#### Gamma

• The finite-difference formula for gamma is

$$e^{-r\tau} E\left[\frac{P(S+\epsilon) - 2 \times P(S) + P(S-\epsilon)}{\epsilon^2}\right]$$

• For a correlation option with multiple underlying assets, the finite-difference formula for the cross gammas  $\partial^2 P(S_1, S_2, \dots)/(\partial S_1 \partial S_2)$  is:

$$e^{-r\tau} E\left[\frac{P(S_1+\epsilon_1, S_2+\epsilon_2) - P(S_1-\epsilon_1, S_2+\epsilon_2)}{4\epsilon_1\epsilon_2}\right]$$

$$\frac{-P(S_1+\epsilon_1,S_2-\epsilon_2)+P(S_1-\epsilon_1,S_2-\epsilon_2)}{2}$$

## Gamma (concluded)

- Choosing an  $\epsilon$  of the right magnitude can be challenging.
  - If  $\epsilon$  is too large, inaccurate Greeks result.
  - If  $\epsilon$  is too small, unstable Greeks result.

## Biases in Pricing Continuously Monitored Options with Monte Carlo

- We are asked to price a continuously monitored up-and-out call with barrier H.
- The Monte Carlo method samples the stock price at n discrete time points  $t_1, t_2, \ldots, t_n$ .
- A sample path  $S(t_0), S(t_1), \ldots, S(t_n)$  is produced.
  - Here,  $t_0 = 0$  is the current time, and  $t_n = T$  is the expiration time of the option.
- If all of the sampled prices are below the barrier, this sample path pays  $\max(S(t_n) X, 0)$ .

## Biases in Pricing Continuously Monitored Options with Monte Carlo (continued)

- Repeating these steps and averaging the payoffs yield a Monte Carlo estimate.
- This estimate is biased.
  - Suppose none of the sampled prices on a sample path equals or exceeds the barrier H.
  - It remains possible for the continuous sample path that passes through them to hit the barrier between sampled time points (see plot on next page).



## Biases in Pricing Continuously Monitored Options with Monte Carlo (concluded)

- The bias can certainly be lowered by increasing the number of observations along the sample path.
- However, even daily sampling may not suffice.
- The computational cost also rises as a result.

#### Brownian Bridge Approach to Pricing Barrier Options

- We desire an unbiased estimate efficiently.
- So the above-mentioned payoff should be multiplied by the probability p that a continuous sample path does not hit the barrier conditional on the sampled prices.
- This methodology is called the Brownian bridge approach.
- Formally, we have

 $p \equiv \operatorname{Prob}[S(t) < H, 0 \le t \le T | S(t_0), S(t_1), \dots, S(t_n)].$ 

• As a barrier is hit over a time interval if and only if the maximum stock price over that period is at least H,

$$p = \operatorname{Prob}\left[\max_{0 \le t \le T} S(t) < H \,|\, S(t_0), S(t_1), \dots, S(t_n)\right].$$

• Luckily, the conditional distribution of the maximum over a time interval given the beginning and ending stock prices is known.

**Lemma 19** Assume S follows  $dS/S = \mu dt + \sigma dW$  and define  $\zeta(x) \equiv \exp\left[-\frac{2\ln(x/S(t))\ln(x/S(t+\Delta t))}{\sigma^2 \Delta t}\right].$ (1) If  $H > \max(S(t), S(t + \Delta t))$ , then  $\operatorname{Prob}\left[\max_{t < u < t + \Delta t} S(u) < H \mid S(t), S(t + \Delta t)\right] = 1 - \zeta(H).$ (2) If  $h < \min(S(t), S(t + \Delta t))$ , then  $\operatorname{Prob}\left[\min_{t \le u \le t + \Delta t} S(u) > h \mid S(t), S(t + \Delta t)\right] = 1 - \zeta(h).$ 

- Lemma 19 gives the probability that the barrier is not hit in a time interval, given the starting and ending stock prices.
- For our up-and-out call, choose n = 1.
- As a result,

$$p = \begin{cases} 1 - \exp\left[-\frac{2\ln(H/S(0))\ln(H/S(T))}{\sigma^2 T}\right], & \text{if } H > \max(S(0), S(T)), \\ 0, & \text{otherwise.} \end{cases}$$

1: C := 0;2: for i = 1, 2, 3, ..., m do 3:  $P := S \times e^{(r-q-\sigma^2/2)T+\sigma\sqrt{T} \xi()};$ 4: if (S < H and P < H) or (S > H and P > H) then 5:  $C := C + \max(P-X, 0) \times \left\{ 1 - \exp\left[ -\frac{2\ln(H/S) \times \ln(H/P)}{\sigma^2 T} \right] \right\};$ 6: end if 7: end for 8: return  $Ce^{-rT}/m;$ 

- The idea can be generalized.
- For example, we can handle more complex barrier options.
- Consider an up-and-out call with barrier  $H_i$  for the time interval  $(t_i, t_{i+1}], 0 \le i < n$ .
- This option thus contains n barriers.
- It is a simple matter of multiplying the probabilities for the *n* time intervals properly to obtain the desired probability adjustment term.

#### Variance Reduction: Antithetic Variates

- We are interested in estimating  $E[g(X_1, X_2, ..., X_n)]$ , where  $X_1, X_2, ..., X_n$  are independent.
- Let  $Y_1$  and  $Y_2$  be random variables with the same distribution as  $g(X_1, X_2, \ldots, X_n)$ .
- Then

$$\operatorname{Var}\left[\frac{Y_1 + Y_2}{2}\right] = \frac{\operatorname{Var}[Y_1]}{2} + \frac{\operatorname{Cov}[Y_1, Y_2]}{2}$$

-  $\operatorname{Var}[Y_1]/2$  is the variance of the Monte Carlo method with two (independent) replications.

• The variance  $\operatorname{Var}[(Y_1 + Y_2)/2]$  is smaller than  $\operatorname{Var}[Y_1]/2$  when  $Y_1$  and  $Y_2$  are negatively correlated.

#### Variance Reduction: Antithetic Variates (continued)

- For each simulated sample path X, a second one is obtained by reusing the random numbers on which the first path is based.
- This yields a second sample path Y.
- Two estimates are then obtained: One based on X and the other on Y.
- If N independent sample paths are generated, the antithetic-variates estimator averages over 2Nestimates.

Variance Reduction: Antithetic Variates (continued)

- Consider process  $dX = a_t dt + b_t \sqrt{dt} \xi$ .
- Let g be a function of n samples  $X_1, X_2, \ldots, X_n$  on the sample path.
- We are interested in  $E[g(X_1, X_2, \ldots, X_n)].$
- Suppose one simulation run has realizations
   ξ<sub>1</sub>, ξ<sub>2</sub>,..., ξ<sub>n</sub> for the normally distributed fluctuation term ξ.
- This generates samples  $x_1, x_2, \ldots, x_n$ .
- The estimate is then  $g(\boldsymbol{x})$ , where  $\boldsymbol{x} \equiv (x_1, x_2 \dots, x_n)$ .

#### Variance Reduction: Antithetic Variates (concluded)

- We do not sample n more numbers from  $\xi$  for the second estimate.
- The antithetic-variates method computes  $g(\mathbf{x}')$  from the sample path  $\mathbf{x}' \equiv (x'_1, x'_2 \dots, x'_n)$  generated by  $-\xi_1, -\xi_2, \dots, -\xi_n$ .
- We then output  $(g(\boldsymbol{x}) + g(\boldsymbol{x}'))/2$ .
- Repeat the above steps for as many times as required by accuracy.

#### Variance Reduction: Conditioning

- We are interested in estimating E[X].
- Suppose here is a random variable Z such that E[X | Z = z] can be efficiently and precisely computed.
- E[X] = E[E[X|Z]] by the law of iterated conditional expectations.
- Hence the random variable E[X | Z] is also an unbiased estimator of E[X].

### Variance Reduction: Conditioning (concluded)

- As  $\operatorname{Var}[E[X | Z]] \leq \operatorname{Var}[X], E[X | Z]$  has a smaller variance than observing X directly.
- First obtain a random observation z on Z.
- Then calculate E[X | Z = z] as our estimate.
  - There is no need to resort to simulation in computing E[X | Z = z].
- The procedure can be repeated a few times to reduce the variance.

#### Control Variates

- Use the analytic solution of a similar yet simpler problem to improve the solution.
- Suppose we want to estimate E[X] and there exists a random variable Y with a known mean  $\mu \equiv E[Y]$ .
- Then  $W \equiv X + \beta(Y \mu)$  can serve as a "controlled" estimator of E[X] for any constant  $\beta$ .
  - $-\beta$  can scale the deviation  $Y \mu$  to arrive at an adjustment for X.
  - However  $\beta$  is chosen, W remains an unbiased estimator of E[X] as

$$E[W] = E[X] + \beta E[Y - \mu] = E[X].$$

### Control Variates (continued)

• Note that

$$\operatorname{Var}[W] = \operatorname{Var}[X] + \beta^{2} \operatorname{Var}[Y] + 2\beta \operatorname{Cov}[X, Y],$$
(64)

• Hence W is less variable than X if and only if

$$\beta^2 \operatorname{Var}[Y] + 2\beta \operatorname{Cov}[X, Y] < 0.$$
(65)

• The success of the scheme clearly depends on both  $\beta$ and the choice of Y.

#### Control Variates (concluded)

- For example, arithmetic average-rate options can be priced by choosing Y to be the otherwise identical geometric average-rate option's price and  $\beta = -1$ .
- This approach is much more effective than the antithetic-variates method.

## Choice of Y

- In general, the choice of Y is ad hoc, and experiments must be performed to confirm the wisdom of the choice.
- Try to match calls with calls and puts with puts.<sup>a</sup>
- On many occasions, Y is a discretized version of the derivative that gives  $\mu$ .
  - Discretely monitored geometric average-rate option vs. the continuously monitored geometric average-rate option given by formulas (29) on p. 327.
- For some choices, the discrepancy can be significant, such as the lookback option.<sup>b</sup>

<sup>a</sup>Contributed by Ms. Teng, Huei-Wen (**R91723054**) on May 25, 2004. <sup>b</sup>Contributed by Mr. Tsai, Hwai (**R92723049**) on May 12, 2004.

## Optimal Choice of $\beta$

• Equation (64) on p. 618 is minimized when

$$\beta = -\operatorname{Cov}[X, Y] / \operatorname{Var}[Y],$$

which was called beta earlier in the book.

• For this specific  $\beta$ ,

$$\operatorname{Var}[W] = \operatorname{Var}[X] - \frac{\operatorname{Cov}[X,Y]^2}{\operatorname{Var}[Y]} = \left(1 - \rho_{X,Y}^2\right) \operatorname{Var}[X],$$

where  $\rho_{X,Y}$  is the correlation between X and Y.

• The stronger X and Y are correlated, the greater the reduction in variance.

## Optimal Choice of $\beta$ (continued)

- For example, if this correlation is nearly perfect (±1), we could control X almost exactly, eliminating practically all of its variance.
- Typically, neither  $\operatorname{Var}[Y]$  nor  $\operatorname{Cov}[X, Y]$  is known.
- Therefore, we cannot obtain the maximum reduction in variance.
- We can guess these values and hope that the resulting W does indeed have a smaller variance than X.
- A second possibility is to use the simulated data to estimate these quantities.

## Optimal Choice of $\beta$ (concluded)

- Observe that  $-\beta$  has the same sign as the correlation between X and Y.
- Hence, if X and Y are positively correlated,  $\beta < 0$ , then X is adjusted downward whenever  $Y > \mu$  and upward otherwise.
- The opposite is true when X and Y are negatively correlated, in which case  $\beta > 0$ .

#### Problems with the Monte Carlo Method

- The error bound is only probabilistic.
- The probabilistic error bound of  $\sqrt{N}$  does not benefit from regularity of the integrand function.
- The requirement that the points be independent random samples are wasteful because of clustering.
- In reality, pseudorandom numbers generated by completely deterministic means are used.
- Monte Carlo simulation exhibits a great sensitivity on the seed of the pseudorandom-number generator.