Backward Induction on the RT Tree

• After the RT tree is constructed, it can be used to price options by backward induction.

• Recall that each node keeps two variances $h^2_{\text{max}}$ and $h^2_{\text{min}}$.

• We now increase that number to $K$ equally spaced variances between $h^2_{\text{max}}$ and $h^2_{\text{min}}$ at each node.

• Besides the minimum and maximum variances, the other $K - 2$ variances in between are linearly interpolated.\(^a\)

\(^a\)In practice, log-linear interpolation works better (Lyuu & C. Wu (R90723065), 2005). Log-cubic interpolation works even better (C. Liu (R92922123), 2005).
Backward Induction on the RT Tree (continued)

- For example, if $K = 3$, then a variance of

$$10.5436 \times 10^{-6}$$

will be added between the maximum and minimum variances at node $(2, 0)$ on p. 936.

- In general, the $k$th variance at node $(i, j)$ is

$$h_{\text{min}}^2(i, j) + k \frac{h_{\text{max}}^2(i, j) - h_{\text{min}}^2(i, j)}{K - 1}, \quad k = 0, 1, \ldots, K - 1.$$ 

- Each interpolated variance’s jump parameter and branching probabilities can be computed as before.

\[a\] Repeated on p. 956.
Backward Induction on the RT Tree (concluded)

• Suppose a variance falls between two of the $K$ variances during backward induction.

• Linear interpolation of the option prices corresponding to the two bracketing variances will be used as the approximate option price.

• The above ideas are reminiscent of the ones on p. 431, where we dealt with Asian options.
Numerical Examples

- We next use the tree on p. 956 to price a European call option with a strike price of 100 and expiring at date 3.
- Recall that the riskless interest rate is zero.
- Assume $K = 2$; hence there are no interpolated variances.
- The pricing tree is shown on p. 959 with a call price of 0.66346.
  - The branching probabilities needed in backward induction can be found on p. 960.
Numerical Examples (continued)

- Let us derive some of the numbers on p. 959.
- A gray line means the updated variance falls strictly between $h_{\text{max}}^2$ and $h_{\text{min}}^2$.
- The option price for a terminal node at date 3 equals $\max(S_3 - 100, 0)$, independent of the variance level.
- Now move on to nodes at date 2.
- The option price at node (2, 3) depends on those at nodes (3, 5), (3, 3), and (3, 1).
- It therefore equals

$$0.1387 \times 5.37392 + 0.7197 \times 3.19054 + 0.1416 \times 1.05240 = 3.19054.$$
Numerical Examples (continued)

• Option prices for other nodes at date 2 can be computed similarly.

• For node \((1, 1)\), the option price for both variances is

\[
0.1237 \times 3.19054 + 0.7499 \times 1.05240 + 0.1264 \times 0.14573 = 1.20241.
\]

• Node \((1, 0)\) is most interesting.

• We knew that a down move from it gives a variance of \(0.000105609\).

• This number falls between the minimum variance \(0.000105173\) and the maximum variance \(0.0001227\) at node \((2, -1)\) on p. 960.
Numerical Examples (continued)

• The option price corresponding to the minimum variance is 0 (p. 960).

• The option price corresponding to the maximum variance is 0.14573.

• The equation

\[ x \times 0.000105173 + (1 - x) \times 0.0001227 = 0.000105609 \]

is satisfied by \( x = 0.9751 \).

• So the option for the down state is approximated by

\[ x \times 0 + (1 - x) \times 0.14573 = 0.00362. \]
Numerical Examples (continued)

- The up move leads to the state with option price 1.05240.
- The middle move leads to the state with option price 0.48366.
- The option price at node (1, 0) is finally calculated as

\[ 0.4775 \times 1.05240 + 0.0400 \times 0.48366 + 0.4825 \times 0.00362 = 0.52360. \]
Numerical Examples (continued)

- A variance following an interpolated variance may exceed the maximum variance or be exceeded by the minimum variance.

- When this happens, the option price corresponding to the maximum or minimum variance will be used during backward induction.\(^a\)

\(^a\)Cakici & Topyan (2000).
Numerical Examples (concluded)

- But an interpolated variance may choose a branch that goes into a node that is not reached in forward induction.\(^a\)

- In this case, the algorithm fails.

- The RT algorithm does not have this problem.
  - This is because all interpolated variances are involved in the forward-induction phase.

- It may be hard to calculate the implied \(\beta_1\) and \(\beta_2\) from option prices.\(^b\)

\(^a\) Lyuu & C. Wu (R90723065) (2005).
\(^b\) Y. Chang (B89704039, R93922034) (2006).
Complexities of GARCH Models\textsuperscript{a}

- The RT algorithm explodes exponentially if \( n \) is big enough (p. 932).

- The mean-tracking tree of Lyuu and Wu (2005) makes sure explosion does not happen if \( n \) is not too large.\textsuperscript{b}

- The next page summarizes the situations for many GARCH option pricing models.
  - Our earlier treatment is for NGARCH only.

\textsuperscript{a}Lyuu & C. Wu (R90723065) (2003, 2005).
\textsuperscript{b}Similar to, but earlier than, the binomial-trinomial tree on pp. 725ff.
Complexities of GARCH Models (concluded)\textsuperscript{a}

<table>
<thead>
<tr>
<th>Model</th>
<th>Explosion</th>
<th>Non-explosion</th>
</tr>
</thead>
<tbody>
<tr>
<td>NGARCH</td>
<td>$\beta_1 + \beta_2 n &gt; 1$</td>
<td>$\beta_1 + \beta_2 (\sqrt{n} + \lambda + c)^2 \leq 1$</td>
</tr>
<tr>
<td>LGARCH</td>
<td>$\beta_1 + \beta_2 n &gt; 1$</td>
<td>$\beta_1 + \beta_2 (\sqrt{n} + \lambda)^2 \leq 1$</td>
</tr>
<tr>
<td>AGARCH</td>
<td>$\beta_1 + \beta_2 n &gt; 1$</td>
<td>$\beta_1 + \beta_2 (\sqrt{n} + \lambda)^2 \leq 1$</td>
</tr>
<tr>
<td>GJR-GARCH</td>
<td>$\beta_1 + \beta_2 n &gt; 1$</td>
<td>$\beta_1 + (\beta_2 + \beta_3)(\sqrt{n} + \lambda)^2 \leq 1$</td>
</tr>
<tr>
<td>TS-GARCH</td>
<td>$\beta_1 + \beta_2 \sqrt{n} &gt; 1$</td>
<td>$\beta_1 + \beta_2 (\lambda + \sqrt{n}) \leq 1$</td>
</tr>
<tr>
<td>TGARCH</td>
<td>$\beta_1 + \beta_2 \sqrt{n} &gt; 1$</td>
<td>$\beta_1 + (\beta_2 + \beta_3)(\lambda + \sqrt{n}) \leq 1$</td>
</tr>
<tr>
<td>Heston-Nandi</td>
<td>$\beta_1 + \beta_2 (c - \frac{1}{2})^2 &gt; 1$ &amp; $c \leq \frac{1}{2}$</td>
<td>$\beta_1 + \beta_2 c^2 \leq 1$</td>
</tr>
<tr>
<td>VGARCH</td>
<td>$\beta_1 + (\beta_2 / 4) &gt; 1$</td>
<td>$\beta_1 \leq 1$</td>
</tr>
</tbody>
</table>

\textsuperscript{a}Y. C. Chen (R95723051) (2008); Y. C. Chen (R95723051), Lyuu, & Wen (D94922003) (2012).
Introduction to Term Structure Modeling
The fox often ran to the hole by which they had come in, to find out if his body was still thin enough to slip through it.

— Grimm’s Fairy Tales
And the worst thing you can have is models and spreadsheets.

Outline

- Use the binomial interest rate tree to model stochastic term structure.
  - Illustrates the basic ideas underlying future models.
  - Applications are generic in that pricing and hedging methodologies can be easily adapted to other models.

- Although the idea is similar to the earlier one used in option pricing, the current task is more complicated.
  - The evolution of an entire term structure, not just a single stock price, is to be modeled.
  - Interest rates of various maturities cannot evolve arbitrarily, or arbitrage profits may occur.
Issues

• A stochastic interest rate model performs two tasks.
  – Provides a stochastic process that defines future term structures without arbitrage profits.
  – “Consistent” with the observed term structures.
History

• The methodology was founded by Merton (1970).

• Modern interest rate modeling is often traced to 1977 when Vasicek and Cox, Ingersoll, and Ross developed simultaneously their influential models.

• Early models have fitting problems because they may not price today’s benchmark bonds correctly.

• An alternative approach pioneered by Ho and Lee (1986) makes fitting the market yield curve mandatory.

• Models based on such a paradigm are called (somewhat misleadingly) arbitrage-free or no-arbitrage models.
Binomial Interest Rate Tree

- Goal is to construct a no-arbitrage interest rate tree consistent with the yields and/or yield volatilities of zero-coupon bonds of all maturities.
  - This procedure is called calibration.$^a$

- Pick a binomial tree model in which the logarithm of the future short rate obeys the binomial distribution.
  - Exactly like the CRR tree.

- The limiting distribution of the short rate at any future time is hence lognormal.

---

$^a$Derman (2004), “complexity without calibration is pointless.”
Binomial Interest Rate Tree (continued)

- A binomial tree of future short rates is constructed.
- Every short rate is followed by two short rates in the following period (p. 977).
- In the figure on p. 977, node A coincides with the start of period $j$ during which the short rate $r$ is in effect.
- At the conclusion of period $j$, a new short rate goes into effect for period $j + 1$. 
period $j - 1$  |  period $j$  |  period $j + 1$

time $j - 1$  |  time $j$

\[ r \]  
\[ 0.5 \]  
\[ r \]  
\[ 0.5 \]  
\[ r_{\ell} \]  
\[ r_{h} \]
Binomial Interest Rate Tree (continued)

- This may take one of two possible values:
  - $r_\ell$: the “low” short-rate outcome at node B.
  - $r_h$: the “high” short-rate outcome at node C.
- Each branch has a 50% chance of occurring in a risk-neutral economy.
- We require that the paths combine as the binomial process unfolds.
- This model can be traced to Salomon Brothers.\(^a\)

\(^a\)Tuckman (2002).
Binomial Interest Rate Tree (continued)

• The short rate $r$ can go to $r_h$ and $r_\ell$ with equal risk-neutral probability $1/2$ in a period of length $\Delta t$.

• Hence the volatility of $\ln r$ after $\Delta t$ time is

$$\sigma = \frac{1}{2} \frac{1}{\sqrt{\Delta t}} \ln \left( \frac{r_h}{r_\ell} \right). \quad (127)$$

• Above, $\sigma$ is annualized, whereas $r_\ell$ and $r_h$ are period based.

\[ ^a \text{See Exercise 23.2.3 in text.} \]
Binomial Interest Rate Tree (continued)

- Note that
  \[ \frac{r_h}{r_\ell} = e^{2\sigma\sqrt{\Delta t}}. \]

- Thus greater volatility, hence uncertainty, leads to larger \( r_h/r_\ell \) and wider ranges of possible short rates.

- The ratio \( r_h/r_\ell \) may depend on time if the volatility is a function of time.

- Note that \( r_h/r_\ell \) has nothing to do with the current short rate \( r \) if \( \sigma \) is independent of \( r \).
Binomial Interest Rate Tree (continued)

- In general there are \( j \) possible rates\(^a\) for period \( j \),
  \[
  r_j, r_j v_j, r_j v_j^2, \ldots, r_j v_j^{j-1},
  \]
  where
  \[
  v_j \triangleq e^{2\sigma_j \sqrt{\Delta t}}
  \]
  (128)
  is the multiplicative ratio for the rates in period \( j \) (see figure on next page).
- We shall call \( r_j \) the baseline rates.
- The subscript \( j \) in \( \sigma_j \) above is meant to emphasize that the short rate volatility may be time dependent.

\(^{a}\)Not \( j + 1 \).
Binomial Interest Rate Tree (concluded)

• In the limit, the short rate follows

\[ r(t) = \mu(t) e^{\sigma(t) W(t)}. \] (129)

  – The (percent) short rate volatility \( \sigma(t) \) is a deterministic function of time.

• The expected value of \( r(t) \) equals \( \mu(t) e^{\sigma(t)^2 (t/2)} \).

• Hence a *declining* short rate volatility is usually imposed to preclude the short rate from assuming implausibly high values.

• Incidentally, this is how the binomial interest rate tree achieves mean reversion to some long-term mean.
Memory Issues

- Path independency: The term structure at any node is independent of the path taken to reach it.

- So only the baseline rates $r_i$ and the multiplicative ratios $v_i$ need to be stored in computer memory.

- This takes up only $O(n)$ space.$^a$

- Storing the whole tree would take up $O(n^2)$ space.
  - Daily interest rate movements for 30 years require roughly $(30 \times 365)^2/2 \approx 6 \times 10^7$ double-precision floating-point numbers (half a gigabyte!).

$^a$Throughout, $n$ denotes the depth of the tree.
Set Things in Motion

• The abstract process is now in place.

• We need the yields to maturities of the riskless bonds that make up the benchmark yield curve and their volatilities.

• In the U.S., for example, the on-the-run yield curve obtained by the most recently issued Treasury securities may be used as the benchmark curve.
Set Things in Motion (concluded)

- The term structure of (yield) volatilities\textsuperscript{a} can be estimated from:
  - Historical data (historical volatility).
  - Or interest rate option prices such as cap prices (implied volatility).
- The binomial tree should be found that is consistent with both term structures.
- Here we focus on the term structure of interest rates.

\textsuperscript{a}Or simply the volatility (term) structure.
Model Term Structures

• The model price is computed by backward induction.

• Refer back to the figure on p. 977.

• Given that the values at nodes B and C are $P_B$ and $P_C$, respectively, the value at node A is then

$$\frac{P_B + P_C}{2(1 + r)} + \text{cash flow at node A}.$$

• We compute the values column by column (see next page).

• This takes $O(n^2)$ time and $O(n)$ space.
Cash flows:

C

D

A

B

P_1

P_2

P_3

P_4

C + \frac{P_1 + P_2}{2(1+r)}

C + \frac{P_2 + P_3}{2(1+rv)}

C + \frac{P_3 + P_4}{2(1+rv^2)}
Term Structure Dynamics

- An \( n \)-period zero-coupon bond’s price can be computed by assigning $1 to every node at period \( n \) and then applying backward induction.
- Repeating this step for \( n = 1, 2, \ldots \), one obtains the market discount function implied by the tree.
- The tree therefore determines a term structure.
- It also contains a term structure dynamics.
  - Taking any node in the tree as the current state induces a binomial interest rate tree and, again, a term structure.
Sample Term Structure

- We shall construct interest rate trees consistent with the sample term structure in the following table.
  - This is calibration (the reverse of pricing).

- Assume the short rate volatility is such that
  \[ v \equiv \frac{r_h}{r_\ell} = 1.5, \]
  independent of time.

<table>
<thead>
<tr>
<th>Period</th>
<th>1</th>
<th>2</th>
<th>3</th>
</tr>
</thead>
<tbody>
<tr>
<td>Spot rate (%)</td>
<td>4</td>
<td>4.2</td>
<td>4.3</td>
</tr>
<tr>
<td>One-period forward rate (%)</td>
<td>4</td>
<td>4.4</td>
<td>4.5</td>
</tr>
<tr>
<td>Discount factor</td>
<td>0.96154</td>
<td>0.92101</td>
<td>0.88135</td>
</tr>
</tbody>
</table>
An Approximate Calibration Scheme

- Start with the implied one-period forward rates.
- Equate the expected short rate with the forward rate.\textsuperscript{a}
- For the first period, the forward rate is today’s one-period spot rate.
- In general, let $f_j$ denote the forward rate in period $j$.
- This forward rate can be derived from the market discount function via\textsuperscript{b}

$$f_j = \frac{d(j)}{d(j + 1)} - 1.$$ 

\textsuperscript{a}See Exercise 5.6.6 in text.
\textsuperscript{b}See Exercise 5.6.3 in text.
An Approximate Calibration Scheme (continued)

• Since the $i$th short rate $r_j v_j^{i-1}$, $1 \leq i \leq j$, occurs with probability $2^{-(j-1)} \binom{j-1}{i-1}$, this means

$$
\sum_{i=1}^{j} 2^{-(j-1)} \binom{j-1}{i-1} r_j v_j^{i-1} = f_j.
$$

• Thus

$$
r_j = \left( \frac{2}{1 + v_j} \right)^{j-1} f_j. \quad (130)
$$

• This binomial interest rate tree is trivial to set up (implicitly), in $O(n)$ time.
An Approximate Calibration Scheme (continued)

- The ensuing tree for the sample term structure appears in figure next page.
- For example, the price of the zero-coupon bond paying $1 at the end of the third period is
  \[ \frac{1}{4} \times \frac{1}{1.04} \times \left( \frac{1}{1.0352} \times \left( \frac{1}{1.0288} + \frac{1}{1.0432} \right) + \frac{1}{1.0528} \times \left( \frac{1}{1.0432} + \frac{1}{1.0648} \right) \right) \]
  or 0.88155, which exceeds discount factor 0.88135.
- The tree is thus not calibrated.
Baseline rates:

- A: 4.0%
- B: 3.52%
- C: 2.88%
- D: 4.32%
- E: 5.28%
- F: 6.48%

Implied forward rates:

- Period 1: 4.0%
- Period 2: 4.4%
- Period 3: 4.5%
An Approximate Calibration Scheme (concluded)

• Indeed, this bias is inherent: The tree *overprices* the bonds.\(^a\)

• Suppose we replace the baseline rates \(r_j\) by \(r_jv_j\).

• Then the resulting tree *underprices* the bonds.\(^b\)

• The true baseline rates are thus bounded between \(r_j\) and \(r_jv_j\).

---

\(^a\)See Exercise 23.2.4 in text.

Issues in Calibration

- The model prices generated by the binomial interest rate tree should match the observed market prices.
- Perhaps the most crucial aspect of model building.
- Treat the backward induction for the model price of the $m$-period zero-coupon bond as computing some function $f(r_m)$ of the unknown baseline rate $r_m$ for period $m$.
- A root-finding method is applied to solve $f(r_m) = P$ for $r_m$ given the zero’s price $P$ and $r_1, r_2, \ldots, r_{m-1}$.
- This procedure is carried out for $m = 1, 2, \ldots, n$.
- It runs in $O(n^3)$ time.
Binomial Interest Rate Tree Calibration

- Calibration can be accomplished in $O(n^2)$ time by the use of forward induction.\(^a\)

- The scheme records how much $1 at a node contributes to the model price.

- This number is called the state price, the Arrow-Debreu price, or Green’s function.
  - It is the price of a state contingent claim that pays $1 at that particular node (state) and 0 elsewhere.

- The column of state prices will be established by moving *forward* from time 0 to time $n$.

Binomial Interest Rate Tree Calibration (continued)

• Suppose we are at time $j$ and there are $j+1$ nodes.
  – The unknown baseline rate for period $j$ is $r \overset{\Delta}{=} r_j$.
  – The multiplicative ratio is $v \overset{\Delta}{=} v_j$.
  – $P_1, P_2, \ldots, P_j$ are the known state prices at earlier time $j-1$.
  – They have rates $r, rv, \ldots, rv^{j-1}$ for period $j$.

• By definition, $\sum_{i=1}^{j} P_i$ is the price of the $(j-1)$-period zero-coupon bond.

• We want to find $r$ based on $P_1, P_2, \ldots, P_j$ and the price of the $j$-period zero-coupon bond.

\[a\text{Recall p. 982.}\]
Binomial Interest Rate Tree Calibration (continued)

• One dollar at time $j$ has a known market value of $1/[1 + S(j)]^j$, where $S(j)$ is the $j$-period spot rate.

• Alternatively, this dollar has a present value of

$$g(r) \triangleq \frac{P_1}{(1+r)} + \frac{P_2}{(1+rv)} + \frac{P_3}{(1+rv^2)} + \cdots + \frac{P_j}{(1+rv^{j-1})}$$

(see next plot).

• So we solve

$$g(r) = \frac{1}{[1 + S(j)]^j} \quad (131)$$

for $r$. 
Binomial Interest Rate Tree Calibration (continued)

- Given a decreasing market discount function, a unique positive solution for $r$ is guaranteed.

- The state prices at time $j$ can now be calculated (see panel (a) next page).

- We call a tree with these state prices a binomial state price tree (see panel (b) next page).

- The calibrated tree is depicted on p. 1003.
(a) The diagram shows a financial model with nodes A, B, C, and D. The probabilities and rates are calculated as follows:

\[ P_i = \frac{P_i}{2(1+r)} \]

\[ r \]

\[ r_v \]

\[ P_1 = \frac{P_1}{2(1+r)} + \frac{P_2}{2(1+r)} \]

(b) The implied forward rates are:

- Period 1: 4.0%
- Period 2: 4.4%
- Period 3: 4.5%

The values are:

- D: 0.112832
- C: 0.232197
- B: 0.333501
- A: 0.327842
- 2.895%
- 3.526%
- 0.228308
- 0.107173
Implied forward rates: 4.0% 4.4% 4.5%

period 1 period 2 period 3
Binomial Interest Rate Tree Calibration (concluded)

- The Newton-Raphson method can be used to solve for the $r$ in Eq. (131) on p. 999 as $g'(r)$ is easy to evaluate.

- The monotonicity and the convexity of $g(r)$ also facilitate root finding.

- The total running time is $O(n^2)$, as each root-finding routine consumes $O(j)$ time.

- With a good initial guess,\textsuperscript{a} the Newton-Raphson method converges in only a few steps.\textsuperscript{b}

\textsuperscript{a} Such as the $r_j = (\frac{2}{1+v_j})^j - 1 f_j$ on p. 992.

\textsuperscript{b} Lyuu (1999).
A Numerical Example

• One dollar at the end of the second period should have a present value of 0.92101 by the sample term structure.

• The baseline rate for the second period, $r_2$, satisfies

$$\frac{0.480769}{1 + r_2} + \frac{0.480769}{1 + 1.5 \times r_2} = 0.92101.$$ 

• The result is $r_2 = 3.526\%$.

• This is used to derive the next column of state prices shown in panel (b) on p. 1002 as 0.232197, 0.460505, and 0.228308.

• Their sum gives the correct market discount factor 0.92101.
A Numerical Example (concluded)

• The baseline rate for the third period, $r_3$, satisfies
  \[
  \frac{0.232197}{1 + r_3} + \frac{0.460505}{1 + 1.5 \times r_3} + \frac{0.228308}{1 + (1.5)^2 \times r_3} = 0.88135.
  \]

• The result is $r_3 = 2.895\%$.

• Now, redo the calculation on p. 993 using the new rates:
  \[
  \frac{1}{4} \times \frac{1}{1.04} \times \left[ \frac{1}{1.03526} \times \left( \frac{1}{1.02895} + \frac{1}{1.04343} \right) + \frac{1}{1.05289} \times \left( \frac{1}{1.04343} + \frac{1}{1.06514} \right) \right],
  \]
  which equals 0.88135, an exact match.

• The tree on p. 1003 prices without bias the benchmark securities.
Spread of Nonbenchmark Bonds

- Model prices calculated by the calibrated tree as a rule do not match market prices of nonbenchmark bonds.
- The incremental return over the benchmark bonds is called spread.
- If we add the spread uniformly over the short rates in the tree, the model price will equal the market price.
- We will apply the spread concept to option-free bonds next.
Spread of Nonbenchmark Bonds (continued)

• We illustrate the idea with an example.

• Start with the tree on p. 1009.

• Consider a security with cash flow $C_i$ at time $i$ for $i = 1, 2, 3$.

• Its model price is $p(s)$, which is equal to

$$\frac{1}{1.04 + s} \times \left[ C_1 + \frac{1}{2} \times \frac{1}{1.03526 + s} \times \left( C_2 + \frac{1}{2} \left( \frac{C_3}{1.02895 + s} + \frac{C_3}{1.04343 + s} \right) \right) + \frac{1}{2} \times \frac{1}{1.05289 + s} \times \left( C_2 + \frac{1}{2} \left( \frac{C_3}{1.04343 + s} + \frac{C_3}{1.06514 + s} \right) \right) \right].$$

• Given a market price of $P$, the spread is the $s$ that solves $P = p(s)$.
Implied forward rates: 4.0%  4.4%  4.5%

period 1  period 2  period 3
Spread of Nonbenchmark Bonds (continued)

- The model price $p(s)$ is a monotonically decreasing, convex function of $s$.

- We will employ the Newton-Raphson root-finding method to solve

  $$p(s) - P = 0$$

  for $s$.

- But a quick look at the equation for $p(s)$ reveals that evaluating $p'(s)$ directly is infeasible.

- Fortunately, the tree can be used to evaluate both $p(s)$ and $p'(s)$ during backward induction.
Spread of Nonbenchmark Bonds (continued)

• Consider an arbitrary node A in the tree associated with the short rate $r$.

• In the process of computing the model price $p(s)$, a price $p_A(s)$ is computed at A.

• Prices computed at A’s two successor nodes B and C are discounted by $r + s$ to obtain $p_A(s)$ as follows,

$$p_A(s) = c + \frac{p_B(s) + p_C(s)}{2(1 + r + s)},$$

where $c$ denotes the cash flow at A.
Spread of Nonbenchmark Bonds (continued)

- To compute $p'_A(s)$ as well, node A calculates

$$p'_A(s) = \frac{p'_B(s) + p'_C(s)}{2(1 + r + s)} - \frac{p_B(s) + p_C(s)}{2(1 + r + s)^2}.$$  \hspace{1cm} (132)

- This is easy if $p'_B(s)$ and $p'_C(s)$ are also computed at nodes B and C.

- When A is a terminal node, simply use the payoff function for $p_A(s)$.

\[a\]Contributed by Mr. Chou, Ming-Hsin (R02723073) on May 28, 2014.
\[ p_A(s) = c + \frac{p^b_A(s) + p^c_A(s)}{2(1+r+s)} \]

\[ p'_A(s) = \frac{p^b'_A(s) + p^c'_A(s)}{2(1+r+s)} - \frac{p^b_A(s) + p^c_A(s)}{2(1+r+s)^2} \]
Spread of Nonbenchmark Bonds (continued)

- Apply the above procedure inductively to yield $p(s)$ and $p'(s)$ at the root (p. 1013).

- This is called the differential tree method.\(^a\)
  - Similar ideas can be found in automatic differentiation (AD)\(^b\) and backpropagation\(^c\) in artificial neural networks.

- The total running time is $O(n^2)$.

- The memory requirement is $O(n)$.

\(^a\)Lyuu (1999).
\(^b\)Rall (1981).
\(^c\)Werbos (1974); Rumelhart, Hinton, & Williams (1986).
## Spread of Nonbenchmark Bonds (continued)

<table>
<thead>
<tr>
<th>Number of partitions $n$</th>
<th>Running time (s)</th>
<th>Number of iterations</th>
<th>Number of partitions</th>
<th>Running time (s)</th>
<th>Number of iterations</th>
</tr>
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<td>........................</td>
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</tr>
</tbody>
</table>

75MHz Sun SPARCstation 20.
Spread of Nonbenchmark Bonds (concluded)

- Consider a three-year, 5% bond with a market price of 100.569.
- Assume the bond pays annual interest.
- The spread can be shown to be 50 basis points over the tree (p. 1017).
- Note that the idea of spread does not assume parallel shifts in the term structure.
- It also differs from the yield spread (p. 130) and static spread (p. 131) of the nonbenchmark bond over an otherwise identical benchmark bond.
More Applications of the Differential Tree: Calculating Implied Volatility (in seconds)\textsuperscript{a}

<table>
<thead>
<tr>
<th>Number of partitions</th>
<th>Running time</th>
<th>Number of iterations</th>
<th>Number of partitions</th>
<th>Running time</th>
<th>Number of iterations</th>
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</tbody>
</table>

Intel 166MHz Pentium, running on Microsoft Windows 95.

\textsuperscript{a}Lyuu (1999).
Fixed-Income Options

• Consider a 2-year 99 European call on the 3-year, 5% Treasury.

• Assume the Treasury pays annual interest.

• From p. 1020 the 3-year Treasury’s price minus the $5 interest at year 2 could be $102.046, $100.630, or $98.579 two years from now.
  – The accrued interest is not included as it belongs to the original bondholder.

• Now compare the strike price against the bond prices.

• The call is in the money in the first two scenarios out of the money in the third.
Fixed-Income Options (continued)

- The option value is calculated to be $1.458 on p. 1020(a).

- European interest rate puts can be valued similarly.

- Consider a two-year 99 European put on the same security.

- At expiration, the put is in the money only when the Treasury is worth $98.579 without the accrued interest.

- The option value is computed to be $0.096 on p. 1020(b).
Fixed-Income Options (concluded)

- The present value of the strike price is \( PV(X) = 99 \times 0.92101 = 91.18 \).
- The Treasury is worth \( B = 101.955 \).
- The present value of the interest payments during the life of the options is
  \[
  PV(I) = 5 \times 0.96154 + 5 \times 0.92101 = 9.41275.
  \]
- The call and the put are worth \( C = 1.458 \) and \( P = 0.096 \), respectively.
- Hence the put-call parity is preserved:
  \[
  C = P + B - PV(I) - PV(X).
  \]
Delta or Hedge Ratio

- How much does the option price change in response to changes in the *price* of the underlying bond?
- This relation is called delta (or hedge ratio) defined as
  \[
  \frac{O_h - O_\ell}{P_h - P_\ell}.
  \]
- In the above, $P_h$ and $P_\ell$ denote the bond prices if the short rate moves up and down, respectively.
- Similarly, $O_h$ and $O_\ell$ denote the option values if the short rate moves up and down, respectively.
Delta or Hedge Ratio (concluded)

• Delta measures the sensitivity of the option value to changes in the underlying bond price.

• So it shows how to hedge one with the other.

• Take the call and put on p. 1020 as examples.

• Their deltas are

\[
\begin{align*}
\frac{0.774 - 2.258}{99.350 - 102.716} &= 0.441, \\
\frac{0.200 - 0.000}{99.350 - 102.716} &= -0.059,
\end{align*}
\]

respectively.
Volatility Term Structures

- The binomial interest rate tree can be used to calculate the yield volatility of zero-coupon bonds.
- Consider an $n$-period zero-coupon bond.
- First find its yield to maturity $y_h$ ($y_\ell$, respectively) at the end of the initial period if the short rate rises (declines, respectively).
- The yield volatility for our model is defined as

$$\frac{1}{2} \ln \left( \frac{y_h}{y_\ell} \right).$$

(133)
Volatility Term Structures (continued)

- For example, based on the tree on p. 1003, the two-year zero’s yield at the end of the first period is 5.289% if the rate rises and 3.526% if the rate declines.

- Its yield volatility is therefore

$$\frac{1}{2} \ln \left( \frac{0.05289}{0.03526} \right) = 20.273\%.$$
Volatility Term Structures (continued)

• Consider the three-year zero-coupon bond.

• If the short rate rises, the price of the zero one year from now will be

\[
\frac{1}{2} \times \frac{1}{1.05289} \times \left( \frac{1}{1.04343} + \frac{1}{1.06514} \right) = 0.90096.
\]

• Thus its yield is \( \sqrt{\frac{1}{0.90096}} - 1 = 0.053531 \).

• If the short rate declines, the price of the zero one year from now will be

\[
\frac{1}{2} \times \frac{1}{1.03526} \times \left( \frac{1}{1.02895} + \frac{1}{1.04343} \right) = 0.93225.
\]
Volatility Term Structures (continued)

- Thus its yield is \( \sqrt{\frac{1}{0.93225}} - 1 = 0.0357 \).
- The yield volatility is hence
  \[
  \frac{1}{2} \ln \left( \frac{0.053531}{0.0357} \right) = 20.256\% ,
  \]
  slightly less than the one-year yield volatility.
- This is consistent with the reality that longer-term bonds typically have lower yield volatilities than shorter-term bonds.\(^a\)
- The procedure can be repeated for longer-term zeros to obtain their yield volatilities.

\(^a\)The relation is reversed for price volatilities (duration).
Short rate volatility given flat 10% volatility term structure.
Volatility Term Structures (concluded)

• We started with $v_i$ and then derived the volatility term structure.

• In practice, the steps are reversed.

• The volatility term structure is supplied by the user along with the term structure.

• The $v_i$—hence the short rate volatilities via Eq. (128) on p. 981—and the $r_i$ are then simultaneously determined.

• The result is the Black-Derman-Toy model of Goldman Sachs.$^a$

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$^a$Black, Derman, & Toy (1990).