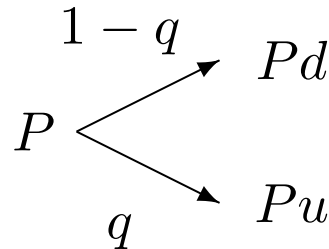


## The Binomial Model

- The analytical framework can be nicely illustrated with the binomial model.
- Suppose the bond price  $P$  can move with probability  $q$  to  $Pu$  and probability  $1 - q$  to  $Pd$ , where  $u > d$ :



## The Binomial Model (continued)

- Over the period, the bond's expected rate of return is

$$\hat{\mu} \equiv \frac{qPu + (1 - q)Pd}{P} - 1 = qu + (1 - q)d - 1. \quad (112)$$

- The variance of that return rate is

$$\hat{\sigma}^2 \equiv q(1 - q)(u - d)^2. \quad (113)$$

## The Binomial Model (continued)

- In particular, the bond whose maturity is one period away will move from a price of  $1/(1+r)$  to its par value \$1.
- This is the money market account modeled by the short rate  $r$ .
- The market price of risk is defined as  $\lambda \equiv (\hat{\mu} - r)/\hat{\sigma}$ .
- As in the continuous-time case, it can be shown that  $\lambda$  is independent of the maturity of the bond (see text).

## The Binomial Model (concluded)

- Now change the probability from  $q$  to

$$p \equiv q - \lambda \sqrt{q(1-q)} = \frac{(1+r) - d}{u-d}, \quad (114)$$

which is independent of bond maturity and  $q$ .

– Recall the BOPM.

- The bond's expected rate of return becomes

$$\frac{pPu + (1-p)Pd}{P} - 1 = pu + (1-p)d - 1 = r.$$

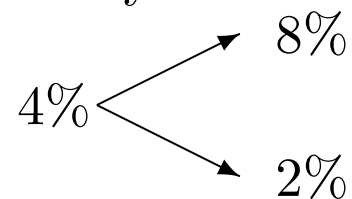
- The local expectations theory hence holds under the new probability measure  $p$ .

## Numerical Examples

- Assume this spot rate curve:

Year	1	2
Spot rate	4%	5%

- Assume the one-year rate (short rate) can move up to 8% or down to 2% after a year:



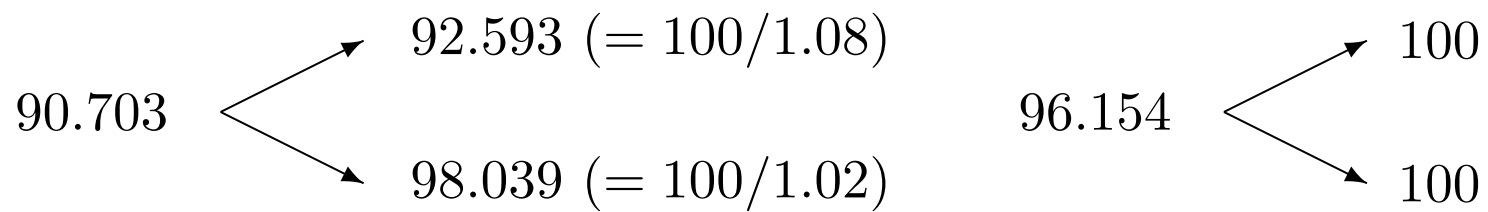
## Numerical Examples (continued)

- No real-world probabilities are specified.
- The prices of one- and two-year zero-coupon bonds are, respectively,

$$\begin{aligned}100/1.04 &= 96.154, \\ 100/(1.05)^2 &= 90.703.\end{aligned}$$

- They follow the binomial processes on p. 925.

## Numerical Examples (continued)



The price process of the two-year zero-coupon bond is on the left; that of the one-year zero-coupon bond is on the right.

## Numerical Examples (continued)

- The pricing of derivatives can be simplified by assuming investors are risk-neutral.
- Suppose all securities have the same expected one-period rate of return, the riskless rate.
- Then

$$(1 - p) \times \frac{92.593}{90.703} + p \times \frac{98.039}{90.703} - 1 = 4\%,$$

where  $p$  denotes the risk-neutral probability of a down move in rates.

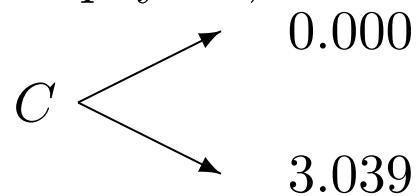


## Numerical Examples (concluded)

- Solving the equation leads to  $p = 0.319$ .
- Interest rate contingent claims can be priced under this probability.

## Numerical Examples: Fixed-Income Options

- A one-year European call on the two-year zero with a \$95 strike price has the payoffs,



- To solve for the option value  $C$ , we replicate the call by a portfolio of  $x$  one-year and  $y$  two-year zeros.

## Numerical Examples: Fixed-Income Options (continued)

- This leads to the simultaneous equations,

$$x \times 100 + y \times 92.593 = 0.000,$$

$$x \times 100 + y \times 98.039 = 3.039.$$

- They give  $x = -0.5167$  and  $y = 0.5580$ .
- Consequently,

$$C = x \times 96.154 + y \times 90.703 \approx 0.93$$

to prevent arbitrage.

## Numerical Examples: Fixed-Income Options (continued)

- This price is derived without assuming any version of an expectations theory.
- Instead, the arbitrage-free price is derived by replication.
- The price of an interest rate contingent claim does not depend directly on the real-world probabilities.
- The dependence holds only indirectly via the current bond prices.

## Numerical Examples: Fixed-Income Options (concluded)

- An equivalent method is to utilize risk-neutral pricing.
- The above call option is worth

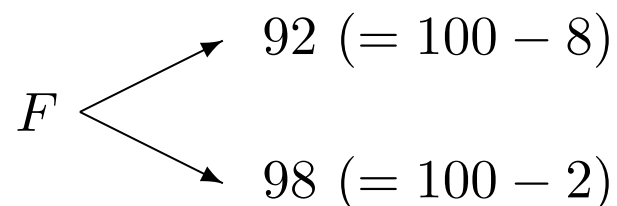
$$C = \frac{(1 - p) \times 0 + p \times 3.039}{1.04} \approx 0.93,$$

the same as before.

- This is not surprising, as arbitrage freedom and the existence of a risk-neutral economy are equivalent.

## Numerical Examples: Futures and Forward Prices

- A one-year futures contract on the one-year rate has a payoff of  $100 - r$ , where  $r$  is the one-year rate at maturity:



- As the futures price  $F$  is the expected future payoff (see text or p. 464),

$$F = (1 - p) \times 92 + p \times 98 = 93.914.$$

## Numerical Examples: Futures and Forward Prices (concluded)

- The forward price for a one-year forward contract on a one-year zero-coupon bond is<sup>a</sup>

$$90.703/96.154 = 94.331\%.$$

- The forward price exceeds the futures price.<sup>b</sup>

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<sup>a</sup>See Eq. (100) on p. 898.

<sup>b</sup>Recall p. 410.

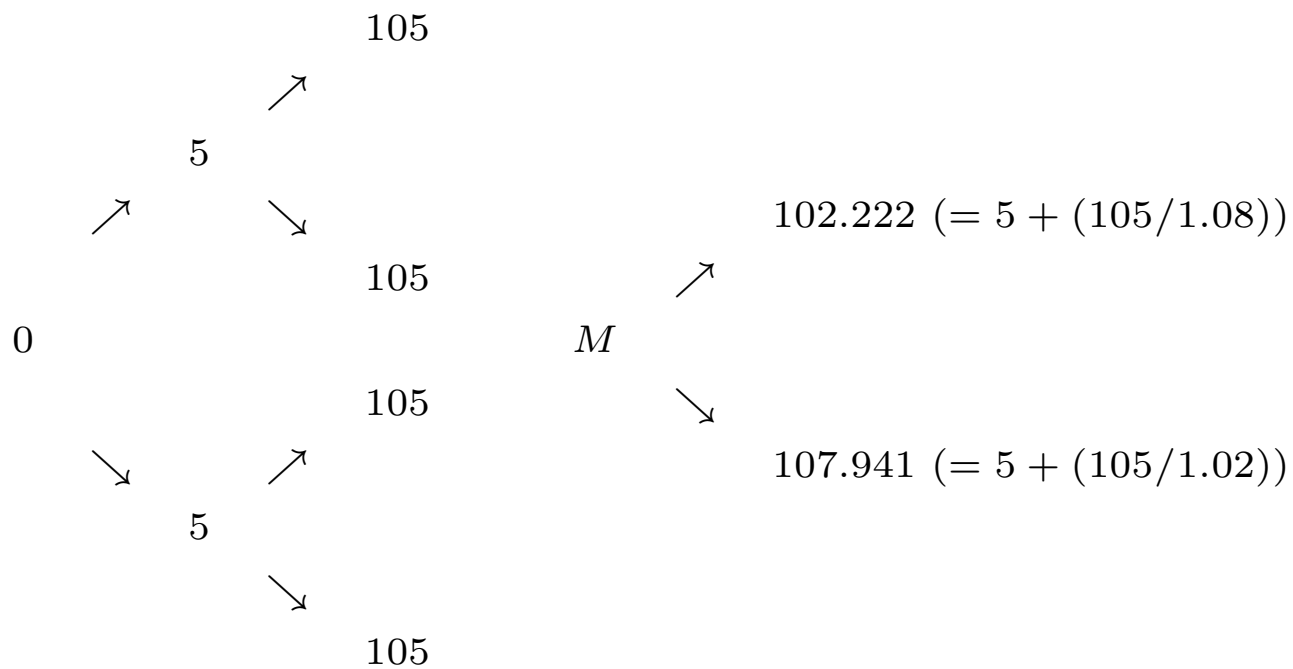
## Numerical Examples: Mortgage-Backed Securities

- Consider a 5%-coupon, two-year mortgage-backed security without amortization, prepayments, and default risk.
- Its cash flow and price process are illustrated on p. 935.
- Its fair price is

$$M = \frac{(1 - p) \times 102.222 + p \times 107.941}{1.04} = 100.045.$$

- Identical results could have been obtained via arbitrage considerations.

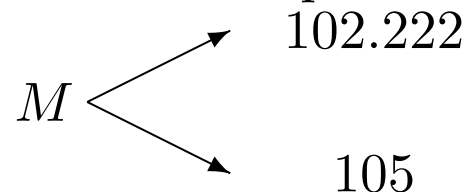




The left diagram depicts the cash flow; the right diagram illustrates the price process.

## Numerical Examples: MBSs (continued)

- Suppose that the security can be prepaid at par.
- It will be prepaid only when its price is higher than par.
- Prepayment will hence occur only in the “down” state when the security is worth 102.941 (excluding coupon).
- The price therefore follows the process,



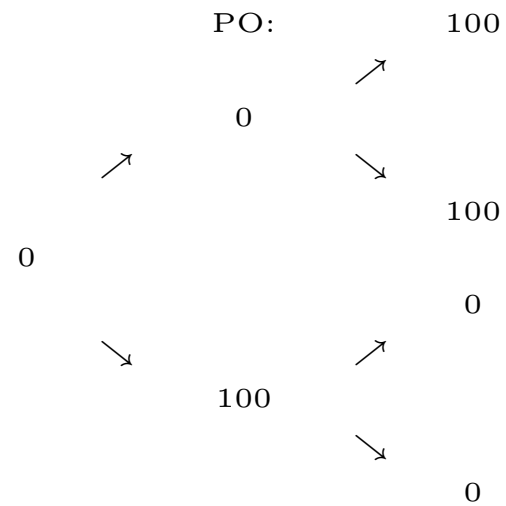
- The security is worth

$$M = \frac{(1 - p) \times 102.222 + p \times 105}{1.04} = 99.142.$$

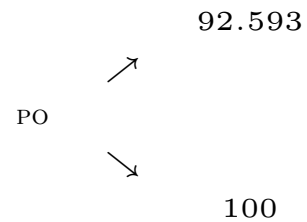
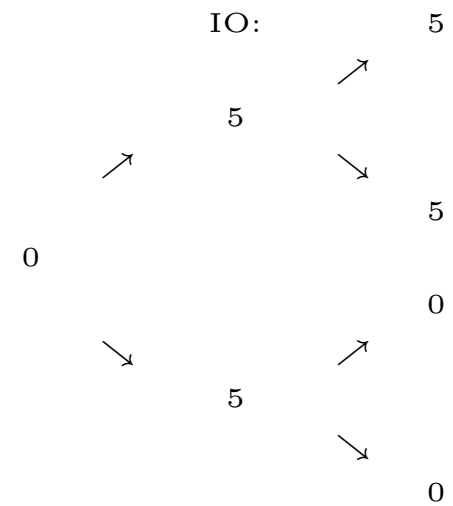
## Numerical Examples: MBSs (continued)

- The cash flow of the principal-only (PO) strip comes from the mortgage's principal cash flow.
- The cash flow of the interest-only (IO) strip comes from the interest cash flow (p. 938(a)).
- Their prices hence follow the processes on p. 938(b).
- The fair prices are

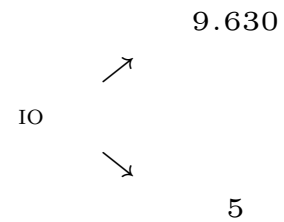
$$\begin{aligned} \text{PO} &= \frac{(1 - p) \times 92.593 + p \times 100}{1.04} = 91.304, \\ \text{IO} &= \frac{(1 - p) \times 9.630 + p \times 5}{1.04} = 7.839. \end{aligned}$$



(a)



(b)



The price 9.630 is derived from  $5 + (5/1.08)$ .

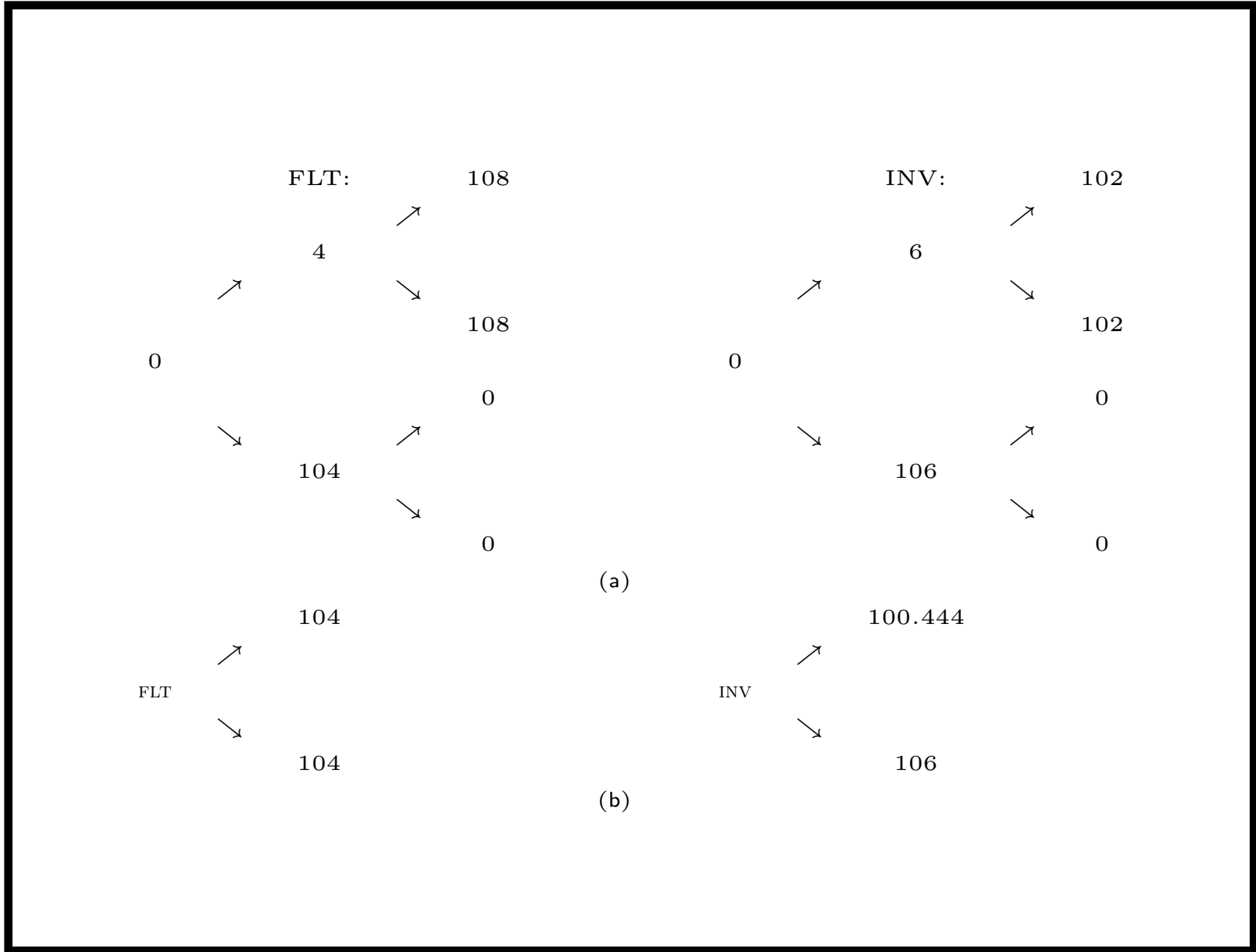
## Numerical Examples: MBSs (continued)

- Suppose the mortgage is split into half floater and half inverse floater.
- Let the floater (FLT) receive the one-year rate.
- Then the inverse floater (INV) must have a coupon rate of

$$(10\% - \text{one-year rate})$$

to make the overall coupon rate 5%.

- Their cash flows as *percentages of par* and values are shown on p. 940.



## Numerical Examples: MBSs (concluded)

- On p. 940, the floater's price in the up node, 104, is derived from  $4 + (108/1.08)$ .
- The inverse floater's price 100.444 is derived from  $6 + (102/1.08)$ .
- The current prices are

$$\text{FLT} = \frac{1}{2} \times \frac{104}{1.04} = 50,$$

$$\text{INV} = \frac{1}{2} \times \frac{(1 - p) \times 100.444 + p \times 106}{1.04} = 49.142.$$

# *Equilibrium Term Structure Models*



8. What's your problem? Any moron  
can understand bond pricing models.  
— *Top Ten Lies Finance Professors  
Tell Their Students*

## Introduction

- This chapter surveys equilibrium models.
- Since the spot rates satisfy

$$r(t, T) = -\frac{\ln P(t, T)}{T - t},$$

the discount function  $P(t, T)$  suffices to establish the spot rate curve.

- All models to follow are short rate models.
- Unless stated otherwise, the processes are risk-neutral.

## The Vasicek Model<sup>a</sup>

- The short rate follows

$$dr = \beta(\mu - r) dt + \sigma dW.$$

- The short rate is pulled to the long-term mean level  $\mu$  at rate  $\beta$ .
- Superimposed on this “pull” is a normally distributed stochastic term  $\sigma dW$ .
- Since the process is an Ornstein-Uhlenbeck process,

$$E[r(T) | r(t) = r] = \mu + (r - \mu) e^{-\beta(T-t)}$$

from Eq. (58) on p. 523.

---

<sup>a</sup>Vasicek (1977).

## The Vasicek Model (continued)

- The price of a zero-coupon bond paying one dollar at maturity can be shown to be

$$P(t, T) = A(t, T) e^{-B(t, T) r(t)}, \quad (115)$$

where

$$A(t, T) = \begin{cases} \exp \left[ \frac{(B(t, T) - T + t)(\beta^2 \mu - \sigma^2 / 2)}{\beta^2} - \frac{\sigma^2 B(t, T)^2}{4\beta} \right] & \text{if } \beta \neq 0, \\ \exp \left[ \frac{\sigma^2 (T - t)^3}{6} \right] & \text{if } \beta = 0. \end{cases}$$

and

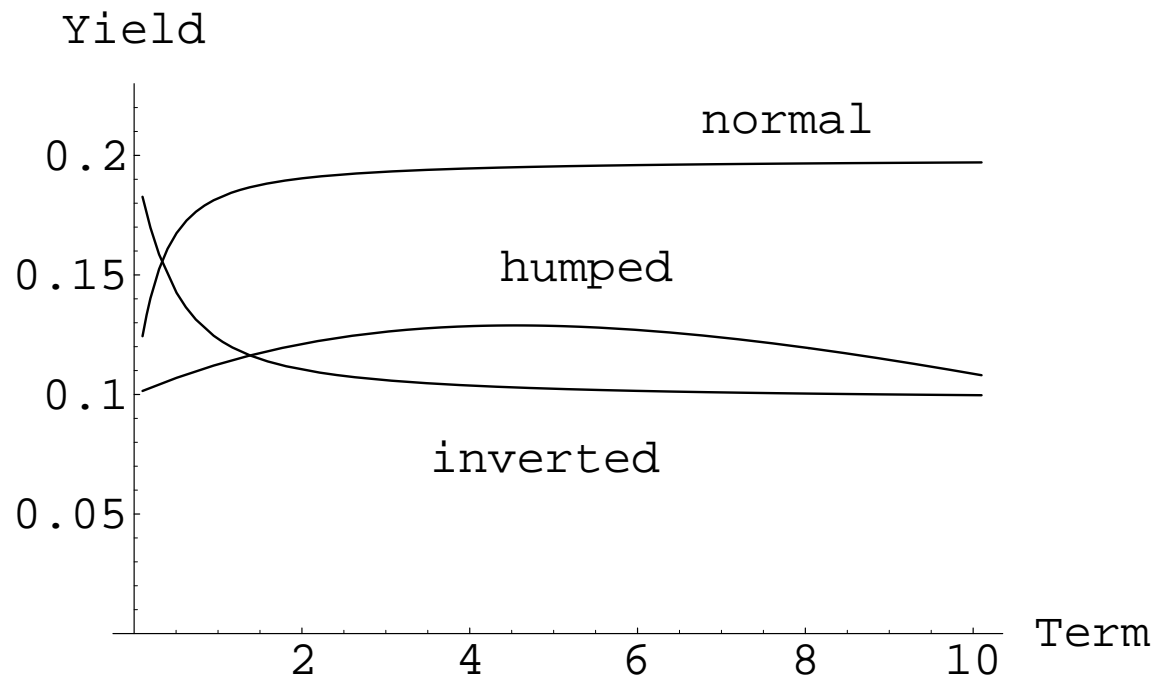
$$B(t, T) = \begin{cases} \frac{1 - e^{-\beta(T-t)}}{\beta} & \text{if } \beta \neq 0, \\ T - t & \text{if } \beta = 0. \end{cases}$$

## The Vasicek Model (concluded)

- If  $\beta = 0$ , then  $P$  goes to infinity as  $T \rightarrow \infty$ .
- Sensibly,  $P$  goes to zero as  $T \rightarrow \infty$  if  $\beta \neq 0$ .
- Even if  $\beta \neq 0$ ,  $P$  may exceed one for a finite  $T$ .
- The spot rate volatility structure is the curve

$$(\partial r(t, T) / \partial r) \sigma = \sigma B(t, T) / (T - t).$$

- When  $\beta > 0$ , the curve tends to decline with maturity.
- The speed of mean reversion,  $\beta$ , controls the shape of the curve.
- Indeed, higher  $\beta$  leads to greater attenuation of volatility with maturity.



## The Vasicek Model: Options on Zeros<sup>a</sup>

- Consider a European call with strike price  $X$  expiring at time  $T$  on a zero-coupon bond with par value \$1 and maturing at time  $s > T$ .
- Its price is given by

$$P(t, s) N(x) - X P(t, T) N(x - \sigma_v).$$

---

<sup>a</sup>Jamshidian (1989).

## The Vasicek Model: Options on Zeros (concluded)

- Above

$$\begin{aligned}x &\equiv \frac{1}{\sigma_v} \ln \left( \frac{P(t, s)}{P(t, T) X} \right) + \frac{\sigma_v}{2}, \\ \sigma_v &\equiv v(t, T) B(T, s), \\ v(t, T)^2 &\equiv \begin{cases} \frac{\sigma^2 [1 - e^{-2\beta(T-t)}]}{2\beta}, & \text{if } \beta \neq 0 \\ \sigma^2 (T - t), & \text{if } \beta = 0 \end{cases}.\end{aligned}$$

- By the put-call parity, the price of a European put is

$$XP(t, T) N(-x + \sigma_v) - P(t, s) N(-x).$$



## Binomial Vasicek

- Consider a binomial model for the short rate in the time interval  $[0, T]$  divided into  $n$  identical pieces.
- Let  $\Delta t \equiv T/n$  and

$$p(r) \equiv \frac{1}{2} + \frac{\beta(\mu - r) \sqrt{\Delta t}}{2\sigma}.$$

- The following binomial model converges to the Vasicek model,<sup>a</sup>

$$r(k + 1) = r(k) + \sigma \sqrt{\Delta t} \xi(k), \quad 0 \leq k < n.$$

---

<sup>a</sup>Nelson and Ramaswamy (1990).

## Binomial Vasicek (continued)

- Above,  $\xi(k) = \pm 1$  with

$$\text{Prob}[\xi(k) = 1] = \begin{cases} p(r(k)) & \text{if } 0 \leq p(r(k)) \leq 1 \\ 0 & \text{if } p(r(k)) < 0 \\ 1 & \text{if } 1 < p(r(k)) \end{cases} .$$

- Observe that the probability of an up move,  $p$ , is a decreasing function of the interest rate  $r$ .
- This is consistent with mean reversion.

## Binomial Vasicek (concluded)

- The rate is the same whether it is the result of an up move followed by a down move or a down move followed by an up move.
- The binomial tree combines.
- The key feature of the model that makes it happen is its *constant* volatility,  $\sigma$ .
- For a general process  $Y$  with nonconstant volatility, the resulting binomial tree may not combine, as we will see next.

## The Cox-Ingersoll-Ross Model<sup>a</sup>

- It is the following square-root short rate model:

$$dr = \beta(\mu - r) dt + \sigma\sqrt{r} dW. \quad (116)$$

- The diffusion differs from the Vasicek model by a multiplicative factor  $\sqrt{r}$ .
- The parameter  $\beta$  determines the speed of adjustment.
- The short rate can reach zero only if  $2\beta\mu < \sigma^2$ .
- See text for the bond pricing formula.

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<sup>a</sup>Cox, Ingersoll, and Ross (1985).

## Binomial CIR

- We want to approximate the short rate process in the time interval  $[0, T]$ .
- Divide it into  $n$  periods of duration  $\Delta t \equiv T/n$ .
- Assume  $\mu, \beta \geq 0$ .
- A direct discretization of the process is problematic because the resulting binomial tree will *not* combine.

## Binomial CIR (continued)

- Instead, consider the transformed process

$$x(r) \equiv 2\sqrt{r}/\sigma.$$

- It follows

$$dx = m(x) dt + dW,$$

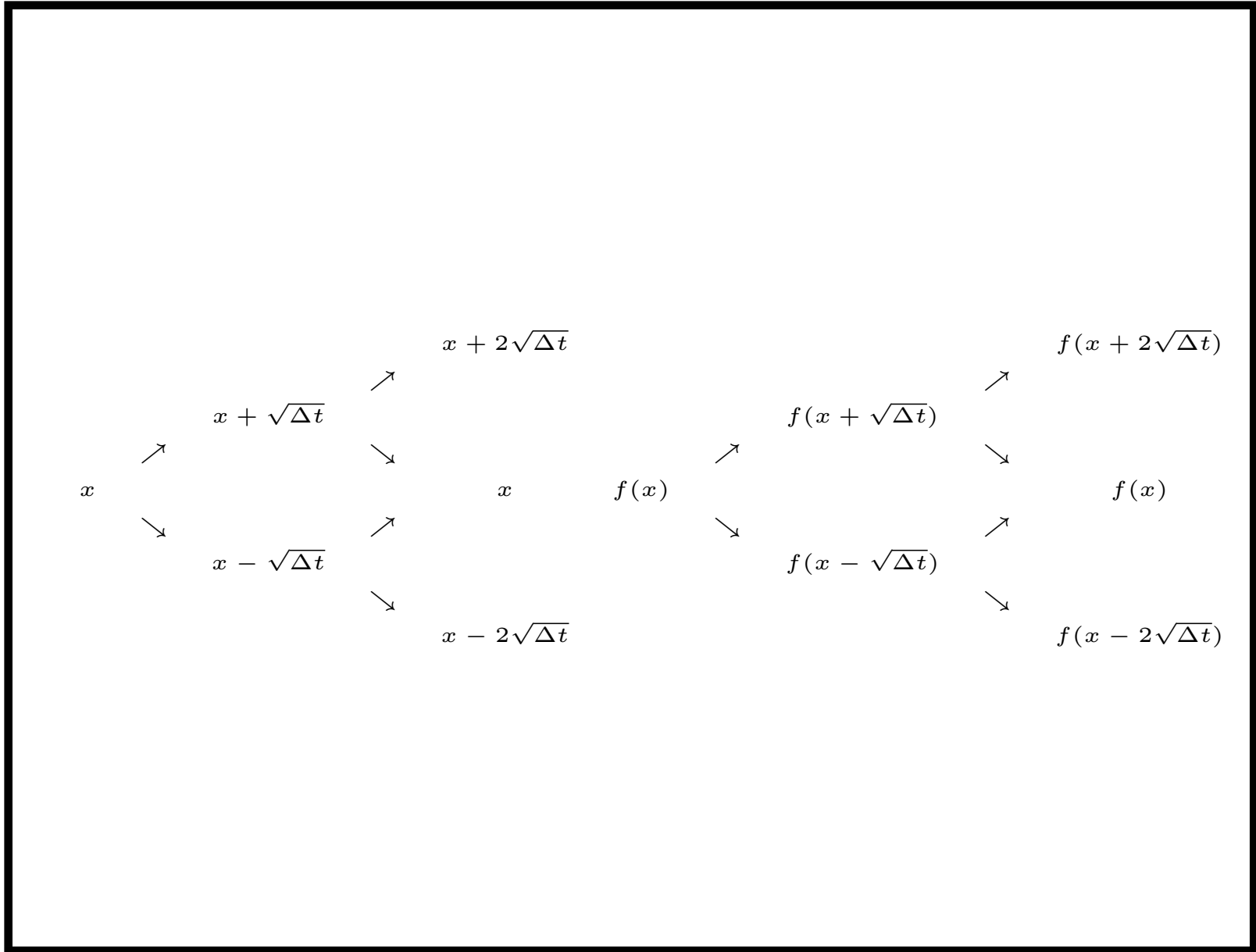
where

$$m(x) \equiv 2\beta\mu/(\sigma^2 x) - (\beta x/2) - 1/(2x).$$

- Since this new process has a constant volatility, its associated binomial tree combines.

## Binomial CIR (continued)

- Construct the combining tree for  $r$  as follows.
- First, construct a tree for  $x$ .
- Then transform each node of the tree into one for  $r$  via the inverse transformation  $r = f(x) \equiv x^2\sigma^2/4$  (p. 958).





## Binomial CIR (concluded)

- The probability of an up move at each node  $r$  is

$$p(r) \equiv \frac{\beta(\mu - r) \Delta t + r - r^-}{r^+ - r^-}. \quad (117)$$

- $r^+ \equiv f(x + \sqrt{\Delta t})$  denotes the result of an up move from  $r$ .
- $r^- \equiv f(x - \sqrt{\Delta t})$  the result of a down move.
- Finally, set the probability  $p(r)$  to one as  $r$  goes to zero to make the probability stay between zero and one.

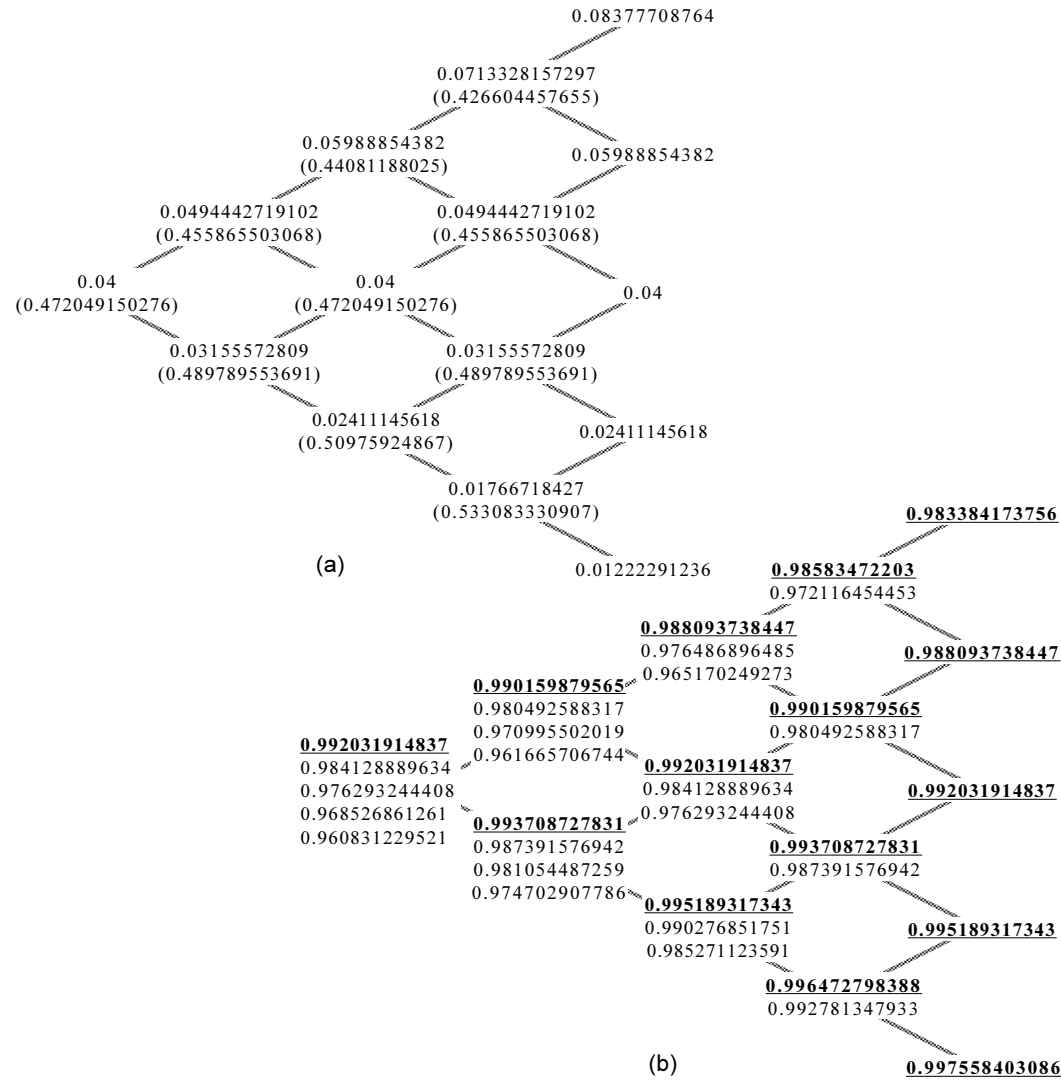
## Numerical Examples

- Consider the process,

$$0.2(0.04 - r) dt + 0.1\sqrt{r} dW,$$

for the time interval  $[0, 1]$  given the initial rate  $r(0) = 0.04$ .

- We shall use  $\Delta t = 0.2$  (year) for the binomial approximation.
- See p. 961(a) for the resulting binomial short rate tree with the up-move probabilities in parentheses.



## Numerical Examples (continued)

- Consider the node which is the result of an up move from the root.
- Since the root has  $x = 2\sqrt{r(0)}/\sigma = 4$ , this particular node's  $x$  value equals  $4 + \sqrt{\Delta t} = 4.4472135955$ .
- Use the inverse transformation to obtain the short rate  $x^2 \times (0.1)^2 / 4 \approx 0.0494442719102$ .

## Numerical Examples (concluded)

- Once the short rates are in place, computing the probabilities is easy.
- Note that the up-move probability decreases as interest rates increase and increases as interest rates decline.
- This phenomenon agrees with mean reversion.
- Convergence is quite good (see text).

## A General Method for Constructing Binomial Models<sup>a</sup>

- We are given a continuous-time process,

$$dy = \alpha(y, t) dt + \sigma(y, t) dW.$$

- Make sure the binomial model's drift and diffusion converge to the above process by setting the probability of an up move to

$$\frac{\alpha(y, t) \Delta t + y - y_d}{y_u - y_d}.$$

- Here  $y_u \equiv y + \sigma(y, t)\sqrt{\Delta t}$  and  $y_d \equiv y - \sigma(y, t)\sqrt{\Delta t}$  represent the two rates that follow the current rate  $y$ .
- The displacements are identical, at  $\sigma(y, t)\sqrt{\Delta t}$ .

---

<sup>a</sup>Nelson and Ramaswamy (1990).

## A General Method (continued)

- But the binomial tree may not combine as

$$\begin{aligned} & \sigma(y, t)\sqrt{\Delta t} - \sigma(y_u, t + \Delta t)\sqrt{\Delta t} \\ \neq & -\sigma(y, t)\sqrt{\Delta t} + \sigma(y_d, t + \Delta t)\sqrt{\Delta t} \end{aligned}$$

in general.

- When  $\sigma(y, t)$  is a constant independent of  $y$ , equality holds and the tree combines.

## A General Method (continued)

- To achieve this, define the transformation

$$x(y, t) \equiv \int^y \sigma(z, t)^{-1} dz.$$

- Then  $x$  follows

$$dx = m(y, t) dt + dW$$

for some  $m(y, t)$  (see text).

- The key is that the diffusion term is now a constant, and the binomial tree for  $x$  combines.
- The transformation that turns a 1-dim stochastic process into one with a constant diffusion term is unique.<sup>a</sup>

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<sup>a</sup>Chiu (R98723059) (2012).



## A General Method (concluded)

- The probability of an up move remains

$$\frac{\alpha(y(x, t), t) \Delta t + y(x, t) - y_d(x, t)}{y_u(x, t) - y_d(x, t)},$$

where  $y(x, t)$  is the inverse transformation of  $x(y, t)$  from  $x$  back to  $y$ .

- Note that  $y_u(x, t) \equiv y(x + \sqrt{\Delta t}, t + \Delta t)$  and  $y_d(x, t) \equiv y(x - \sqrt{\Delta t}, t + \Delta t)$ .

## Examples

- The transformation is

$$\int^r (\sigma\sqrt{z})^{-1} dz = 2\sqrt{r}/\sigma$$

for the CIR model.

- The transformation is

$$\int^S (\sigma z)^{-1} dz = (1/\sigma) \ln S$$

for the Black-Scholes model.

- The familiar binomial option pricing model in fact discretizes  $\ln S$  not  $S$ .

## On One-Factor Short Rate Models

- By using only the short rate, they ignore other rates on the yield curve.
- Such models also restrict the volatility to be a function of interest rate *levels* only.
- The prices of all bonds move in the same direction at the same time (their magnitudes may differ).
- The returns on all bonds thus become highly correlated.
- In reality, there seems to be a certain amount of independence between short- and long-term rates.

## On One-Factor Short Rate Models (continued)

- One-factor models therefore cannot accommodate nondegenerate correlation structures across maturities.
- Derivatives whose values depend on the correlation structure will be mispriced.
- The calibrated models may not generate term structures as concave as the data suggest.
- The term structure empirically changes in slope and curvature as well as makes parallel moves.
- This is inconsistent with the restriction that all segments of the term structure be perfectly correlated.

## On One-Factor Short Rate Models (concluded)

- Multi-factor models lead to families of yield curves that can take a greater variety of shapes and can better represent reality.
- But they are much harder to think about and work with.
- They also take much more computer time—the curse of dimensionality.
- These practical concerns limit the use of multifactor models to two-factor ones.

## Options on Coupon Bonds<sup>a</sup>

- Assume a one-factor short rate model.
- The price of a European option on a coupon bond can be calculated from those on zero-coupon bonds.
- Consider a European call expiring at time  $T$  on a bond with par value \$1.
- Let  $X$  denote the strike price.
- The bond has cash flows  $c_1, c_2, \dots, c_n$  at times  $t_1, t_2, \dots, t_n$ , where  $t_i > T$  for all  $i$ .

---

<sup>a</sup>Jamshidian (1989).

## Options on Coupon Bonds (continued)

- The payoff for the option is

$$\max \left( \sum_{i=1}^n c_i P(r(T), T, t_i) - X, 0 \right).$$

- At time  $T$ , there is a unique value  $r^*$  for  $r(T)$  that renders the coupon bond's price equal the strike price  $X$ .
- This  $r^*$  can be obtained by solving

$$X = \sum_{i=1}^n c_i P(r, T, t_i)$$

numerically for  $r$ .

## Options on Coupon Bonds (continued)

- The solution is unique for one-factor models whose bond price is a monotonically decreasing function of  $r$ .
- Let

$$X_i \equiv P(r^*, T, t_i),$$

the value at time  $T$  of a zero-coupon bond with par value \$1 and maturing at time  $t_i$  if  $r(T) = r^*$ .

- Note that  $P(r(T), T, t_i) \geq X_i$  if and only if  $r(T) \leq r^*$ .



## Options on Coupon Bonds (concluded)

- As  $X = \sum_i c_i X_i$ , the option's payoff equals

$$\begin{aligned} & \max \left( \sum_{i=1}^n c_i P(r(T), T, t_i) - \sum_i c_i X_i, 0 \right) \\ &= \sum_{i=1}^n c_i \times \max(P(r(T), T, t_i) - X_i, 0). \end{aligned}$$

- Thus the call is a package of  $n$  options on the underlying zero-coupon bond.
- Why can't we do the same thing for Asian options?<sup>a</sup>

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<sup>a</sup>Contributed by Mr. Yang, Jui-Chung (D97723002) on May 20, 2009.

# *No-Arbitrage Term Structure Models*

How much of the structure of our theories  
really tells us about things in nature,  
and how much do we contribute ourselves?  
— Arthur Eddington (1882–1944)

## Motivations

- Recall the difficulties facing equilibrium models mentioned earlier.
  - They usually require the estimation of the market price of risk.
  - They cannot fit the market term structure.
  - But consistency with the market is often mandatory in practice.

## No-Arbitrage Models<sup>a</sup>

- No-arbitrage models utilize the full information of the term structure.
- They accept the observed term structure as consistent with an unobserved and unspecified equilibrium.
- From there, arbitrage-free movements of interest rates or bond prices over time are modeled.
- By definition, the market price of risk must be reflected in the current term structure; hence the resulting interest rate process is risk-neutral.

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<sup>a</sup>Ho and Lee (1986). Thomas Lee is a “billionaire founder” of Thomas H. Lee Partners LP, according to *Bloomberg* on May 26, 2012.

## No-Arbitrage Models (concluded)

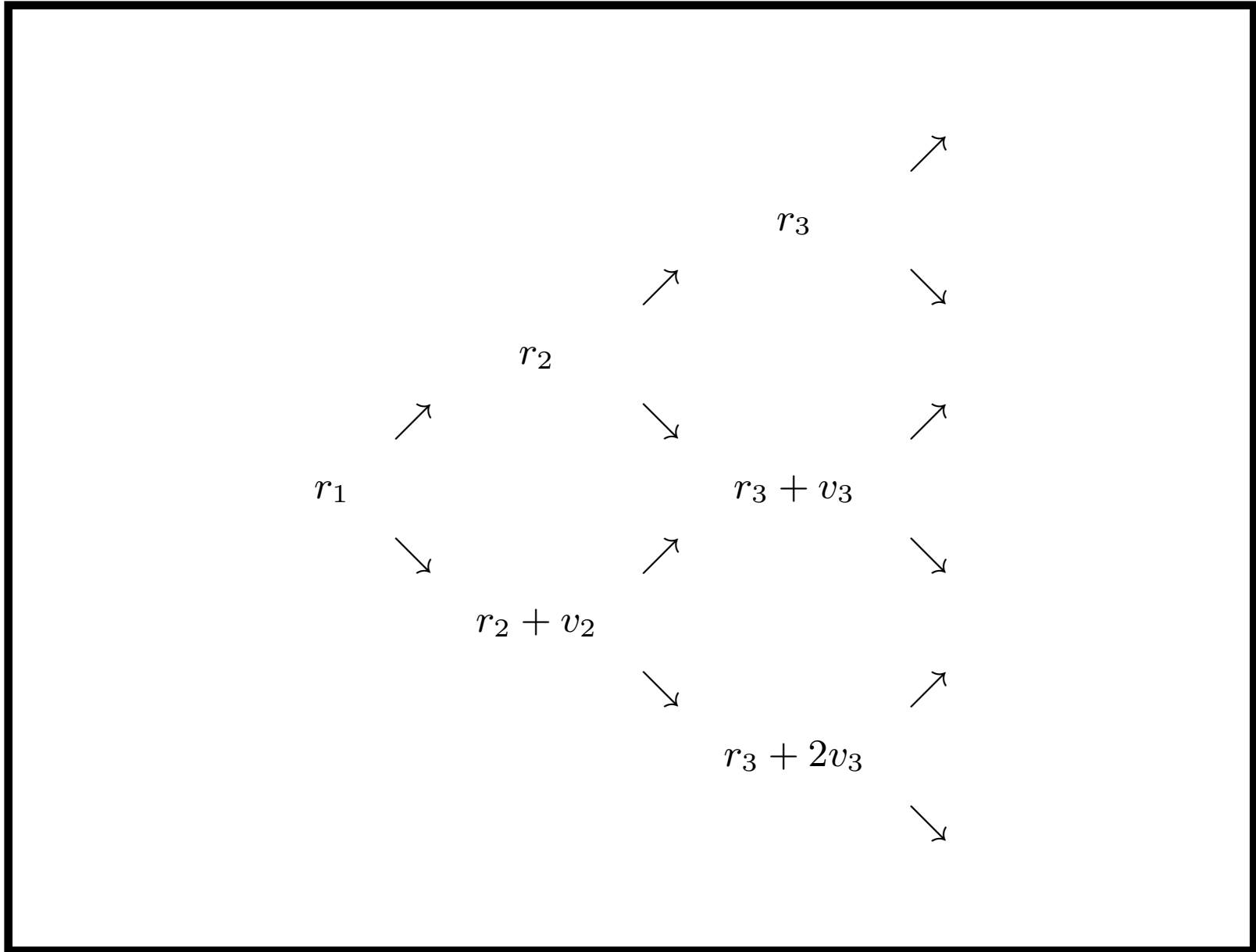
- No-arbitrage models can specify the dynamics of zero-coupon bond prices, forward rates, or the short rate.
- Bond price and forward rate models are usually non-Markovian (path dependent).
- In contrast, short rate models are generally constructed to be explicitly Markovian (path independent).
- Markovian models are easier to handle computationally.

## The Ho-Lee Model<sup>a</sup>

- The short rates at any given time are evenly spaced.
- Let  $p$  denote the risk-neutral probability that the short rate makes an up move.
- We shall adopt continuous compounding.

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<sup>a</sup>Ho and Lee (1986).





## The Ho-Lee Model (continued)

- The Ho-Lee model starts with zero-coupon bond prices  $P(t, t + 1), P(t, t + 2), \dots$  at time  $t$  identified with the root of the tree.

- Let the discount factors in the next period be

$$P_d(t + 1, t + 2), P_d(t + 1, t + 3), \dots \quad \text{if short rate moves down}$$

$$P_u(t + 1, t + 2), P_u(t + 1, t + 3), \dots \quad \text{if short rate moves up}$$

- By backward induction, it is not hard to see that for  $n \geq 2$ ,

$$P_u(t + 1, t + n) = P_d(t + 1, t + n) e^{-(v_2 + \dots + v_n)} \quad (118)$$

(see text).

## The Ho-Lee Model (continued)

- It is also not hard to check that the  $n$ -period zero-coupon bond has yields

$$y_d(n) \equiv -\frac{\ln P_d(t+1, t+n)}{n-1}$$

$$y_u(n) \equiv -\frac{\ln P_u(t+1, t+n)}{n-1} = y_d(n) + \frac{v_2 + \cdots + v_n}{n-1}$$

- The volatility of the yield to maturity for this bond is therefore

$$\begin{aligned}\kappa_n &\equiv \sqrt{py_u(n)^2 + (1-p)y_d(n)^2 - [py_u(n) + (1-p)y_d(n)]^2} \\ &= \sqrt{p(1-p)} (y_u(n) - y_d(n)) \\ &= \sqrt{p(1-p)} \frac{v_2 + \cdots + v_n}{n-1}.\end{aligned}$$

## The Ho-Lee Model (concluded)

- In particular, the short rate volatility is determined by taking  $n = 2$ :

$$\sigma = \sqrt{p(1-p)} v_2. \quad (119)$$

- The variance of the short rate therefore equals  $p(1-p)(r_u - r_d)^2$ , where  $r_u$  and  $r_d$  are the two successor rates.<sup>a</sup>

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<sup>a</sup>Contrast this with the lognormal model.

## The Ho-Lee Model: Volatility Term Structure

- The volatility term structure is composed of  $\kappa_2, \kappa_3, \dots$ 
  - It is independent of the  $r_i$ .
- It is easy to compute the  $v_i$ s from the volatility structure, and vice versa.
- The  $r_i$ s can be computed by forward induction.
- The volatility structure is supplied by the market.

## The Ho-Lee Model: Bond Price Process

- In a risk-neutral economy, the initial discount factors satisfy

$$P(t, t+n) = (pP_u(t+1, t+n) + (1-p)P_d(t+1, t+n))P(t, t+1)$$

- Combine the above with Eq. (118) on p. 983 and assume  $p = 1/2$  to obtain<sup>a</sup>

$$P_d(t+1, t+n) = \frac{P(t, t+n)}{P(t, t+1)} \frac{2 \times \exp[v_2 + \cdots + v_n]}{1 + \exp[v_2 + \cdots + v_n]}, \quad (120)$$

$$P_u(t+1, t+n) = \frac{P(t, t+n)}{P(t, t+1)} \frac{2}{1 + \exp[v_2 + \cdots + v_n]}. \quad (120')$$

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<sup>a</sup>In the limit, only the volatility matters.

## The Ho-Lee Model: Bond Price Process (concluded)

- The bond price tree combines.
- Suppose all  $v_i$  equal some constant  $v$  and  $\delta \equiv e^v > 0$ .
- Then

$$P_d(t+1, t+n) = \frac{P(t, t+n)}{P(t, t+1)} \frac{2\delta^{n-1}}{1 + \delta^{n-1}},$$
$$P_u(t+1, t+n) = \frac{P(t, t+n)}{P(t, t+1)} \frac{2}{1 + \delta^{n-1}}.$$

- Short rate volatility  $\sigma$  equals  $v/2$  by Eq. (119) on p. 985.
- Price derivatives by taking expectations under the risk-neutral probability.

## The Ho-Lee Model: Yields and Their Covariances

- The one-period rate of return of an  $n$ -period zero-coupon bond is

$$r(t, t + n) \equiv \ln \left( \frac{P(t + 1, t + n)}{P(t, t + n)} \right).$$

- Its value is either  $\ln \frac{P_d(t+1, t+n)}{P(t, t+n)}$  or  $\ln \frac{P_u(t+1, t+n)}{P(t, t+n)}$ .
- Thus the variance of return is

$$\text{Var}[r(t, t + n)] = p(1 - p)((n - 1)v)^2 = (n - 1)^2 \sigma^2.$$

## The Ho-Lee Model: Yields and Their Covariances (concluded)

- The covariance between  $r(t, t + n)$  and  $r(t, t + m)$  is  $(n - 1)(m - 1) \sigma^2$  (see text).
- As a result, the correlation between any two one-period rates of return is unity.
- Strong correlation between rates is inherent in all one-factor Markovian models.



## The Ho-Lee Model: Short Rate Process

- The continuous-time limit of the Ho-Lee model is

$$dr = \theta(t) dt + \sigma dW.$$

- This is Vasicek's model with the mean-reverting drift replaced by a deterministic, time-dependent drift.
- A nonflat term structure of volatilities can be achieved if the short rate volatility is also made time varying, i.e.,  
$$dr = \theta(t) dt + \sigma(t) dW.$$
- This corresponds to the discrete-time model in which  $v_i$  are not all identical.

## The Ho-Lee Model: Some Problems

- Future (nominal) interest rates may be negative.
- The short rate volatility is independent of the rate level.

## Problems with No-Arbitrage Models in General

- Interest rate movements should reflect shifts in the model's state variables (factors) not its parameters.
- Model *parameters*, such as the drift  $\theta(t)$  in the continuous-time Ho-Lee model, should be stable over time.
- But in practice, no-arbitrage models capture yield curve shifts through the recalibration of parameters.
  - A new model is thus born everyday.

## Problems with No-Arbitrage Models in General (concluded)

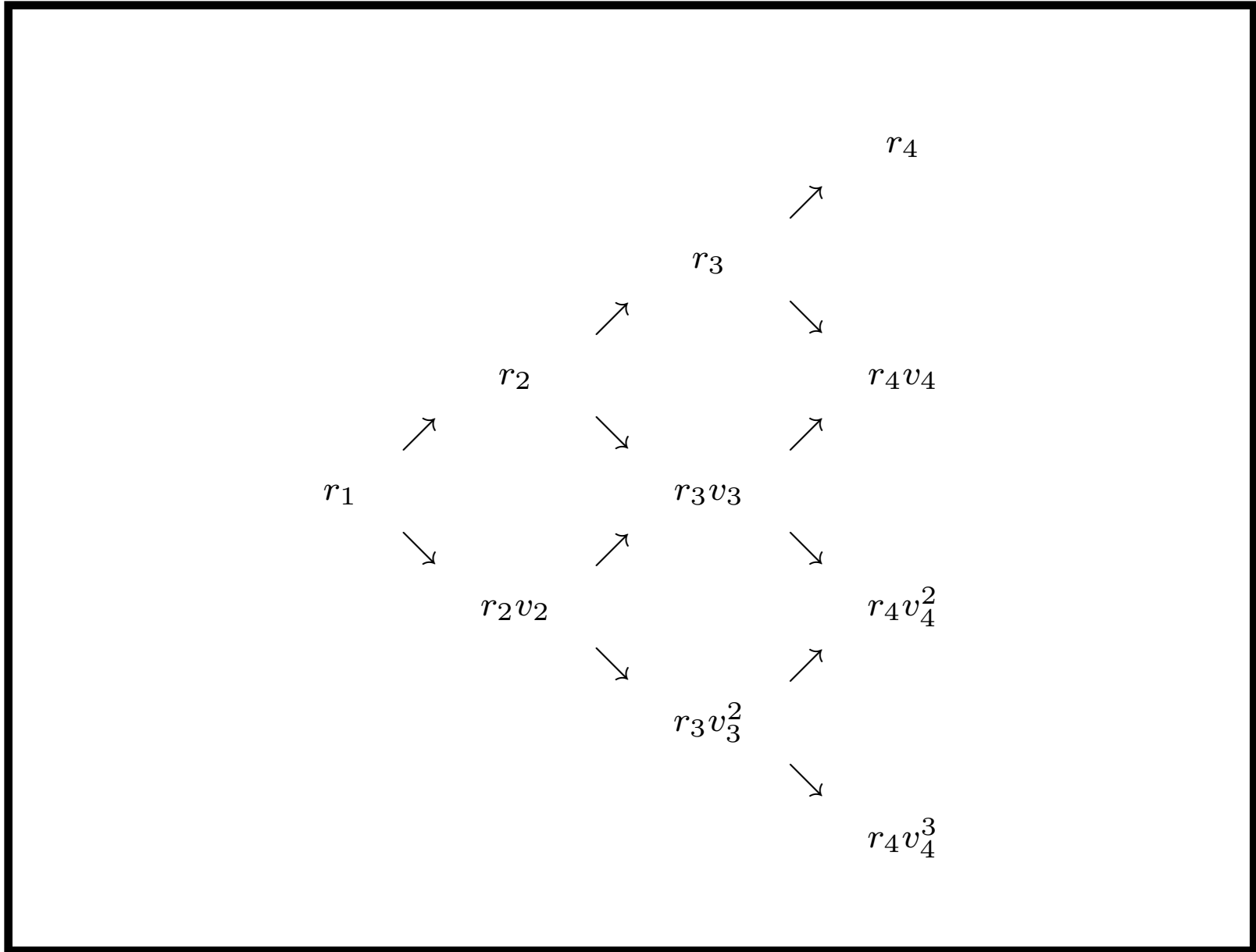
- This in effect says the model estimated at some time does not describe the term structure of interest rates and their volatilities at other times.
- Consequently, a model's intertemporal behavior is suspect, and using it for hedging and risk management may be unreliable.

## The Black-Derman-Toy Model<sup>a</sup>

- This model is extensively used by practitioners.
- The BDT short rate process is the lognormal binomial interest rate process described on pp. 834ff (repeated on next page).
- The volatility structure is given by the market.
- From it, the short rate volatilities (thus  $v_i$ ) are determined together with  $r_i$ .

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<sup>a</sup>Black, Derman, and Toy (BDT) (1990), but essentially finished in 1986 according to Mehrling (2005).



## The Black-Derman-Toy Model (concluded)

- Our earlier binomial interest rate tree, in contrast, assumes  $v_i$  are given a priori.
  - A related model of Salomon Brothers takes  $v_i$  to be a given constant.<sup>a</sup>
- Lognormal models preclude negative short rates.

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<sup>a</sup>Tuckman (2002).

## The BDT Model: Volatility Structure

- The volatility structure defines the yield volatilities of zero-coupon bonds of various maturities.
- Let the yield volatility of the  $i$ -period zero-coupon bond be denoted by  $\kappa_i$ .
- $P_u$  is the price of the  $i$ -period zero-coupon bond one period from now if the short rate makes an up move.
- $P_d$  is the price of the  $i$ -period zero-coupon bond one period from now if the short rate makes a down move.



## The BDT Model: Volatility Structure (concluded)

- Corresponding to these two prices are the following yields to maturity,

$$y_u \equiv P_u^{-1/(i-1)} - 1,$$

$$y_d \equiv P_d^{-1/(i-1)} - 1.$$

- The yield volatility is defined as

$$\kappa_i \equiv \frac{\ln(y_u/y_d)}{2}.$$

## The BDT Model: Calibration

- The inputs to the BDT model are riskless zero-coupon bond yields and their volatilities.
- For economy of expression, all numbers are period based.
- Suppose inductively that we have calculated

$$(r_1, v_1), (r_2, v_2), \dots, (r_{i-1}, v_{i-1}).$$

- They define the binomial tree up to period  $i - 1$ .
- We now proceed to calculate  $r_i$  and  $v_i$  to extend the tree to period  $i$ .

## The BDT Model: Calibration (continued)

- Assume the price of the  $i$ -period zero can move to  $P_u$  or  $P_d$  one period from now.
- Let  $y$  denote the current  $i$ -period spot rate, which is known.
- In a risk-neutral economy,

$$\frac{P_u + P_d}{2(1 + r_1)} = \frac{1}{(1 + y)^i}. \quad (121)$$

- Obviously,  $P_u$  and  $P_d$  are functions of the unknown  $r_i$  and  $v_i$ .

## The BDT Model: Calibration (continued)

- Viewed from now, the future  $(i - 1)$ -period spot rate at time 1 is uncertain.
- Recall that  $y_u$  and  $y_d$  represent the spot rates at the up node and the down node, respectively (p. 999).
- With  $\kappa^2$  denoting their variance, we have

$$\kappa_i = \frac{1}{2} \ln \left( \frac{P_u^{-1/(i-1)} - 1}{P_d^{-1/(i-1)} - 1} \right). \quad (122)$$

## The BDT Model: Calibration (continued)

- We will employ forward induction to derive a quadratic-time calibration algorithm.<sup>a</sup>
- Recall that forward induction inductively figures out, by moving *forward* in time, how much \$1 at a node contributes to the price (review p. 860(a)).
- This number is called the state price and is the price of the claim that pays \$1 at that node and zero elsewhere.

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<sup>a</sup>Chen (R84526007) and Lyuu (1997); Lyuu (1999).

## The BDT Model: Calibration (continued)

- Let the unknown baseline rate for period  $i$  be  $r_i = r$ .
- Let the unknown multiplicative ratio be  $v_i = v$ .
- Let the state prices at time  $i - 1$  be  $P_1, P_2, \dots, P_i$ , corresponding to rates  $r, rv, \dots, rv^{i-1}$  for period  $i$ , respectively.
- One dollar at time  $i$  has a present value of

$$f(r, v) \equiv \frac{P_1}{1 + r} + \frac{P_2}{1 + rv} + \frac{P_3}{1 + rv^2} + \dots + \frac{P_i}{1 + rv^{i-1}}.$$

## The BDT Model: Calibration (continued)

- The yield volatility is

$$g(r, v) \equiv \frac{1}{2} \ln \left( \frac{\left( \frac{P_{u,1}}{1+rv} + \frac{P_{u,2}}{1+rv^2} + \dots + \frac{P_{u,i-1}}{1+rv^{i-1}} \right)^{-1/(i-1)} - 1}{\left( \frac{P_{d,1}}{1+r} + \frac{P_{d,2}}{1+rv} + \dots + \frac{P_{d,i-1}}{1+rv^{i-2}} \right)^{-1/(i-1)} - 1} \right).$$

- Above,  $P_{u,1}, P_{u,2}, \dots$  denote the state prices at time  $i - 1$  of the subtree rooted at the up node (like  $r_2 v_2$  on p. 996).
- And  $P_{d,1}, P_{d,2}, \dots$  denote the state prices at time  $i - 1$  of the subtree rooted at the down node (like  $r_2$  on p. 996).

## The BDT Model: Calibration (concluded)

- Note that every node maintains 3 state prices.
- Now solve

$$\begin{aligned}f(r, v) &= \frac{1}{(1 + y)^i}, \\g(r, v) &= \kappa_i,\end{aligned}$$

for  $r = r_i$  and  $v = v_i$ .

- This  $O(n^2)$ -time algorithm appears in the text.



## The BDT Model: Continuous-Time Limit

- The continuous-time limit of the BDT model is

$$d \ln r = \left( \theta(t) + \frac{\sigma'(t)}{\sigma(t)} \ln r \right) dt + \sigma(t) dW.$$

- The short rate volatility clearly should be a declining function of time for the model to display mean reversion.
  - That makes  $\sigma'(t) < 0$ .
- In particular, constant volatility will not attain mean reversion.