .
- Let  $h_{\min}^2(i, j)$  denote the minimum variance at node  $(i, j)$ .
- Initially,  $h_{\max}^2(0, 0) = h_{\min}^2(0, 0) = h_0^2$ .
- The resulting 3-day tree is depicted on p. 983.



## Numerical Examples (continued)

- A top number inside a gray box refers to the minimum variance  $h_{\min}^2$  for the node.
- A bottom number inside a gray box refers to the maximum variance  $h_{\max}^2$  for the node.
- Variances are multiplied by 100,000 for readability.
- The top number inside a white box refers to the  $\eta$  for  $h_{\min}^2$ .
- The bottom number inside a white box refers to the  $\eta$  for  $h_{\max}^2$ .



## Numerical Examples (continued)

- Let us see how the numbers are calculated.
- Start with the root node, node  $(0, 0)$ .
- Try  $\eta = 1$  in Eqs. (132)–(134) on p. 967 first to obtain

$$p_u = 0.4974,$$

$$p_m = 0,$$

$$p_d = 0.5026.$$

- As they are valid, the three branches from the root node take single jumps.

## Numerical Examples (continued)

- Move on to node  $(1, 1)$ .
- It has one predecessor node—node  $(0, 0)$ —and it takes an up move to reach node  $(1, 1)$ .
- So apply updating rule (136) on p. 973 with  $\ell = 1$  and  $h_t^2 = h^2(0, 0)$ .
- The result is  $h^2(1, 1) = 0.000109645$ .

## Numerical Examples (continued)

- Because  $\lceil h(1, 1)/\gamma \rceil = 2$ , we try  $\eta = 2$  in Eqs. (132)–(134) on p. 967 first to obtain

$$p_u = 0.1237,$$

$$p_m = 0.7499,$$

$$p_d = 0.1264.$$

- As they are valid, the three branches from node  $(1, 1)$  take double jumps.

## Numerical Examples (continued)

- Carry out similar calculations for node  $(1, 0)$  with  $\ell = 0$  in updating rule (136) on p. 973.
- Carry out similar calculations for node  $(1, -1)$  with  $\ell = -1$  in updating rule (136).
- Single jump  $\eta = 1$  works for both nodes.
- The resulting variances are

$$\begin{aligned}h^2(1, 0) &= 0.000105215, \\h^2(1, -1) &= 0.000109553.\end{aligned}$$

## Numerical Examples (continued)

- Node  $(2, 0)$  has 2 predecessor nodes,  $(1, 0)$  and  $(1, -1)$ .
- Both have to be considered in deriving the variances.
- Let us start with node  $(1, 0)$ .
- Because it takes a middle move to reach node  $(2, 0)$ , we apply updating rule (136) on p. 973 with  $\ell = 0$  and  $h_t^2 = h^2(1, 0)$ .
- The result is  $h_{t+1}^2 = 0.000101269$ .

## Numerical Examples (continued)

- Now move on to the other predecessor node  $(1, -1)$ .
- Because it takes an up move to reach node  $(2, 0)$ , apply updating rule (136) on p. 973 with  $\ell = 1$  and  $h_t^2 = h^2(1, -1)$ .
- The result is  $h_{t+1}^2 = 0.000109603$ .
- We hence record

$$h_{\min}^2(2, 0) = 0.000101269,$$

$$h_{\max}^2(2, 0) = 0.000109603.$$

## Numerical Examples (continued)

- Consider state  $h_{\max}^2(2, 0)$  first.
- Because  $\lceil h_{\max}(2, 0)/\gamma \rceil = 2$ , we first try  $\eta = 2$  in Eqs. (132)–(134) on p. 967 to obtain

$$p_u = 0.1237,$$

$$p_m = 0.7500,$$

$$p_d = 0.1263.$$

- As they are valid, the three branches from node  $(2, 0)$  with the maximum variance take double jumps.

## Numerical Examples (continued)

- Now consider state  $h_{\min}^2(2, 0)$ .
- Because  $\lceil h_{\min}(2, 0)/\gamma \rceil = 1$ , we first try  $\eta = 1$  in Eqs. (132)–(134) on p. 967 to obtain

$$p_u = 0.4596,$$

$$p_m = 0.0760,$$

$$p_d = 0.4644.$$

- As they are valid, the three branches from node  $(2, 0)$  with the minimum variance take single jumps.



## Numerical Examples (continued)

- Node  $(2, -1)$  has 3 predecessor nodes.
- Start with node  $(1, 1)$ .
- Because it takes *one* down move to reach node  $(2, -1)$ , we apply updating rule (136) on p. 973 with  $\ell = -1$  and  $h_t^2 = h^2(1, 1)$ .<sup>a</sup>
- The result is  $h_{t+1}^2 = 0.0001227$ .

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<sup>a</sup>Note that it is *not*  $\ell = -2$ . The reason is that  $h(1, 1)$  has  $\eta = 2$  (p. 987).

## Numerical Examples (continued)

- Now move on to predecessor node  $(1, 0)$ .
- Because it also takes a down move to reach node  $(2, -1)$ , we apply updating rule (136) on p. 973 with  $\ell = -1$  and  $h_t^2 = h^2(1, 0)$ .
- The result is  $h_{t+1}^2 = 0.000105609$ .

## Numerical Examples (continued)

- Finally, consider predecessor node  $(1, -1)$ .
- Because it takes a middle move to reach node  $(2, -1)$ , we apply updating rule (136) on p. 973 with  $\ell = 0$  and  $h_t^2 = h^2(1, -1)$ .
- The result is  $h_{t+1}^2 = 0.000105173$ .
- We hence record

$$\begin{aligned}h_{\min}^2(2, -1) &= 0.000105173, \\h_{\max}^2(2, -1) &= 0.0001227.\end{aligned}$$

## Numerical Examples (continued)

- Consider state  $h_{\max}^2(2, -1)$ .
- Because  $\lceil h_{\max}(2, -1)/\gamma \rceil = 2$ , we first try  $\eta = 2$  in Eqs. (132)–(134) on p. 967 to obtain

$$p_u = 0.1385,$$

$$p_m = 0.7201,$$

$$p_d = 0.1414.$$

- As they are valid, the three branches from node  $(2, -1)$  with the maximum variance take double jumps.

## Numerical Examples (continued)

- Next, consider state  $h_{\min}^2(2, -1)$ .
- Because  $\lceil h_{\min}(2, -1)/\gamma \rceil = 1$ , we first try  $\eta = 1$  in Eqs. (132)–(134) on p. 967 to obtain

$$p_u = 0.4773,$$

$$p_m = 0.0404,$$

$$p_d = 0.4823.$$

- As they are valid, the three branches from node  $(2, -1)$  with the minimum variance take single jumps.

## Numerical Examples (concluded)

- Other nodes at dates 2 and 3 can be handled similarly.
- In general, if a node has  $k$  predecessor nodes, then up to  $2k$  variances will be calculated using the updating rule.
  - This is because each predecessor node keeps *two* variance numbers.
- But only the maximum and minimum variances will be kept.

## Negative Aspects of the RT Algorithm Revisited<sup>a</sup>

- Recall the problems mentioned on p. 979.
- In our case, combinatorial explosion occurs when

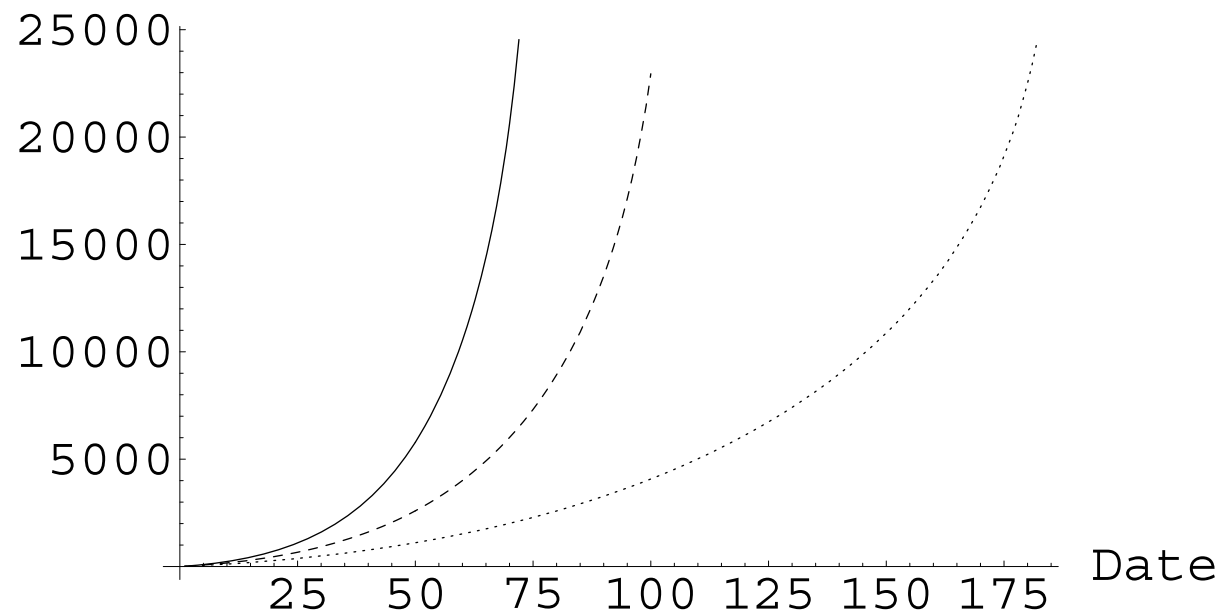
$$n > \frac{1 - \beta_1}{\beta_2} = \frac{1 - 0.9}{0.04} = 2.5$$

(see the next plot).

- Suppose we are willing to accept the exponential running time and pick  $n = 100$  to seek accuracy.
- But the problem of shortened maturity forces the tree to stop at date 9!

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<sup>a</sup>Lyyu & C. Wu (R90723065) (2003, 2005).



Dotted line:  $n = 3$ ; dashed line:  $n = 4$ ; solid line:  $n = 5$ .



## Backward Induction on the RT Tree

- After the RT tree is constructed, it can be used to price options by backward induction.
- Recall that each node keeps two variances  $h_{\max}^2$  and  $h_{\min}^2$ .
- We now increase that number to  $K$  equally spaced variances between  $h_{\max}^2$  and  $h_{\min}^2$  at each node.
- Besides the minimum and maximum variances, the other  $K - 2$  variances in between are linearly interpolated.<sup>a</sup>

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<sup>a</sup>Log-linear interpolation works better in practice (Lyu & C. Wu (R90723065), 2005). Log-cubic interpolation works even better (C. Liu (R92922123), 2005).

## Backward Induction on the RT Tree (continued)

- For example, if  $K = 3$ , then a variance of

$$10.5436 \times 10^{-6}$$

will be added between the maximum and minimum variances at node  $(2, 0)$  on p. 983.<sup>a</sup>

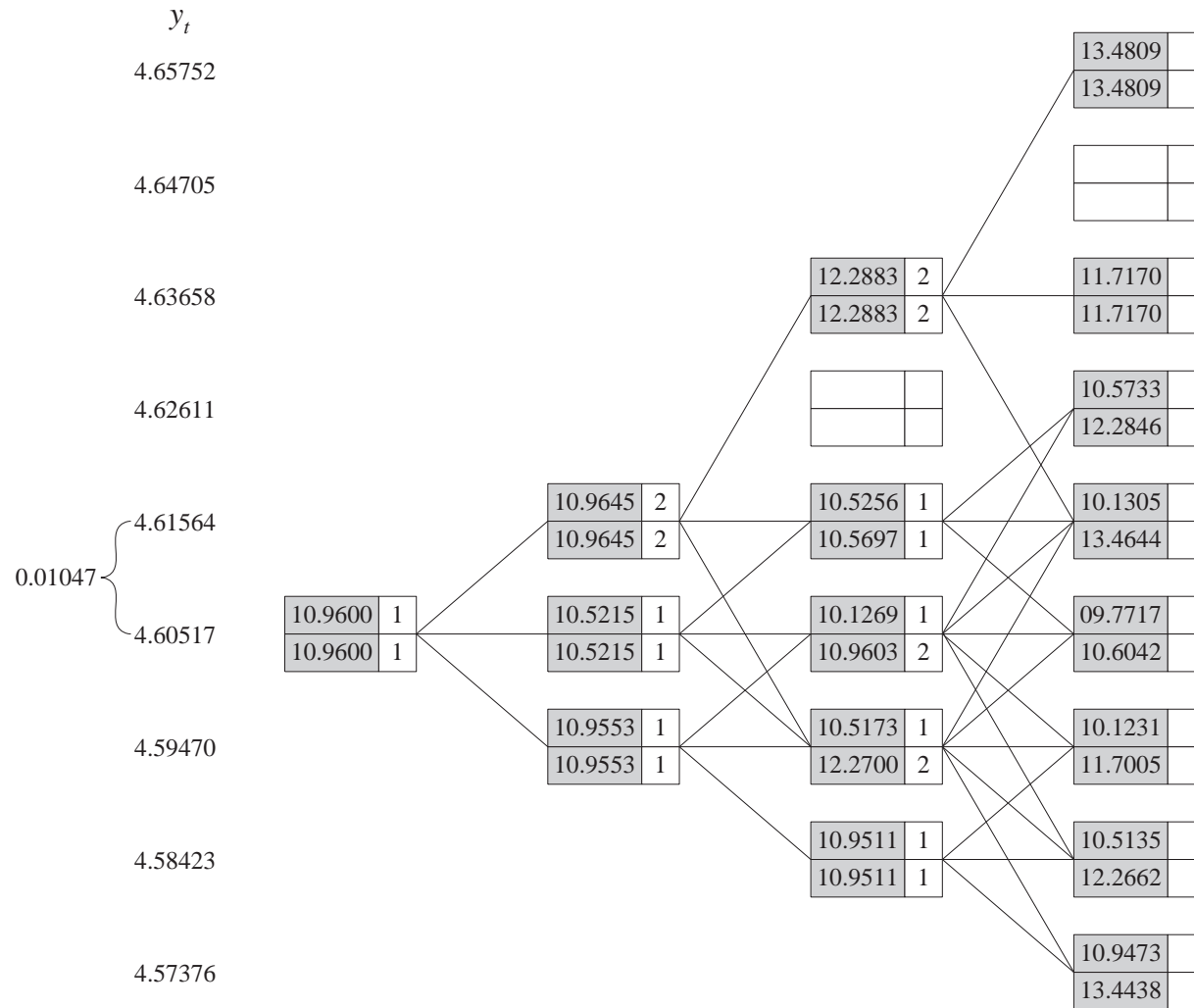
- In general, the  $k$ th variance at node  $(i, j)$  is

$$h_{\min}^2(i, j) + k \frac{h_{\max}^2(i, j) - h_{\min}^2(i, j)}{K - 1}, \quad k = 0, 1, \dots, K - 1.$$

- Each interpolated variance's jump parameter and branching probabilities can be computed as before.

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<sup>a</sup>Repeated on p. 1003.



## Backward Induction on the RT Tree (concluded)

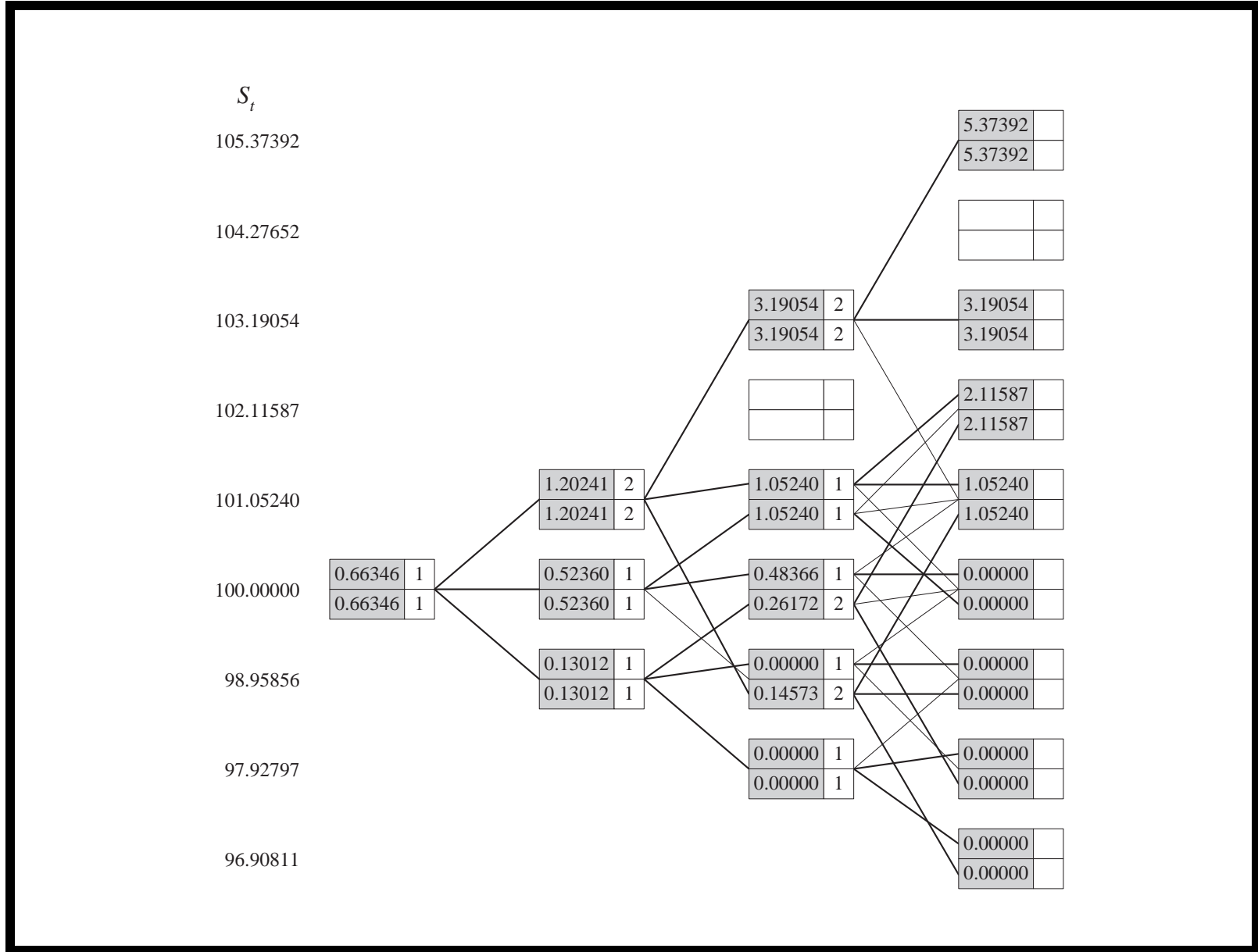
- Suppose a variance falls between two of the  $K$  variances during backward induction.
- Linear interpolation of the option prices corresponding to the two bracketing variances will be used as the approximate option price.
- The above idea is reminiscent of the one in dealing with Asian options.<sup>a</sup>

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<sup>a</sup>Recall p. 454.

## Numerical Examples

- We next use the tree on p. 1003 to price a European call option with a strike price of 100 and expiring at date 3.
- Recall that the riskless interest rate is zero.
- Assume  $K = 2$ ; hence there are no interpolated variances.
- The pricing tree is shown on p. 1006 with a call price of 0.66346.
  - The branching probabilities needed in backward induction can be found on p. 1007.



$rb[i][0]$			
$rb[i][1]$			
$rb[0][ ]$	$rb[1][ ]$	$rb[2][ ]$	$rb[3][ ]$
0	-1	-2	-3
0	1	3	5

$h^2[i][j][0]$			
$h^2[i][j][1]$			
$h^2[3][ ][ ]$			
13.4809			
13.4809			
$h^2[2][ ][ ]$			
12.2883			
11.7170			
12.2883			
11.7170			
$h^2[1][ ][ ]$			
10.9645			
10.5256			
10.1305			
10.9645			
10.5697			
13.4644			
$h^2[0][ ][ ]$			
10.9600			
10.5215			
10.1269			
09.7717			
10.9600			
10.5215			
10.9603			
10.6042			
10.9553			
10.5173			
10.1231			
10.9553			
12.2700			
11.7005			
10.9511			
10.5135			
10.1231			
10.9511			
12.2662			
10.9473			
13.4438			

$\eta[i][j][0]$		
$\eta[i][j][1]$		
$\eta[2][ ][ ]$		
2		
2		
$\eta[1][ ][ ]$		
2		
1		
1		
$\eta[0][ ][ ]$		
2		
1		
1		
2		
1		
2		
1		

$p[i][j][0][1]$			
$p[i][j][1][1]$			
$p[i][j][0][0]$			
$p[i][j][1][0]$			
$p[i][j][0][-1]$			
$p[i][j][1][-1]$			
$p[2][ ][ ][ ]$			
0.1387 0.1387			
0.7197 0.7197			
0.1416 0.1416			
$p[1][ ][ ][ ]$			
0.1237 0.1237			
0.7499 0.7499			
0.1264 0.1264			
0.4777 0.4797			
0.0396 0.0356			
0.4827 0.4847			
$p[0][ ][ ][ ]$			
0.4974 0.4974			
0.0000 0.0000			
0.5026 0.5026			
0.4775 0.4775			
0.0400 0.0400			
0.4825 0.4825			
0.4596 0.1237			
0.0760 0.7500			
0.4644 0.1263			
0.4972 0.4972			
0.0004 0.0004			
0.5024 0.5024			
0.4773 0.1385			
0.0404 0.7201			
0.4823 0.1414			
0.4970 0.4970			
0.0008 0.0008			
0.5022 0.5022			

## Numerical Examples (continued)

- Let us derive some of the numbers on p. 1006.
- A gray line means the updated variance falls strictly between  $h_{\max}^2$  and  $h_{\min}^2$ .
- The option price for a terminal node at date 3 equals  $\max(S_3 - 100, 0)$ , independent of the variance level.
- Now move on to nodes at date 2.
- The option price at node  $(2, 3)$  depends on those at nodes  $(3, 5)$ ,  $(3, 3)$ , and  $(3, 1)$ .
- It therefore equals

$$0.1387 \times 5.37392 + 0.7197 \times 3.19054 + 0.1416 \times 1.05240 = 3.19054.$$



## Numerical Examples (continued)

- Option prices for other nodes at date 2 can be computed similarly.
- For node  $(1, 1)$ , the option price for both variances is
$$0.1237 \times 3.19054 + 0.7499 \times 1.05240 + 0.1264 \times 0.14573 = 1.20241.$$
- Node  $(1, 0)$  is most interesting.
- We knew that a down move from it gives a variance of 0.000105609.
- This number falls between the minimum variance 0.000105173 and the maximum variance 0.0001227 at node  $(2, -1)$  on p. 1003.

## Numerical Examples (continued)

- The option price corresponding to the minimum variance is 0 (p. 1006).
- The option price corresponding to the maximum variance is 0.14573.
- The equation

$$x \times 0.000105173 + (1 - x) \times 0.0001227 = 0.000105609$$

is satisfied by  $x = 0.9751$ .

- So the option for the down state is approximated by

$$x \times 0 + (1 - x) \times 0.14573 = 0.00362.$$

## Numerical Examples (continued)

- The up move leads to the state with option price 1.05240.
- The middle move leads to the state with option price 0.48366.
- The option price at node  $(1, 0)$  is finally calculated as

$$0.4775 \times 1.05240 + 0.0400 \times 0.48366 + 0.4825 \times 0.00362 = 0.52360.$$

## Numerical Examples (continued)

- A variance following an interpolated variance may exceed the maximum variance or be lower than the minimum variance.
- When this happens, the option price corresponding to the maximum or minimum variance will be used during backward induction.<sup>a</sup>
- This act tends to reduce the dynamic range of the variance, however.

---

<sup>a</sup>Cakici & Topyan (2000).

## Numerical Examples (concluded)

- Worse, an interpolated variance may choose a branch that goes into a node that is *not* reached in forward induction.<sup>a</sup>
- In this case, the algorithm fails.
- The RT algorithm does not have this problem.
  - This is because all interpolated variances are involved in the forward-induction phase.
- It may be hard to calculate the implied  $\beta_1$  and  $\beta_2$  from option prices.<sup>b</sup>

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<sup>a</sup>Lyu & C. Wu (R90723065) (2005).

<sup>b</sup>Y. Chang (B89704039, R93922034) (2006).

## Complexities of GARCH Models<sup>a</sup>

- The RT algorithm explodes exponentially even for moderate  $n$ .<sup>b</sup>
- The mean-tracking tree of Lyuu and Wu (2005) guarantees explosion not to happen for  $n$  not too large.
  - That tree is similar to, but earlier than, the binomial-trinomial tree.<sup>c</sup>
  - In fact, we can use the binomial-trinomial tree here, and everything goes through.<sup>d</sup>

---

<sup>a</sup>Lyuu & C. Wu (R90723065) (2003, 2005).

<sup>b</sup>Recall p. 979.

<sup>c</sup>Recall pp. 764ff.

<sup>d</sup>Contributed by Mr. Lu, Zheng-Liang (D00922011) on August 12, 2021.

## Complexities of GARCH Models (continued)

- The next page summarizes the situations for many GARCH option pricing models other than NGARCH.

## Complexities of GARCH Models (concluded)<sup>a</sup>

Model	Explosion	Non-explosion
NGARCH	$\beta_1 + \beta_2 n > 1$	$\beta_1 + \beta_2(\sqrt{n} + \lambda + c)^2 \leq 1$
LGARCH	$\beta_1 + \beta_2 n > 1$	$\beta_1 + \beta_2(\sqrt{n} + \lambda)^2 \leq 1$
AGARCH	$\beta_1 + \beta_2 n > 1$	$\beta_1 + \beta_2(\sqrt{n} + \lambda)^2 \leq 1$
GJR-GARCH	$\beta_1 + \beta_2 n > 1$	$\beta_1 + (\beta_2 + \beta_3)(\sqrt{n} + \lambda)^2 \leq 1$
TS-GARCH	$\beta_1 + \beta_2\sqrt{n} > 1$	$\beta_1 + \beta_2(\lambda + \sqrt{n}) \leq 1$
TGARCH	$\beta_1 + \beta_2\sqrt{n} > 1$	$\beta_1 + (\beta_2 + \beta_3)(\lambda + \sqrt{n}) \leq 1$
Heston-Nandi	$\beta_1 + \beta_2(c - \frac{1}{2})^2 > 1$ & $c \leq \frac{1}{2}$	$\beta_1 + \beta_2 c^2 \leq 1$
VGARCH	$\beta_1 + (\beta_2/4) > 1$	$\beta_1 \leq 1$

<sup>a</sup>Y. C. Chen (R95723051) (2008); Y. C. Chen (R95723051), Lyuu, & Wen (D94922003) (2012).



## Obtaining Profit and Loss of Delta Hedge

- Profit and loss of any hedging strategy should be calculated under the real-world probability measure.<sup>a</sup>
- But hedging parameters such as delta should be computed under the risk-neutral measure.
- Say we want the distribution of profit and loss for the delta hedge under the GARCH model.
- If a tree is built for each sampled stock price to obtain the delta, the complexity will be astronomical.<sup>b</sup>
- How to do it efficiently?<sup>c</sup>

---

<sup>a</sup>Recall p. 713.

<sup>b</sup>Augustyniak, Badescu, & Guo (2021).

<sup>c</sup>Lu (D00922011), Lyuu, & Yang (D09922005) (2021).

*Introduction to Term Structure Modeling*

The fox often ran to the hole  
by which they had come in,  
to find out if his body was still thin enough  
to slip through it.  
— *Grimm's Fairy Tales*

And the worst thing you can have  
is models and spreadsheets.  
— Warren Buffet (2008, May 3)

Renaissance is 100% model driven.<sup>a</sup>  
James Simons (2015, May 13, 37:09)

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<sup>a</sup><https://www.youtube.com/watch?v=QNznD9hMEh0>

## Outline

- Use the binomial interest rate tree to model stochastic term structure.
  - Illustrates the basic ideas underlying future models.
  - Applications are generic in that pricing and hedging methodologies can be easily adapted to other models.
- Although the idea is similar to the earlier one used in option pricing, the current task is more complicated.
  - The evolution of an entire term structure, not just a single stock price, is to be modeled.
  - Interest rates of various maturities cannot evolve arbitrarily, or arbitrage profits may occur.

## Goals

- A stochastic interest rate model performs two tasks.
  - Provides a stochastic process that defines future term structures without arbitrage profits.
  - “Consistent” with the observed term structures.

## History

- The methodology was founded by Merton (1970).
- Modern interest rate modeling is often traced to 1977 when Vasicek and Cox, Ingersoll, and Ross developed simultaneously their influential models.
- Early models have fitting problems because they may not price today's benchmark bonds correctly.
- An alternative approach pioneered by Ho and Lee (1986) makes fitting the market yield curve mandatory.
- Models based on such a paradigm are called arbitrage-free or no-arbitrage models.<sup>a</sup>

---

<sup>a</sup>Somewhat misleadingly.

## Binomial Interest Rate Tree

- Goal is to construct a no-arbitrage interest rate tree consistent with the yields — and sometimes yield volatilities — of zero-coupon bonds of all maturities.
  - This procedure is called calibration.<sup>a</sup>
- Pick a binomial tree model in which the logarithm of the future short rate obeys the binomial distribution.
  - Like the CRR tree for pricing options.
- The limiting distribution of the short rate at any future time is hence lognormal.

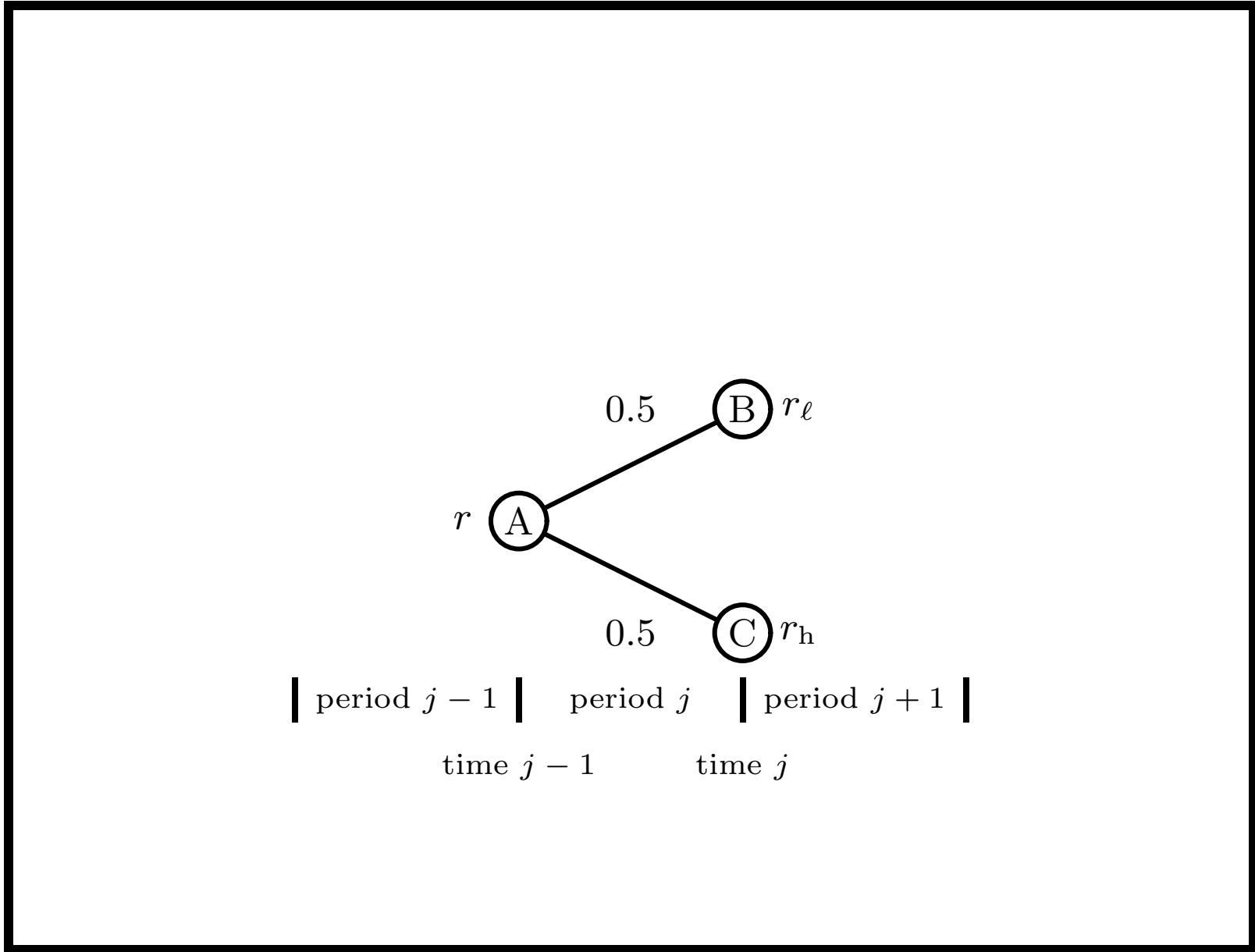
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<sup>a</sup>Derman (2004), “complexity without calibration is pointless.”



## Binomial Interest Rate Tree (continued)

- A binomial tree of future short rates is constructed.
- Every short rate is followed by two short rates in the following period.
- In the figure on p. 1026, node A coincides with the start of period  $j$  during which the short rate  $r$  is in effect.
- At the conclusion of period  $j$ , a new short rate goes into effect for period  $j + 1$ .



## Binomial Interest Rate Tree (continued)

- This may take one of two possible values:
  - $r_\ell$ : the “low” short-rate outcome at node B.
  - $r_h$ : the “high” short-rate outcome at node C.
- Each branch has a 50% chance of occurring in a risk-neutral economy.
- We require that the paths combine as the binomial process unfolds.
- Tuckman (2002) attributes this model to Salomon Brothers.

## Binomial Interest Rate Tree (continued)

- The short rate  $r$  can go to  $r_h$  and  $r_\ell$  with equal risk-neutral probability  $1/2$  in a period of length  $\Delta t$ .
- Hence the volatility of  $\ln r$  after  $\Delta t$  time is<sup>a</sup>

$$\sigma = \frac{1}{2} \frac{1}{\sqrt{\Delta t}} \ln \left( \frac{r_h}{r_\ell} \right). \quad (137)$$

- Above,  $\sigma$  is annualized,<sup>b</sup> whereas  $r_\ell$  and  $r_h$  are period based.

---

<sup>a</sup>See Exercise 23.2.3 in text.

<sup>b</sup>You may remove the  $1/\sqrt{\Delta t}$  term to return it to being period based.

## Binomial Interest Rate Tree (continued)

- Note that

$$\frac{r_h}{r_\ell} = e^{2\sigma\sqrt{\Delta t}}.$$

- Thus greater volatility, hence uncertainty, leads to larger  $r_h/r_\ell$  and wider ranges of possible short rates.
- The ratio  $r_h/r_\ell$  may depend on time if the volatility is a function of time.
- Note that  $r_h/r_\ell$  has nothing to do with the current short rate  $r$  if  $\sigma$  is independent of  $r$ .

## Binomial Interest Rate Tree (continued)

- In general there are  $j$  possible rates for *period*  $j$ ,<sup>a</sup>

$$r_j, r_j v_j, r_j v_j^2, \dots, r_j v_j^{j-1},$$

where

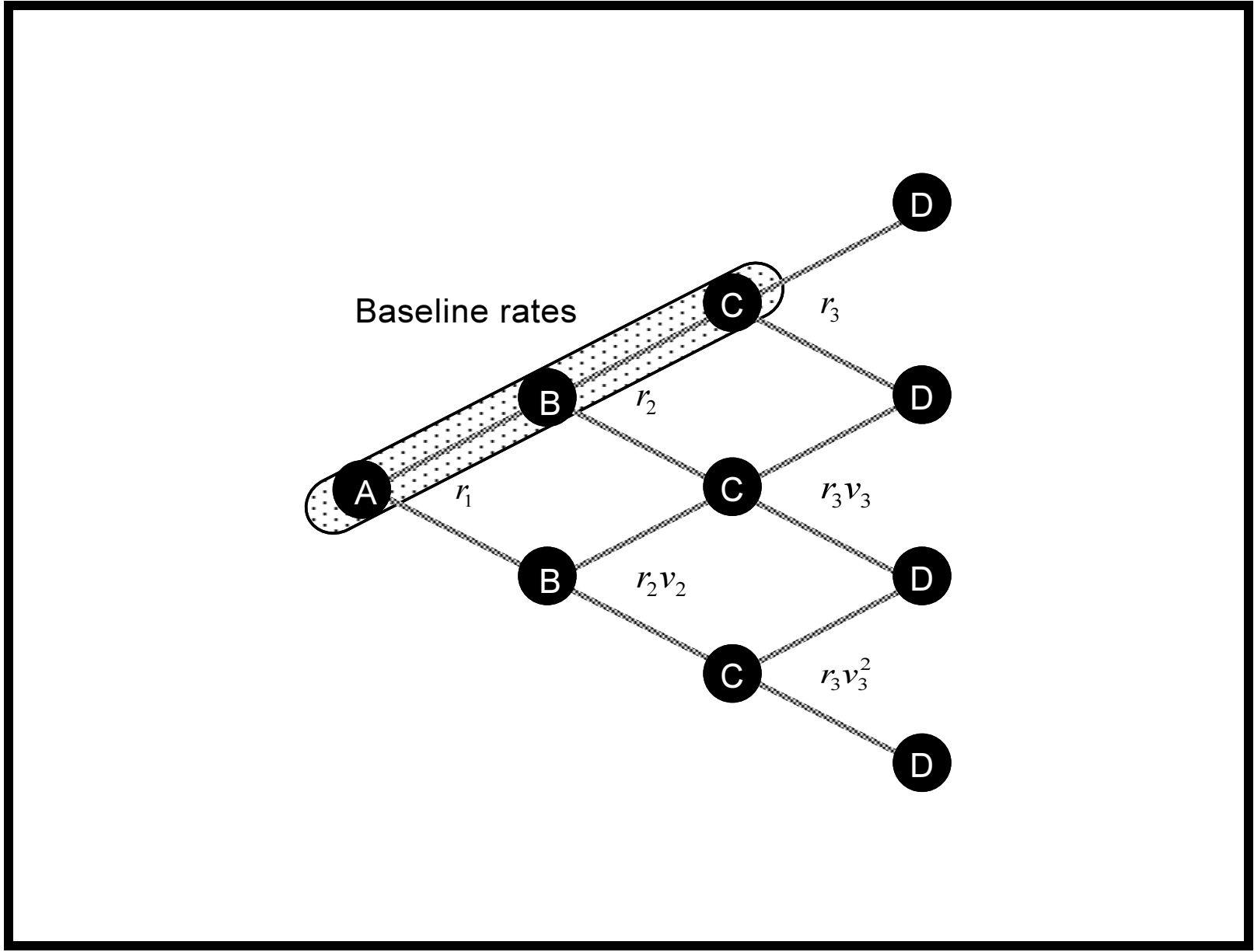
$$v_j \triangleq e^{2\sigma_j \sqrt{\Delta t}} = 1 + O\left(\sqrt{\Delta t}\right) \quad (138)$$

is the multiplicative ratio for the rates in period  $j$  (see figure on next page).

- We shall call  $r_j$  the baseline rates.
- The subscript  $j$  in  $\sigma_j$  means to emphasize that the short rate volatility may be time dependent.

---

<sup>a</sup>Not  $j + 1$ .



## Binomial Interest Rate Tree (concluded)

- In the limit, the short rate follows

$$r(t) = \mu(t) e^{\sigma(t) W(t)}. \quad (139)$$

- The (percent) short rate volatility  $\sigma(t)$  is a deterministic function of time.
- The expected value of  $r(t)$  equals  $\mu(t) e^{\sigma(t)^2(t/2)}$ .
- Hence a *declining* short rate volatility is needed to preclude the short rate from assuming implausibly high values.
- This is how the binomial interest rate tree achieves mean reversion to some long-term mean.



## Memory Issues

- Path independency: The term structure at any node is independent of the path taken to reach it.
- So only the baseline rates  $r_i$  and the multiplicative ratios  $v_i$  need to be stored in computer memory.
- This takes up only  $O(n)$  space.<sup>a</sup>
- Storing the whole tree would take up  $O(n^2)$  space.
  - Daily interest rate movements for 30 years require roughly  $(30 \times 365)^2/2 \approx 6 \times 10^7$  double-precision floating-point numbers (half a gigabyte!).

---

<sup>a</sup>Throughout,  $n$  denotes the depth of the tree.

## Set Things in Motion

- The abstract process is now in place.
- We need the yields to maturities of the riskless bonds that make up the benchmark yield curve and their volatilities.
- In the U.S., for example, the on-the-run yield curve obtained by the most recently issued Treasury securities may be used as the benchmark curve.

## Set Things in Motion (concluded)

- The term structure of (yield) volatilities<sup>a</sup> can be estimated from:
  - Historical data (historical volatility).
  - Or interest rate option prices such as cap prices (implied volatility).
- The binomial tree should be found that is consistent with both term structures.
- Here we focus on the term structure of interest rates.

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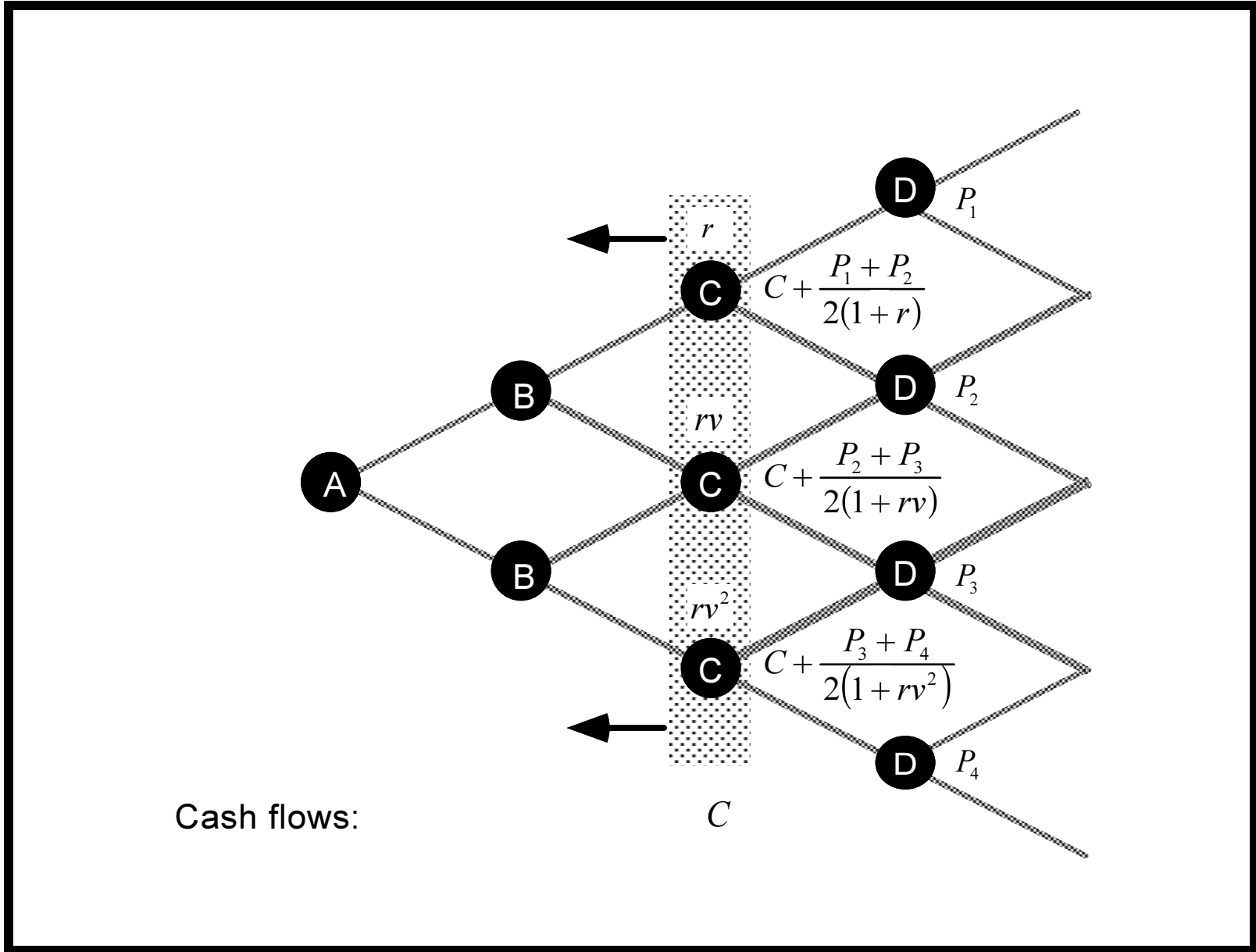
<sup>a</sup>Or simply the volatility (term) structure.

## Model Term Structures

- The model price is computed by backward induction.
- Refer back to the figure on p. 1026.
- Given that the values at nodes B and C are  $P_B$  and  $P_C$ , respectively, the value at node A is then

$$\frac{P_B + P_C}{2(1 + r)} + \text{cash flow at node A.}$$

- We compute the values column by column (see next page).
- This takes  $O(n^2)$  time and  $O(n)$  space.



## Term Structure Dynamics

- An  $n$ -period zero-coupon bond's price can be computed by assigning \$1 to every node at time  $n$  and then applying backward induction.
- Repeat this step for  $n = 1, 2, \dots$  to obtain the market discount function implied by the tree.
- The tree therefore determines a term structure.
- It also contains a term structure dynamics.
  - Every node in the tree induces a binomial interest rate tree and a term structure.

## Sample Term Structure

- We shall construct interest rate trees consistent with the sample term structure in the table below.
  - This is calibration (the reverse of pricing).
- Assume the short rate volatility is such that

$$v \triangleq \frac{\Delta r_h}{r_\ell} = 1.5,$$

independent of time.

Period	1	2	3
Spot rate (%)	4	4.2	4.3
One-period forward rate (%)	4	4.4	4.5
Discount factor	0.96154	0.92101	0.88135

## An Approximate Calibration Scheme

- Start with the implied one-period forward rates.
- Equate the expected short rate with the forward rate.<sup>a</sup>
- For the first period, the forward rate is today's one-period spot rate.
- In general, let  $f_j$  denote the forward rate in period  $j$ .
- This forward rate can be derived from the market discount function via<sup>b</sup>

$$f_j = \frac{d(j)}{d(j+1)} - 1.$$

---

<sup>a</sup>See Exercise 5.6.6 in text for the motivation.

<sup>b</sup>See Exercise 5.6.3 in text.



## An Approximate Calibration Scheme (continued)

- As the  $i$ th short rate  $r_j v_j^{i-1}$ ,  $1 \leq i \leq j$ , occurs with probability  $2^{-(j-1)} \binom{j-1}{i-1}$ , we set up

$$\sum_{i=1}^j 2^{-(j-1)} \binom{j-1}{i-1} r_j v_j^{i-1} = f_j.$$

- Thus

$$r_j = \left( \frac{2}{1 + v_j} \right)^{j-1} f_j. \quad (140)$$

- This binomial interest rate tree is trivial to set up (implicitly), in  $O(n)$  time.

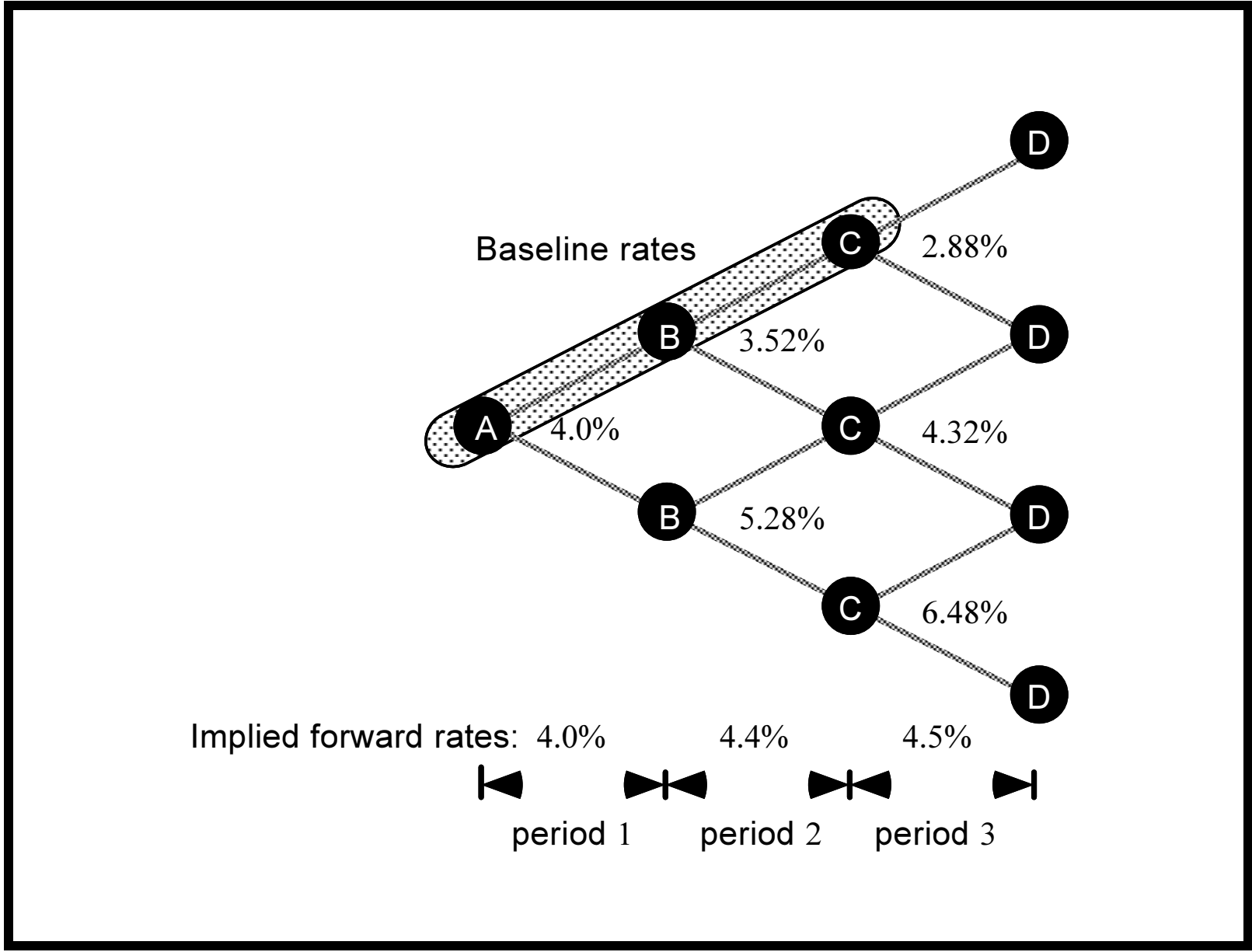
## An Approximate Calibration Scheme (continued)

- The ensuing tree for the sample term structure appears in figure on the next page.
- For example, the price of the zero-coupon bond paying \$1 at the end of the third period is

$$\frac{1}{4} \times \frac{1}{1.04} \times \left( \frac{1}{1.0352} \times \left( \frac{1}{1.0288} + \frac{1}{1.0432} \right) + \frac{1}{1.0528} \times \left( \frac{1}{1.0432} + \frac{1}{1.0648} \right) \right)$$

or 0.88155, which exceeds discount factor 0.88135.

- The tree is *not* calibrated.



## An Approximate Calibration Scheme (concluded)

- This bias is inherent: The tree *overprices* the bonds.<sup>a</sup>
- Suppose we replace the baseline rates  $r_j$  by  $r_j v_j$ .
- Then the resulting tree *underprices* the bonds.<sup>b</sup>
- The true baseline rates are thus bounded between  $r_j$  and  $r_j v_j$ .

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<sup>a</sup>See Exercise 23.2.4 in text.

<sup>b</sup>Lyu & C. Wang (F95922018) (2009, 2011).