

The Barrier-Too-Close Problem (p. 719) Revisited

- Our idea solves it even if one barrier is very close to S .
 - It runs in linear time, unlike an earlier quadratic-time solution with trinomial trees (pp. 727ff).
 - Unlike an earlier solution using combinatorics (p. 710), now the choice of n is not that restricted.
- So it combines the strengths of binomial and trinomial trees.
- This holds for single-barrier options too.
- Here is how.

The Barrier-Too-Close Problem Revisited (continued)

- We can build the tree treating S as if it were a second barrier.
- So both H and S are matched.
- Alternatively, we can pick $\Delta\tau \triangleq T/m$ as our length of a period Δt without the subsequent adjustment.^a
- Then build the tree from the price H down.
- So H is matched.
- As before, the initial price S will be matched by the *trinomial* structure.

^aThere is no second barrier to match!

The Barrier-Too-Close Problem Revisited (concluded)

- The earlier trinomial tree is impractical as it needs a very large n when the barrier H is very close to S .^a
 - It needs at least one *up* move to connect S to H as its middle branch is flat.
 - But when $S \approx H$, that up move must take a very small step, necessitating a small Δt .
- Our trinomial structure's middle branch is *not* necessarily flat.
- So S can be connected to H via the middle branch, and the need of a very large n no longer exists!

^aRecall the table on p. 720.

Pricing Discrete Barrier Options

- Barrier options whose barrier is monitored only at discrete times are called discrete barrier options.
- They are less common than the continuously monitored versions for single stocks.^a
- The main difficulty with pricing discrete barrier options lies in matching the monitored *times*.
- Here is why.

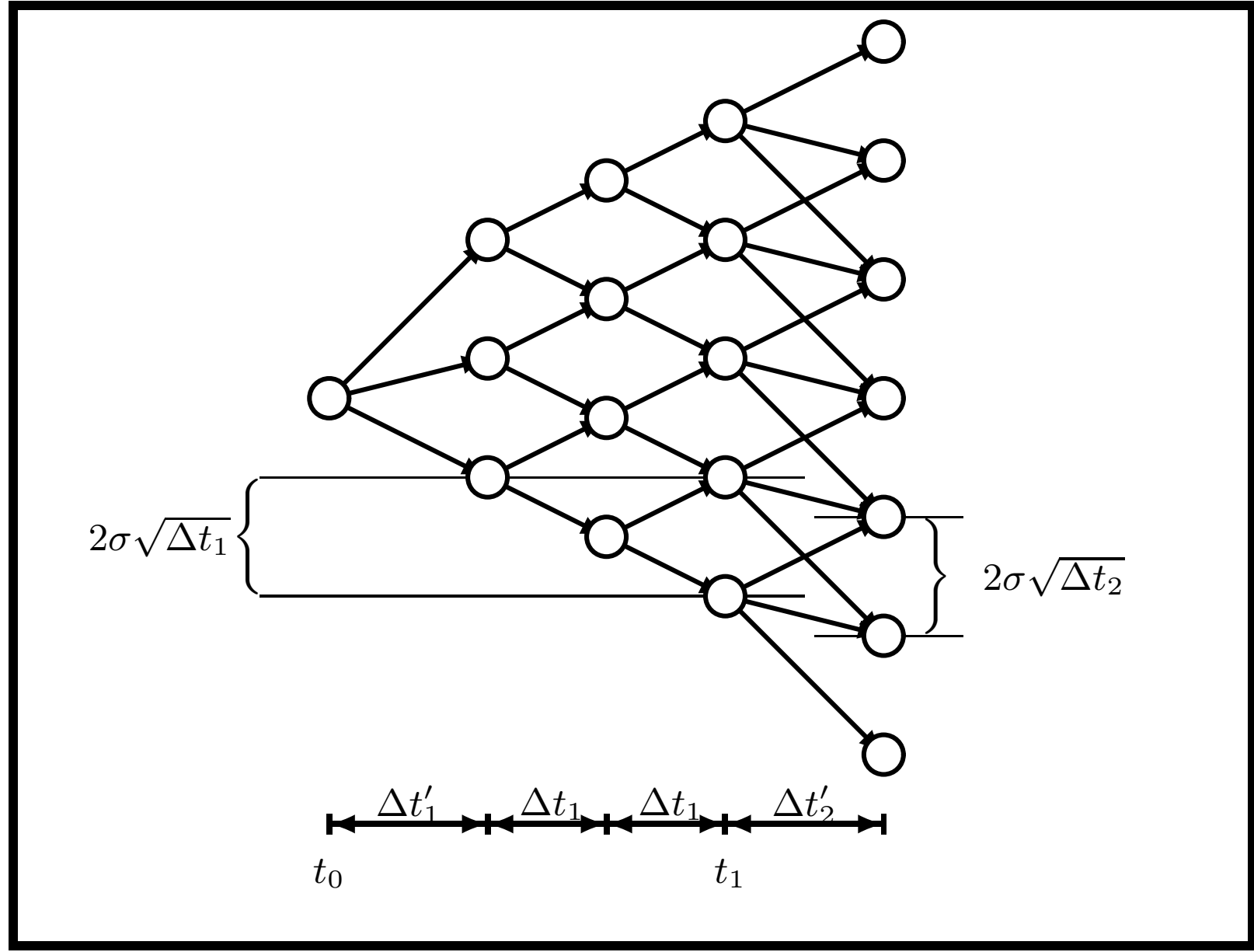
^aBennett (2014).

Pricing Discrete Barrier Options (continued)

- Suppose each period has a duration of Δt and the $\ell > 1$ monitored times are

$$t_0 = 0, t_1, t_2, \dots, t_\ell = T.$$

- It is unlikely that *all* monitored times coincide with the end of a period on the tree, or Δt divides t_i for all i .
- The binomial-trinomial tree can handle discrete options with ease, however.
- Simply build a binomial-trinomial tree from time 0 to time t_1 , followed by one from time t_1 to time t_2 , and so on until time t_ℓ .



Pricing Discrete Barrier Options (concluded)

- This procedure works even if each t_i is associated with a distinct barrier or if each window $[t_i, t_{i+1})$ has its own continuously monitored barrier or double barriers.
- For typical discrete barriers, placing barriers midway between two price levels on the tree may increase accuracy.^a

^aSteiner & Wallmeier (1999); Tavella & Randall (2000).

Options on a Stock That Pays Known Dividends

- Many ad hoc assumptions have been postulated for option pricing with known dividends.^a
 1. The one we saw earlier^b models the stock price minus the present value of the anticipated dividends as following geometric Brownian motion.
 2. One can also model the stock price plus the forward values of the dividends as following geometric Brownian motion.

^aFrishling (2002).

^bOn p. 318.

Options on a Stock That Pays Known Dividends (continued)

- Realistic models assume:
 - The stock price decreases by the amount of the dividend paid at the ex-dividend date.
 - The dividend is part cash and part yield (i.e., $\alpha(t)S_0 + \beta(t)S_t$), for practitioners.^a
- The stock price follows geometric Brownian motion between adjacent ex-dividend dates.
- But they result in binomial trees that grow exponentially (recall p. 317).
- The binomial-trinomial tree can avoid this problem in most cases.

^aHenry-Labordère (2009).

Options on a Stock That Pays Known Dividends (continued)

- Suppose that the known dividend is D dollars and the ex-dividend date is at time t .
- So there are $m = t/\Delta t$ periods between time 0 and the ex-dividend date.^a
- To avoid negative stock prices, we need to make sure the lowest stock price at time t is at least D , i.e.,

$$S e^{-(t/\Delta t)\sigma\sqrt{\Delta t}} \geq D.$$

– Or,

$$\Delta t \geq \left[\frac{t\sigma}{\ln(S/D)} \right]^2.$$

^aThat is, m is an integer input and $\Delta t \triangleq t/m$.

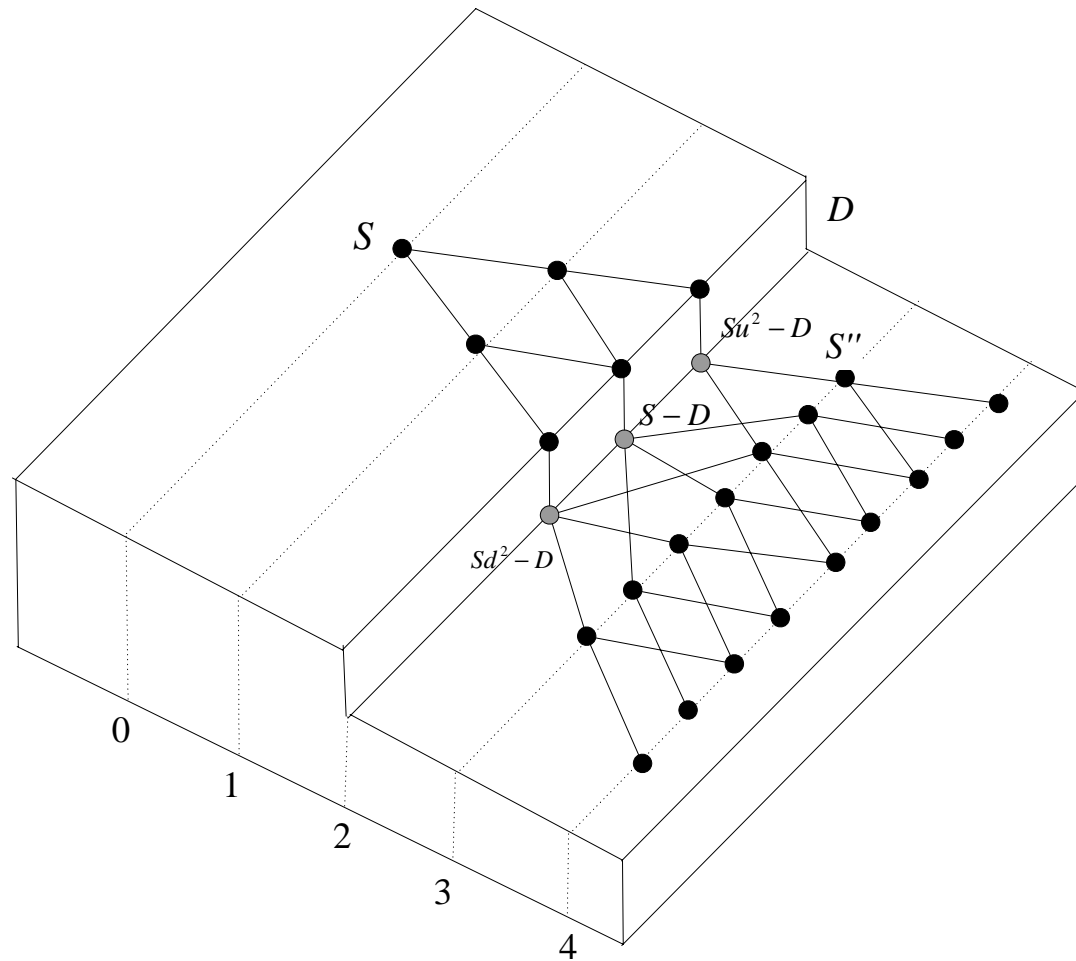
Options on a Stock That Pays Known Dividends (continued)

- Build a CRR tree from time 0 to time t as before.
- Subtract D from all the stock prices on the tree at time t to represent the price drop on the ex-dividend date.
- Assume the top node's price equals S' .
 - As usual, its two successor nodes will have prices $S'u$ and $S'u^{-1}$.
- The remaining nodes' successor nodes will choose from prices

$$S'u, S', S'u^{-1}, S'u^{-2}, S'u^{-3}, \dots,$$

same as the CRR tree.

A Stair Tree



Options on a Stock That Pays Known Dividends (continued)

- For each node at time t below the top node, we build the trinomial connection.
- Note that the binomial-trinomial structure remains valid in the special case when $\Delta t' = \Delta t$ on p. 746.

Options on a Stock That Pays Known Dividends (concluded)

- Hence the construction can be completed.
- From time $t + \Delta t$ onward, the standard binomial tree will be used until the maturity date or the next ex-dividend date when the procedure can be repeated.
- The resulting tree is called the stair tree.^a

^aDai (B82506025, R86526008, D8852600) & Lyuu (2004); Dai (B82506025, R86526008, D8852600) (2009).

Other Applications of Binomial-Trinomial Trees

- Pricing guaranteed minimum withdrawal benefits.^a
- Option pricing with stochastic volatilities.^b
- Efficient Parisian option pricing.^c
- Option pricing with time-varying volatilities and time-varying double barriers of an exponential form.^d
- Defaultable bond pricing.^e

^aH. Wu (R96723058) (2009).

^bC. Huang (R97922073) (2010).

^cY. Huang (R97922081) (2010).

^dC. Chou (R97944012) (2010); C. I. Chen (R98922127) (2011).

^eDai (B82506025, R86526008, D8852600), Lyuu, & C. Wang (F95922018) (2009, 2010, 2014).

Mean Tracking^a

- The general idea behind the binomial-trinomial tree on pp. 745ff is very powerful.
- One finds the successor middle node as the one closest to the mean.
- The two flanking successor nodes are then spaced at $c\sigma\sqrt{\Delta t}$ from the middle node for a suitably large $c > 0$.
- The resulting trinomial structure are then guaranteed to have valid branching probabilities.

^aLyyu & C. Wu (R90723065) (2003, 2005).

Default Boundary as Implied Barrier

- Under the structural model,^a the default boundary is modeled as a barrier.^b
- The constant barrier can be inferred from the closed-form formula given the firm's market capitalization, etc.^c
- More generally, the moving barrier can be inferred from the term structure of default probabilities with the binomial-trinomial tree.^d

^aRecall p. 366.

^bBlack & Cox (1976).

^cBrockman & Turtle (2003).

^dY. Lu (R06723032) (2019).

Time-Varying Double Barriers under Time-Dependent Volatility^a

- More general models allow a time-varying $\sigma(t)$ (p. 309).
- Let the two barriers $L(t)$ and $H(t)$ be functions of time.^b
- They do not have to be differentiable or even continuous.
- Still, we can price double-barrier options in $O(n^2)$ time or less with trinomial trees.
- Continuously monitored double-barrier knock-out options with time-varying barriers are called hot dog options.^c

^aY. Zhang (R05922052) (2019).

^bSo the barriers are continuously monitored.

^cEl Babsiri & Noel (1998).

General Local-Volatility Models and Their Trees

- Consider the general local-volatility model

$$\frac{dS}{S} = (r_t - q_t) dt + \sigma(S, t) dW,$$

where $a \leq \sigma(S, t) \leq b$ for some $a, b > 0$.

- This model has a unique (weak) solution.^a
- The positive lower bound is justifiable because prices fluctuate.

^aAchdou & Pironneau (2005).

General Local-Volatility Models and Their Trees (continued)

- The upper-bound assumption is also reasonable.
- Even on October 19, 1987, the CBOE S&P 100 Volatility Index (VXO) was about 150%, the highest ever.^a
- An efficient quadratic-sized tree for this range-bounded model is straightforward.^b
- Pick any $\sigma' \geq b$.
- Grow the trinomial tree with the node spacing $\sigma' \sqrt{\Delta t}$.^c
- The branching probabilities are guaranteed to be valid.

^aCaprio (2012).

^bLok (D99922028) & Lyuu (2016, 2017, 2020).

^cHaahtela (2010).

General Local-Volatility Models and Their Trees (concluded)

- The same idea can be applied to price double-barrier options.
- Pick any $\sigma' \geq \max[\max_{S, 0 \leq t \leq T} \sigma(S, t), \sqrt{2} \sigma(S_0, 0)]$.
- Grow the trinomial tree with the node spacing $\sigma' \sqrt{\Delta t}$.
- Apply the mean-tracking idea to the first period and Eqs. (97)–(102) on p. 751 to obtain the probabilities

Merton's Jump-Diffusion Model

- Empirically, stock returns tend to have fat tails, inconsistent with the Black-Scholes model's assumptions.
- Stochastic volatility and jump processes have been proposed to address this problem.
- Merton's (1976) jump-diffusion model is our focus.

Merton's Jump-Diffusion Model (continued)

- This model superimposes a jump component on a diffusion component.
- The diffusion component is the familiar geometric Brownian motion.
- The jump component is composed of lognormal jumps driven by a Poisson process.
 - It models the rare but large changes in the stock price because of the arrival of important new information.^a

^aDerman & M. B. Miller (2016), “There is no precise, universally accepted definition of a jump, but it usually comes down to magnitude, duration, and frequency.”

Merton's Jump-Diffusion Model (continued)

- Let S_t be the stock price at time t .
- The risk-neutral jump-diffusion process for the stock price follows^a

$$\frac{dS_t}{S_t} = (r - \lambda \bar{k}) dt + \sigma dW_t + k dq_t. \quad (104)$$

- Above, σ denotes the volatility of the diffusion component.

^aDerman & M. B. Miller (2016), “[M]ost jump-diffusion models simply assume risk-neutral pricing without convincing justification.”

Merton's Jump-Diffusion Model (continued)

- The jump event is governed by a compound Poisson process q_t with intensity λ , where k denotes the magnitude of the *random* jump.
 - The distribution of k obeys

$$\ln(1 + k) \sim N(\gamma, \delta^2)$$

with mean $\bar{k} \triangleq E(k) = e^{\gamma + \delta^2/2} - 1$.

- Note that $k > -1$.
 - Note also that k is not related to dt .
- The model with $\lambda = 0$ reduces to the Black-Scholes model.

Merton's Jump-Diffusion Model (continued)

- The solution to Eq. (104) on p. 786 is

$$S_t = S_0 e^{(r - \lambda \bar{k} - \sigma^2/2)t + \sigma W_t} U(n(t)), \quad (105)$$

where

$$U(n(t)) = \prod_{i=0}^{n(t)} (1 + k_i).$$

- k_i is the magnitude of the i th jump with $\ln(1 + k_i) \sim N(\gamma, \delta^2)$.
- $k_0 = 0$.
- $n(t)$ is a Poisson process with intensity λ .

Merton's Jump-Diffusion Model (concluded)

- Recall that $n(t)$ denotes the number of jumps that occur up to time t .
- It is known that $E[n(t)] = \text{Var}[n(t)] = \lambda t$.
- As $k_i > -1$, stock prices will stay positive.
- The geometric Brownian motion, the lognormal jumps, and the Poisson process are assumed to be independent.

Tree for Merton's Jump-Diffusion Model^a

- Define the S -logarithmic return of the stock price S' as

$$\ln(S'/S).$$

- Define the logarithmic distance between stock prices S' and S as

$$| \ln(S') - \ln(S) | = | \ln(S'/S) |.$$

^aDai (B82506025, R86526008, D8852600), C. Wang (F95922018), Lyuu, & Y. Liu (2010).

Tree for Merton's Jump-Diffusion Model (continued)

- Take the logarithm of Eq. (105) on p. 788:

$$M_t \triangleq \ln \left(\frac{S_t}{S_0} \right) = X_t + Y_t, \quad (106)$$

where

$$X_t \triangleq \left(r - \lambda \bar{k} - \frac{\sigma^2}{2} \right) t + \sigma W_t, \quad (107)$$

$$Y_t \triangleq \sum_{i=0}^{n(t)} \ln (1 + k_i). \quad (108)$$

- It decomposes the S_0 -logarithmic return of S_t into the diffusion component X_t and the jump component Y_t .

Tree for Merton's Jump-Diffusion Model (continued)

- Motivated by decomposition (106) on p. 791, the tree construction divides each period into a diffusion phase followed by a jump phase.
- In the diffusion phase, X_t is approximated by the BOPM.
- So X_t makes an up move to $X_t + \sigma\sqrt{\Delta t}$ with probability p_u or a down move to $X_t - \sigma\sqrt{\Delta t}$ with probability p_d .

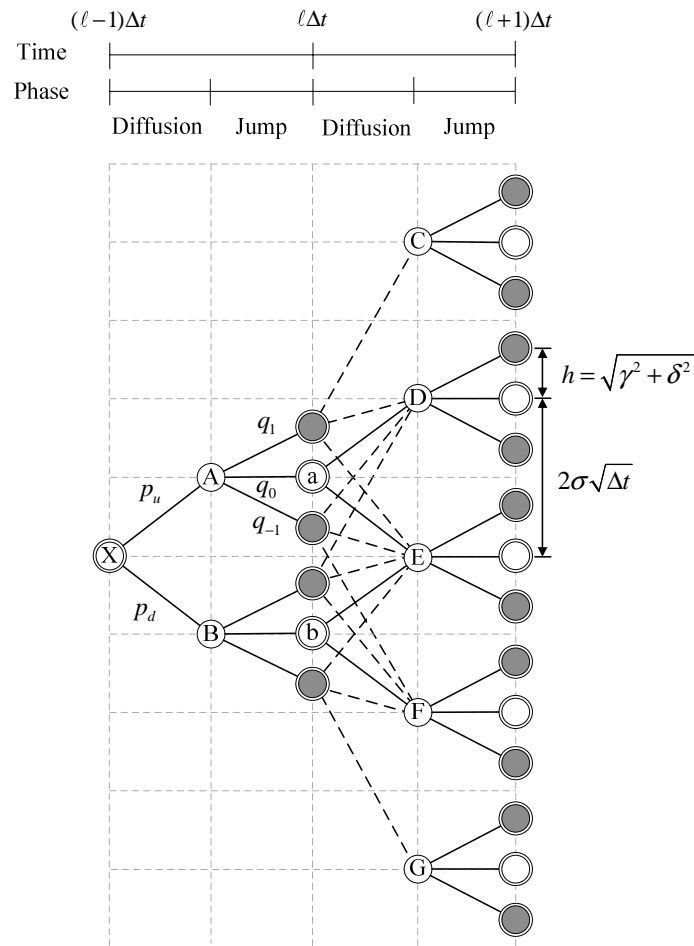
Tree for Merton's Jump-Diffusion Model (continued)

- According to BOPM,

$$\begin{aligned}p_u &= \frac{e^{\mu\Delta t} - d}{u - d}, \\p_d &= 1 - p_u,\end{aligned}$$

except that $\mu = r - \lambda\bar{k}$ here.

- The diffusion component gives rise to diffusion nodes.
- They are spaced at $2\sigma\sqrt{\Delta t}$ apart such as the white nodes A, B, C, D, E, F, and G on p. 794.



White nodes are *diffusion nodes*. Gray nodes are *jump nodes*. In the diffusion phase, the solid black lines denote the binomial structure of BOPM; the dashed lines denote the trinomial structure. Only the double-circled nodes will remain after the construction. Note that a and b are diffusion nodes because no jump occurs in the jump phase.

Tree for Merton's Jump-Diffusion Model (continued)

- In the jump phase, $Y_{t+\Delta t}$ is approximated by moves from *each* diffusion node to $2m$ jump nodes that match the first $2m$ moments of the lognormal jump.
- The m jump nodes above the diffusion node are spaced at $h \triangleq \sqrt{\gamma^2 + \delta^2}$ apart.
- Note that h is independent of Δt .

Tree for Merton's Jump-Diffusion Model (concluded)

- The same holds for the m jump nodes below the diffusion node.
- The gray nodes at time $\ell\Delta t$ on p. 794 are jump nodes.
 - We set $m = 1$ on p. 794.
- The size of the tree is $O(n^{2.5})$.

Multivariate Contingent Claims

- They depend on two or more underlying assets.
- The basket call on m assets has the terminal payoff

$$\max \left(\sum_{i=1}^m \alpha_i S_i(\tau) - X, 0 \right),$$

where α_i is the percentage of asset i .

- Basket options are essentially options on a portfolio of stocks (or index options).^a
- Option on the best of two risky assets and cash has a terminal payoff of $\max(S_1(\tau), S_2(\tau), X)$.

^aExcept that membership and weights do *not* change for basket options (Bennett, 2014).

Multivariate Contingent Claims (concluded)^a

| Name | Payoff |
|------------------------|---|
| Exchange option | $\max(S_1(\tau) - S_2(\tau), 0)$ |
| Better-off option | $\max(S_1(\tau), \dots, S_k(\tau), 0)$ |
| Worst-off option | $\min(S_1(\tau), \dots, S_k(\tau), 0)$ |
| Binary maximum option | $I\{\max(S_1(\tau), \dots, S_k(\tau)) > X\}$ |
| Maximum option | $\max(\max(S_1(\tau), \dots, S_k(\tau)) - X, 0)$ |
| Minimum option | $\max(\min(S_1(\tau), \dots, S_k(\tau)) - X, 0)$ |
| Spread option | $\max(S_1(\tau) - S_2(\tau) - X, 0)$ |
| Basket average option | $\max((S_1(\tau) + \dots + S_k(\tau))/k - X, 0)$ |
| Multi-strike option | $\max(S_1(\tau) - X_1, \dots, S_k(\tau) - X_k, 0)$ |
| Pyramid rainbow option | $\max(S_1(\tau) - X_1 + \dots + S_k(\tau) - X_k - X, 0)$ |
| Madonna option | $\max(\sqrt{(S_1(\tau) - X_1)^2 + \dots + (S_k(\tau) - X_k)^2} - X, 0)$ |

^aLyuu & Teng (R91723054) (2011).

Correlated Trinomial Model^a

- Two risky assets S_1 and S_2 follow

$$\frac{dS_i}{S_i} = r dt + \sigma_i dW_i$$

in a risk-neutral economy, $i = 1, 2$.

- Let

$$\begin{aligned} M_i &\triangleq e^{r\Delta t}, \\ V_i &\triangleq M_i^2(e^{\sigma_i^2\Delta t} - 1). \end{aligned}$$

- $S_i M_i$ is the mean of S_i at time Δt .
- $S_i^2 V_i$ the variance of S_i at time Δt .

^aBoyle, Evnine, & Gibbs (1989).

Correlated Trinomial Model (continued)

- The value of $S_1 S_2$ at time Δt has a joint lognormal distribution with mean $S_1 S_2 M_1 M_2 e^{\rho \sigma_1 \sigma_2 \Delta t}$, where ρ is the correlation between dW_1 and dW_2 .
- Next match the 1st and 2nd moments of the approximating discrete distribution to those of the continuous counterpart.
- At time Δt from now, there are 5 distinct outcomes.

Correlated Trinomial Model (continued)

- The five-point probability distribution of the asset prices is

| Probability | Asset 1 | Asset 2 |
|-------------|-----------|-----------|
| p_1 | $S_1 u_1$ | $S_2 u_2$ |
| p_2 | $S_1 u_1$ | $S_2 d_2$ |
| p_3 | $S_1 d_1$ | $S_2 d_2$ |
| p_4 | $S_1 d_1$ | $S_2 u_2$ |
| p_5 | S_1 | S_2 |

- As usual, impose $u_i d_i = 1$.

Correlated Trinomial Model (continued)

- The probabilities must sum to one, and the means must be matched:

$$1 = p_1 + p_2 + p_3 + p_4 + p_5,$$

$$S_1 M_1 = (p_1 + p_2) S_1 u_1 + p_5 S_1 + (p_3 + p_4) S_1 d_1,$$

$$S_2 M_2 = (p_1 + p_4) S_2 u_2 + p_5 S_2 + (p_2 + p_3) S_2 d_2.$$

Correlated Trinomial Model (concluded)

- Let $R \triangleq M_1 M_2 e^{\rho \sigma_1 \sigma_2 \Delta t}$.
- Match the variances and covariance:

$$\begin{aligned} S_1^2 V_1 &= (p_1 + p_2)((S_1 u_1)^2 - (S_1 M_1)^2) + p_5(S_1^2 - (S_1 M_1)^2) \\ &\quad + (p_3 + p_4)((S_1 d_1)^2 - (S_1 M_1)^2), \end{aligned}$$

$$\begin{aligned} S_2^2 V_2 &= (p_1 + p_4)((S_2 u_2)^2 - (S_2 M_2)^2) + p_5(S_2^2 - (S_2 M_2)^2) \\ &\quad + (p_2 + p_3)((S_2 d_2)^2 - (S_2 M_2)^2), \end{aligned}$$

$$S_1 S_2 R = (p_1 u_1 u_2 + p_2 u_1 d_2 + p_3 d_1 d_2 + p_4 d_1 u_2 + p_5) S_1 S_2.$$

- The solutions appear on p. 246 of the textbook.

Correlated Trinomial Model Simplified^a

- Let $\mu'_i \triangleq r - \sigma_i^2/2$ and $u_i \triangleq e^{\lambda\sigma_i\sqrt{\Delta t}}$ for $i = 1, 2$.
- The following simpler scheme is good enough:

$$\begin{aligned}p_1 &= \frac{1}{4} \left[\frac{1}{\lambda^2} + \frac{\sqrt{\Delta t}}{\lambda} \left(\frac{\mu'_1}{\sigma_1} + \frac{\mu'_2}{\sigma_2} \right) + \frac{\rho}{\lambda^2} \right], \\p_2 &= \frac{1}{4} \left[\frac{1}{\lambda^2} + \frac{\sqrt{\Delta t}}{\lambda} \left(\frac{\mu'_1}{\sigma_1} - \frac{\mu'_2}{\sigma_2} \right) - \frac{\rho}{\lambda^2} \right], \\p_3 &= \frac{1}{4} \left[\frac{1}{\lambda^2} + \frac{\sqrt{\Delta t}}{\lambda} \left(-\frac{\mu'_1}{\sigma_1} - \frac{\mu'_2}{\sigma_2} \right) + \frac{\rho}{\lambda^2} \right], \\p_4 &= \frac{1}{4} \left[\frac{1}{\lambda^2} + \frac{\sqrt{\Delta t}}{\lambda} \left(-\frac{\mu'_1}{\sigma_1} + \frac{\mu'_2}{\sigma_2} \right) - \frac{\rho}{\lambda^2} \right], \\p_5 &= 1 - \frac{1}{\lambda^2}.\end{aligned}$$

^aMadan, Milne, & Shefrin (1989).

Correlated Trinomial Model Simplified (continued)

- All of the probabilities lie between 0 and 1 if and only if

$$-1 + \lambda\sqrt{\Delta t} \left| \frac{\mu'_1}{\sigma_1} + \frac{\mu'_2}{\sigma_2} \right| \leq \rho \leq 1 - \lambda\sqrt{\Delta t} \left| \frac{\mu'_1}{\sigma_1} - \frac{\mu'_2}{\sigma_2} \right| \quad (109)$$

$$1 \leq \lambda \quad (110)$$

- We call a multivariate tree (correlation-) optimal if it guarantees valid probabilities as long as

$$-1 + O(\sqrt{\Delta t}) < \rho < 1 - O(\sqrt{\Delta t}),$$

such as the above one.^a

^aW. Kao (R98922093) (2011); W. Kao (R98922093), Lyuu, & Wen (D94922003) (2014).

Correlated Trinomial Model Simplified (continued)

- But this model cannot price 2-asset 2-barrier options accurately.^a
- Few multivariate trees are both optimal and able to handle multiple barriers.^b
- An alternative is to use orthogonalization.^c

^aSee Y. Chang (B89704039, R93922034), Hsu (R7526001, D89922012), & Lyuu (2006); W. Kao (R98922093), Lyuu, & Wen (D94922003) (2014) for solutions.

^bSee W. Kao (R98922093), Lyuu, & Wen (D94922003) (2014) for one.

^cHull & White (1990); Dai (B82506025, R86526008, D8852600), C. Wang (F95922018), & Lyuu (2013).

Correlated Trinomial Model Simplified (concluded)

- Suppose we allow each asset's volatility to be a function of time.^a
- There are k assets.
- Can you build an optimal multivariate tree that can handle two barriers on each asset in time $O(n^{k+1})$?^b

^aRecall p. 308.

^bSee Y. Zhang (R05922052) (2019) for a complete solution.

Extrapolation

- It is a method to speed up numerical convergence.
- Say $f(n)$ converges to an unknown limit f at rate of $1/n$:

$$f(n) = f + \frac{c}{n} + o\left(\frac{1}{n}\right). \quad (111)$$

- Assume c is an unknown constant independent of n .
 - Convergence is basically monotonic and smooth.

Extrapolation (concluded)

- From two approximations $f(n_1)$ and $f(n_2)$ and ignoring the smaller terms,

$$f(n_1) = f + \frac{c}{n_1},$$

$$f(n_2) = f + \frac{c}{n_2}.$$

- A better approximation to the desired f is

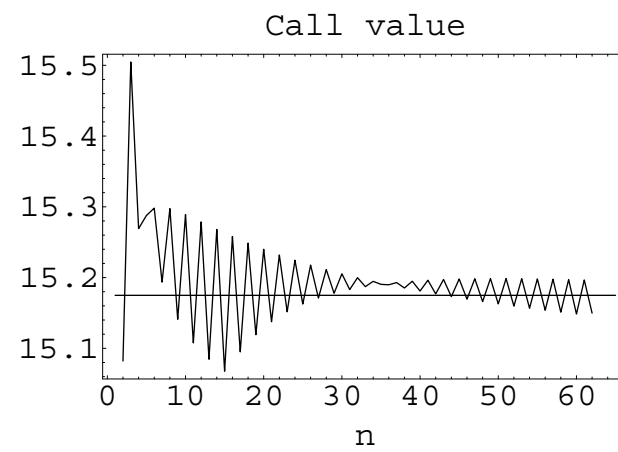
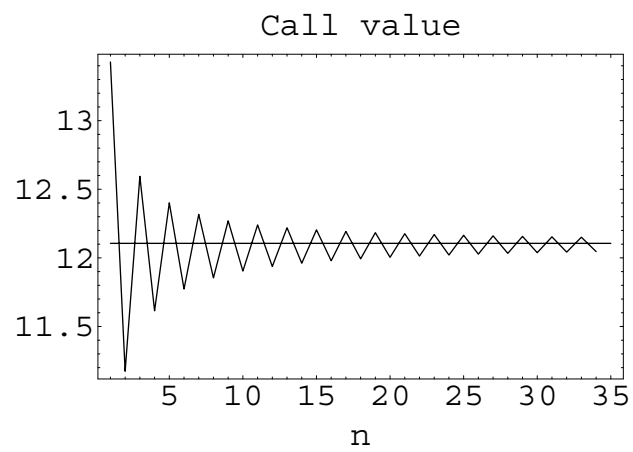
$$f = \frac{n_1 f(n_1) - n_2 f(n_2)}{n_1 - n_2}. \quad (112)$$

- This estimate should converge faster than $1/n$.^a
- The Richardson extrapolation uses $n_2 = 2n_1$.

^aIt is identical to the forward rate formula (23) on p. 150!

Improving BOPM with Extrapolation

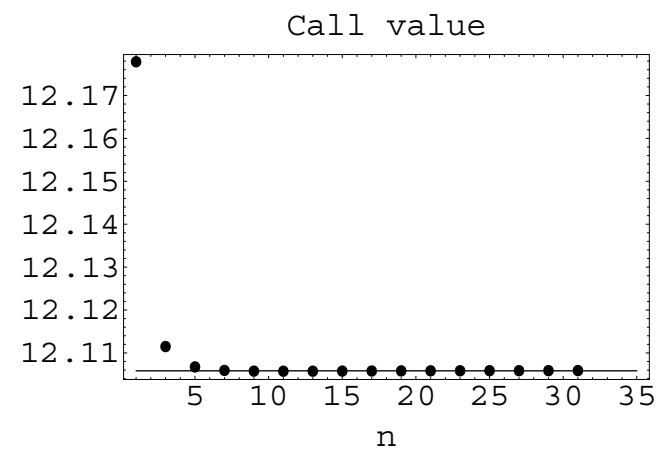
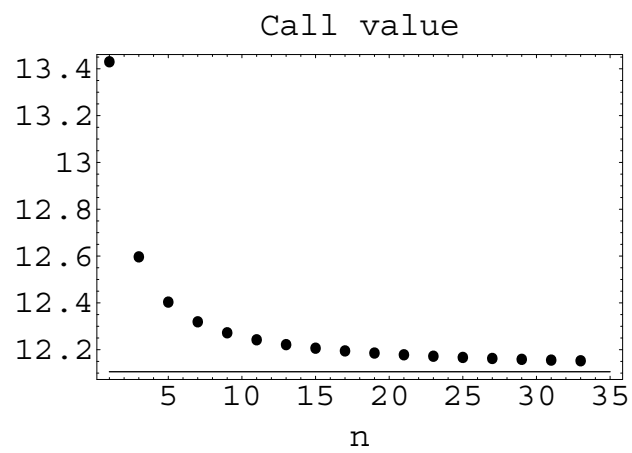
- Consider standard European options.
- Denote the option value under BOPM using n time periods by $f(n)$.
- It is known that BOPM converges at the rate of $1/n$, consistent with Eq. (111) on p. 808.
- The plots on p. 299 (redrawn on next page) show that convergence to the true option value oscillates with n .
- Extrapolation is inapplicable at this stage.



Improving BOPM with Extrapolation (concluded)

- Take the at-the-money option in the left plot on p. 811.
- The sequence with odd n turns out to be monotonic and smooth (see the left plot on p. 813).^a
- Apply extrapolation (112) on p. 809 with $n_2 = n_1 + 2$, where n_1 is odd.
- Result is shown in the right plot on p. 813.
- The convergence rate is amazing.
- See Exercise 9.3.8 of the text (p. 111) for ideas in the general case.

^aThis can be proved (L. Chang & Palmer, 2007).

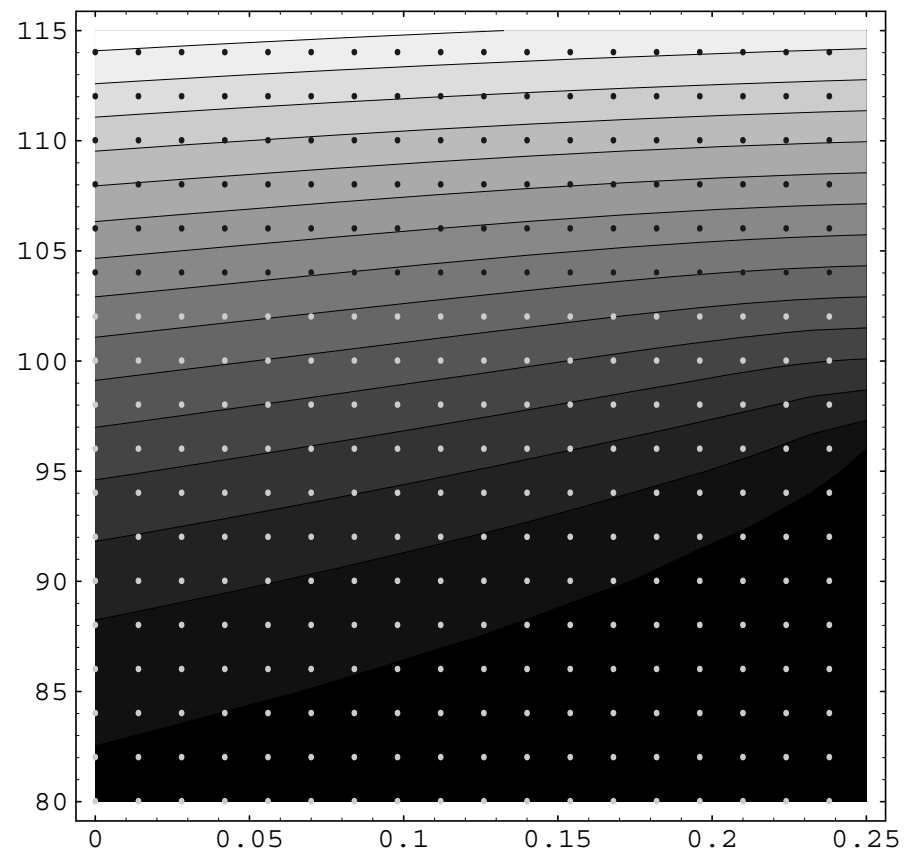


Numerical Methods

All science is dominated
by the idea of approximation.
— Bertrand Russell

Finite-Difference Methods

- Place a grid of points on the space over which the desired function takes value.
- Then approximate the function value at each of these points (p. 817).
- Solve the equation numerically by introducing difference equations in place of derivatives.



Example: Poisson's Equation

- It is $\partial^2\theta/\partial x^2 + \partial^2\theta/\partial y^2 = -\rho(x, y)$, which describes the electrostatic field.
- Replace second derivatives with finite differences through central difference.
- Introduce evenly spaced grid points with distance of Δx along the x axis and Δy along the y axis.
- The finite difference form is

$$-\rho(x_i, y_j) = \frac{\theta(x_{i+1}, y_j) - 2\theta(x_i, y_j) + \theta(x_{i-1}, y_j)}{(\Delta x)^2} + \frac{\theta(x_i, y_{j+1}) - 2\theta(x_i, y_j) + \theta(x_i, y_{j-1})}{(\Delta y)^2}.$$

Example: Poisson's Equation (concluded)

- In the above, $\Delta x \triangleq x_i - x_{i-1}$ and $\Delta y \triangleq y_j - y_{j-1}$ for $i, j = 1, 2, \dots$
- When the grid points are evenly spaced in both axes so that $\Delta x = \Delta y = h$, the difference equation becomes

$$\begin{aligned} -h^2 \rho(x_i, y_j) = & \theta(x_{i+1}, y_j) + \theta(x_{i-1}, y_j) \\ & + \theta(x_i, y_{j+1}) + \theta(x_i, y_{j-1}) - 4\theta(x_i, y_j). \end{aligned}$$

- Given boundary values, we can solve for the x_i s and the y_j s within the square $[\pm L, \pm L]$.
- From now on, $\theta_{i,j}$ will denote the finite-difference approximation to the exact $\theta(x_i, y_j)$.

Explicit Methods

- Consider the diffusion equation
 $D(\partial^2\theta/\partial x^2) - (\partial\theta/\partial t) = 0, D > 0.$
- Use evenly spaced grid points (x_i, t_j) with distances Δx and Δt , where $\Delta x \triangleq x_{i+1} - x_i$ and $\Delta t \triangleq t_{j+1} - t_j$.
- Employ central difference for the second derivative and forward difference for the time derivative to obtain

$$\left. \frac{\partial\theta(x, t)}{\partial t} \right|_{t=t_j} = \frac{\theta(x, t_{j+1}) - \theta(x, t_j)}{\Delta t} + \dots, \quad (113)$$

$$\left. \frac{\partial^2\theta(x, t)}{\partial x^2} \right|_{x=x_i} = \frac{\theta(x_{i+1}, t) - 2\theta(x_i, t) + \theta(x_{i-1}, t))}{(\Delta x)^2} + \dots \quad (114)$$

Explicit Methods (continued)

- Next, assemble Eqs. (113) and (114) into a single equation at (x_i, t_j) .
- But we need to decide how to evaluate x in the first equation and t in the second.
- Since central difference around x_i is used in Eq. (114), we might as well use x_i for x in Eq. (113).
- Two choices are possible for t in Eq. (114).
- The first choice uses $t = t_j$ to yield the following finite-difference equation,

$$\frac{\theta_{i,j+1} - \theta_{i,j}}{\Delta t} = D \frac{\theta_{i+1,j} - 2\theta_{i,j} + \theta_{i-1,j}}{(\Delta x)^2}. \quad (115)$$

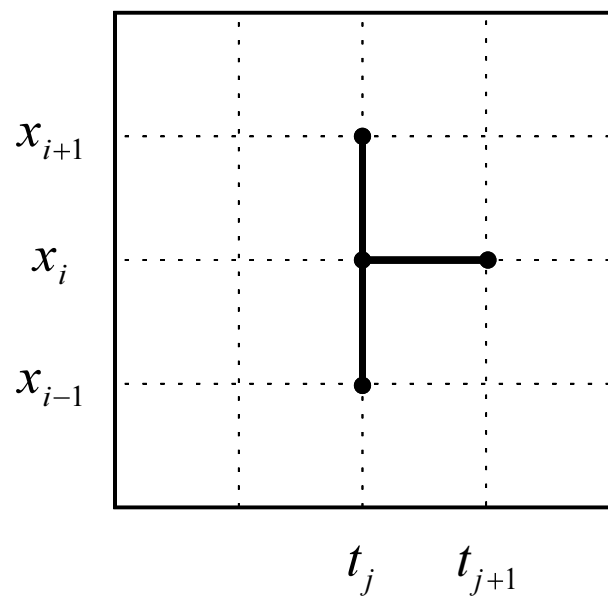
Explicit Methods (continued)

- The stencil of grid points involves four values, $\theta_{i,j+1}$, $\theta_{i,j}$, $\theta_{i+1,j}$, and $\theta_{i-1,j}$.
- Rearrange Eq. (115) on p. 821 as

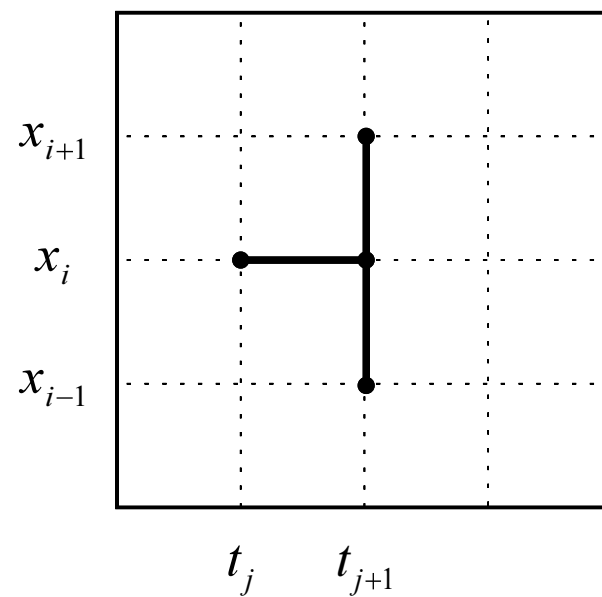
$$\theta_{i,j+1} = \frac{D\Delta t}{(\Delta x)^2} \theta_{i+1,j} + \left(1 - \frac{2D\Delta t}{(\Delta x)^2}\right) \theta_{i,j} + \frac{D\Delta t}{(\Delta x)^2} \theta_{i-1,j}.$$

- We can calculate $\theta_{i,j+1}$ from $\theta_{i,j}$, $\theta_{i+1,j}$, $\theta_{i-1,j}$, at the previous time t_j (see exhibit (a) on next page).

Stencils



(a)



(b)

Explicit Methods (concluded)

- Starting from the initial conditions at t_0 , that is, $\theta_{i,0} = \theta(x_i, t_0)$, $i = 1, 2, \dots$, we calculate

$$\theta_{i,1}, \quad i = 1, 2, \dots .$$

- And then

$$\theta_{i,2}, \quad i = 1, 2, \dots .$$

- And so on.

Stability

- The explicit method is numerically unstable unless

$$\Delta t \leq (\Delta x)^2 / (2D).$$

- A numerical method is unstable if the solution is highly sensitive to changes in initial conditions.
- The stability condition may lead to high running times and memory requirements.
- For instance, halving Δx would imply quadrupling $(\Delta t)^{-1}$, resulting in a running time 8 times as much.

Explicit Method and Trinomial Tree

- Recall that

$$\theta_{i,j+1} = \frac{D\Delta t}{(\Delta x)^2} \theta_{i+1,j} + \left(1 - \frac{2D\Delta t}{(\Delta x)^2}\right) \theta_{i,j} + \frac{D\Delta t}{(\Delta x)^2} \theta_{i-1,j}.$$

- When the stability condition is satisfied, the three coefficients for $\theta_{i+1,j}$, $\theta_{i,j}$, and $\theta_{i-1,j}$ all lie between zero and one and sum to one.
- They can be interpreted as probabilities.
- So the finite-difference equation becomes identical to backward induction on trinomial trees!

Explicit Method and Trinomial Tree (concluded)

- The freedom in choosing Δx corresponds to similar freedom in the construction of trinomial trees.
- The explicit finite-difference equation is also identical to backward induction on a binomial tree.^a
 - Let the binomial tree take 2 steps each of length $\Delta t/2$.
 - It is now a trinomial tree.

^aHilliard (2014).

Implicit Methods

- Suppose we use $t = t_{j+1}$ in Eq. (114) on p. 820 instead.
- The finite-difference equation becomes

$$\frac{\theta_{i,j+1} - \theta_{i,j}}{\Delta t} = D \frac{\theta_{i+1,j+1} - 2\theta_{i,j+1} + \theta_{i-1,j+1}}{(\Delta x)^2}. \quad (116)$$

- The stencil involves $\theta_{i,j}$, $\theta_{i,j+1}$, $\theta_{i+1,j+1}$, and $\theta_{i-1,j+1}$.
- This method is implicit:
 - The value of any one of the three quantities at t_{j+1} cannot be calculated unless the other two are known.
 - See exhibit (b) on p. 823.

Implicit Methods (continued)

- Equation (116) can be rearranged as

$$\theta_{i-1,j+1} - (2 + \gamma) \theta_{i,j+1} + \theta_{i+1,j+1} = -\gamma \theta_{i,j},$$

where $\gamma \triangleq (\Delta x)^2 / (D \Delta t)$.

- This equation is unconditionally stable.
- Suppose the boundary conditions are given at $x = x_0$ and $x = x_{N+1}$.
- After $\theta_{i,j}$ has been calculated for $i = 1, 2, \dots, N$, the values of $\theta_{i,j+1}$ at time t_{j+1} can be computed as the solution to the following tridiagonal linear system,

Implicit Methods (continued)

$$\begin{bmatrix}
 a & 1 & 0 & \cdots & \cdots & \cdots & 0 \\
 1 & a & 1 & 0 & \cdots & \cdots & 0 \\
 0 & 1 & a & 1 & 0 & \cdots & 0 \\
 \vdots & \ddots & \ddots & \ddots & \ddots & \ddots & \vdots \\
 \vdots & \ddots & \ddots & \ddots & \ddots & \ddots & \vdots \\
 0 & \cdots & \cdots & 0 & 1 & a & 1 \\
 0 & \cdots & \cdots & \cdots & 0 & 1 & a
 \end{bmatrix}
 \begin{bmatrix}
 \theta_{1,j+1} \\
 \theta_{2,j+1} \\
 \theta_{3,j+1} \\
 \vdots \\
 \vdots \\
 \vdots \\
 \vdots \\
 \theta_{N,j+1}
 \end{bmatrix}
 =
 \begin{bmatrix}
 -\gamma\theta_{1,j} - \theta_{0,j+1} \\
 -\gamma\theta_{2,j} \\
 -\gamma\theta_{3,j} \\
 \vdots \\
 \vdots \\
 \vdots \\
 -\gamma\theta_{N-1,j} \\
 -\gamma\theta_{N,j} - \theta_{N+1,j+1}
 \end{bmatrix},$$

where $a \triangleq -2 - \gamma$.

Implicit Methods (concluded)

- Tridiagonal systems can be solved in $O(N)$ time and $O(N)$ space.
 - Never invert a matrix to solve a tridiagonal system.
- The matrix above is nonsingular when $\gamma \geq 0$.
 - A square matrix is nonsingular if its inverse exists.

Crank-Nicolson Method

- Take the average of explicit method (115) on p. 821 and implicit method (116) on p. 828:

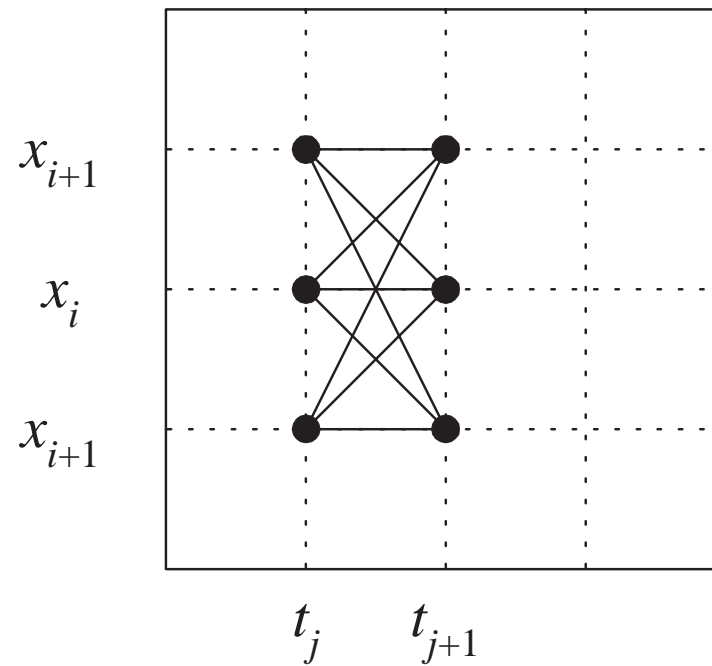
$$\begin{aligned} & \frac{\theta_{i,j+1} - \theta_{i,j}}{\Delta t} \\ = & \frac{1}{2} \left(D \frac{\theta_{i+1,j} - 2\theta_{i,j} + \theta_{i-1,j}}{(\Delta x)^2} + D \frac{\theta_{i+1,j+1} - 2\theta_{i,j+1} + \theta_{i-1,j+1}}{(\Delta x)^2} \right). \end{aligned}$$

- After rearrangement,

$$\gamma\theta_{i,j+1} - \frac{\theta_{i+1,j+1} - 2\theta_{i,j+1} + \theta_{i-1,j+1}}{2} = \gamma\theta_{i,j} + \frac{\theta_{i+1,j} - 2\theta_{i,j} + \theta_{i-1,j}}{2}.$$

- This is an unconditionally stable implicit method with excellent rates of convergence.

Stencil



Numerically Solving the Black-Scholes PDE (87) on p. 669

- See text.
- Brennan and Schwartz (1978) analyze the stability of the implicit method.

Monte Carlo Simulation^a

- Monte Carlo simulation is a sampling scheme.
- In many important applications within finance and without, Monte Carlo is one of the few feasible tools.
- When the time evolution of a stochastic process is not easy to describe analytically, Monte Carlo may very well be the only strategy that succeeds consistently.

^aA top 10 algorithm (Dongarra & Sullivan, 2000).

The Big Idea

- Assume X_1, X_2, \dots, X_n have a joint distribution.
- $\theta \triangleq E[g(X_1, X_2, \dots, X_n)]$ for some function g is desired.
- We generate

$$\left(x_1^{(i)}, x_2^{(i)}, \dots, x_n^{(i)}\right), \quad 1 \leq i \leq N$$

independently with the same joint distribution as (X_1, X_2, \dots, X_n) .

- Set

$$Y_i \triangleq g\left(x_1^{(i)}, x_2^{(i)}, \dots, x_n^{(i)}\right).$$

The Big Idea (concluded)

- Y_1, Y_2, \dots, Y_N are independent and identically distributed random variables.
- Each Y_i has the same distribution as

$$Y \triangleq g(X_1, X_2, \dots, X_n).$$

- Since the average of these N random variables, \bar{Y} , satisfies $E[\bar{Y}] = \theta$, it can be used to estimate θ .
- The strong law of large numbers says that this procedure converges almost surely.
- The number of replications (or independent trials), N , is called the sample size.

Accuracy

- The Monte Carlo estimate and true value may differ owing to two reasons:
 1. Sampling variation.
 2. The discreteness of the sample paths.^a
- The first can be controlled by the number of replications.
- The second can be controlled by the number of observations along the sample path.

^aThis may not be an issue if the financial derivative only requires discrete sampling along the time dimension, such as the *discrete* barrier option.

Accuracy and Number of Replications

- The statistical error of the sample mean \bar{Y} of the random variable Y grows as $1/\sqrt{N}$.
 - Because $\text{Var}[\bar{Y}] = \text{Var}[Y]/N$.
- In fact, this convergence rate is asymptotically optimal.^a
- So the variance of the estimator \bar{Y} can be reduced by a factor of $1/N$ by doing N times as much work.
- This is amazing because the same order of convergence holds independently of the dimension n .

^aThe Berry-Esseen theorem.

Accuracy and Number of Replications (concluded)

- In contrast, classic numerical integration schemes have an error bound of $O(N^{-c/n})$ for some constant $c > 0$.
 - n is the dimension.
- The required number of evaluations thus grows exponentially in n to achieve a given level of accuracy.
 - The curse of dimensionality.
- The Monte Carlo method is more efficient than alternative procedures for multivariate derivatives when n is large.

Monte Carlo Option Pricing

- For the pricing of European options on a dividend-paying stock, we may proceed as follows.

- Assume

$$\frac{dS}{S} = \mu dt + \sigma dW.$$

- Stock prices S_1, S_2, S_3, \dots at times $\Delta t, 2\Delta t, 3\Delta t, \dots$ can be generated via

$$S_{i+1} = S_i e^{(\mu - \sigma^2/2) \Delta t + \sigma \sqrt{\Delta t} \xi}, \quad \xi \sim N(0, 1). \quad (117)$$

Monte Carlo Option Pricing (continued)

- If we discretize $dS/S = \mu dt + \sigma dW$ directly, we will obtain

$$S_{i+1} = S_i + S_i \mu \Delta t + S_i \sigma \sqrt{\Delta t} \xi.$$

- But this is locally normally distributed, not lognormally, hence biased.^a
- In practice, this is not expected to be a major problem as long as Δt is sufficiently small.

^aContributed by Mr. Tai, Hui-Chin (R97723028) on April 22, 2009.

Monte Carlo Option Pricing (continued)

Non-dividend-paying stock prices in a risk-neutral economy can be generated by setting $\mu = r$ and $\Delta t = T$.

```
1:  $C := 0$ ; {Accumulated terminal option value.}
2: for  $i = 1, 2, 3, \dots, N$  do
3:    $P := S \times e^{(r - \sigma^2/2)T + \sigma\sqrt{T}\xi}$ ,  $\xi \sim N(0, 1)$ ;
4:    $C := C + \max(P - X, 0)$ ;
5: end for
6: return  $Ce^{-rT}/N$ ;
```

Monte Carlo Option Pricing (concluded)

Pricing Asian options is also easy.

```
1:  $C := 0$ ;  
2: for  $i = 1, 2, 3, \dots, N$  do  
3:    $P := S$ ;  $M := S$ ;  
4:   for  $j = 1, 2, 3, \dots, n$  do  
5:      $P := P \times e^{(r - \sigma^2/2)(T/n) + \sigma\sqrt{T/n} \xi}$ ;  
6:      $M := M + P$ ;  
7:   end for  
8:    $C := C + \max(M/(n+1) - X, 0)$ ;  
9: end for  
10: return  $Ce^{-rT}/N$ ;
```

How about American Options?

- Standard Monte Carlo simulation is inappropriate for American options because of early exercise.
 - Given a sample path S_0, S_1, \dots, S_n , how to decide which S_i is an early-exercise point?
 - What is the option price at each S_i if the option is not exercised?
- It is difficult to determine the early-exercise point based on one single path.
- But Monte Carlo simulation can be modified to price American options with small biases (pp. 906ff).^a

^aLongstaff & Schwartz (2001).

Delta and Common Random Numbers

- In estimating delta, it is natural to start with the finite-difference estimate

$$e^{-r\tau} \frac{E[P(S + \epsilon)] - E[P(S - \epsilon)]}{2\epsilon}.$$

- $P(x)$ is the terminal payoff of the derivative security when the underlying asset's initial price equals x .
- Use simulation to estimate $E[P(S + \epsilon)]$ first.
- Use another simulation to estimate $E[P(S - \epsilon)]$.
- Finally, apply the formula to approximate the delta.
- This is also called the bump-and-revalue method.

Delta and Common Random Numbers (concluded)

- This method is not recommended because of its high variance.
- A much better approach is to use common random numbers to lower the variance:

$$e^{-r\tau} E \left[\frac{P(S + \epsilon) - P(S - \epsilon)}{2\epsilon} \right].$$

- Here, the *same* random numbers are used for $P(S + \epsilon)$ and $P(S - \epsilon)$.
- This holds for gamma and cross gamma.^a

^aFor multivariate derivatives.

Problems with the Bump-and-Revalue Method

- Consider the binary option with payoff

$$\begin{cases} 1, & \text{if } S(T) > X, \\ 0, & \text{otherwise.} \end{cases}$$

- Then

$$P(S+\epsilon) - P(S-\epsilon) = \begin{cases} 1, & \text{if } S + \epsilon > X \text{ and } S - \epsilon < X, \\ 0, & \text{otherwise.} \end{cases}$$

- So the finite-difference estimate per run for the (undiscounted) delta is 0 or $O(1/\epsilon)$.
- This means high variance.

Problems with the Bump-and-Revalue Method (concluded)

- The price of the binary option equals

$$e^{-r\tau} N(x - \sigma\sqrt{\tau}).$$

- It equals *minus* the derivative of the European call with respect to X .
- It also equals $X\tau$ times the rho of a European call (p. 355).
- Its delta is

$$\frac{N'(x - \sigma\sqrt{\tau})}{S\sigma\sqrt{\tau}}.$$

Gamma

- The finite-difference formula for gamma is

$$e^{-r\tau} E \left[\frac{P(S + \epsilon) - 2 \times P(S) + P(S - \epsilon)}{\epsilon^2} \right].$$

- For a correlation option with multiple underlying assets, the finite-difference formula for the cross gamma $\partial^2 P(S_1, S_2, \dots) / (\partial S_1 \partial S_2)$ is:

$$e^{-r\tau} E \left[\frac{P(S_1 + \epsilon_1, S_2 + \epsilon_2) - P(S_1 - \epsilon_1, S_2 + \epsilon_2)}{4\epsilon_1 \epsilon_2} - \frac{P(S_1 + \epsilon_1, S_2 - \epsilon_2) + P(S_1 - \epsilon_1, S_2 - \epsilon_2)}{4\epsilon_1 \epsilon_2} \right].$$

Gamma (continued)

- Choosing an ϵ of the right magnitude can be challenging.
 - If ϵ is too large, inaccurate Greeks result.
 - If ϵ is too small, unstable Greeks result.
- This phenomenon is sometimes called the curse of differentiation.^a

^aAït-Sahalia & Lo (1998); Bondarenko (2003).

Gamma (continued)

- In general, suppose

$$\frac{\partial^i}{\partial \theta^i} e^{-r\tau} E[P(S)] = e^{-r\tau} E \left[\frac{\partial^i P(S)}{\partial \theta^i} \right]$$

holds for all $i > 0$, where θ is a parameter of interest.^a

– A common requirement is Lipschitz continuity.^b

- Then Greeks become integrals.
- As a result, we avoid ϵ , finite differences, and resimulation.

^a $\partial^i P(S)/\partial \theta^i$ may not be partial differentiation in the classic sense.

^bBroadie & Glasserman (1996).

Gamma (continued)

- This is indeed possible for a broad class of payoff functions.^a
 - Roughly speaking, any payoff function that is equal to a sum of products of differentiable functions and indicator functions with the right kind of support.
 - For example, the payoff of a call is

$$\max(S(T) - X, 0) = (S(T) - X)I_{\{S(T) - X \geq 0\}}.$$

- The results are too technical to cover here (see next page).

^aTeng (R91723054) (2004); Lyuu & Teng (R91723054) (2011).

Gamma (continued)

- Suppose $h(\theta, x) \in \mathcal{H}$ with pdf $f(x)$ for x and $g_j(\theta, x) \in \mathcal{G}$ for $j \in \mathcal{B}$, a finite set of natural numbers.
- Then

$$\begin{aligned}
 & \frac{\partial}{\partial \theta} \int_{\mathfrak{R}} h(\theta, x) \prod_{j \in \mathcal{B}} \mathbf{1}_{\{g_j(\theta, x) > 0\}}(x) f(x) dx \\
 = & \int_{\mathfrak{R}} h_{\theta}(\theta, x) \prod_{j \in \mathcal{B}} \mathbf{1}_{\{g_j(\theta, x) > 0\}}(x) f(x) dx \\
 & + \sum_{l \in \mathcal{B}} \left[h(\theta, x) J_l(\theta, x) \prod_{j \in \mathcal{B} \setminus l} \mathbf{1}_{\{g_j(\theta, x) > 0\}}(x) f(x) \right]_{x=\chi_l(\theta)},
 \end{aligned}$$

where

$$J_l(\theta, x) = \text{sign} \left(\frac{\partial g_l(\theta, x)}{\partial x_k} \right) \frac{\partial g_l(\theta, x) / \partial \theta}{\partial g_l(\theta, x) / \partial x} \text{ for } l \in \mathcal{B}.$$

Gamma (concluded)

- Similar results have been derived for Levy processes.^a
- Formulas are also recently obtained for credit derivatives.^b
- In queueing networks, this is called infinitesimal perturbation analysis (IPA).^c

^aLyu, Teng (R91723054), & S. Wang (2013).

^bLyu, Teng (R91723054), & Tseng (2014, 2018).

^cCao (1985); Y. C. Ho & Cao (1985).