

A Technical Introduction to Gaussian Process Regression

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1 Definition

In regression problems, we are given a sample $S = \{(\mathbf{x}_1, y_1), (\mathbf{x}_2, y_2), \dots, (\mathbf{x}_l, y_l)\}$ in which \mathbf{x}_i denotes the i th observation and y_i is the corresponding target value. The relationship between \mathbf{x}_i and y_i is formulated as

$$y_i = f(\mathbf{x}_i) + \epsilon(\mathbf{x}_i),$$

that is, a function f maps the input vector \mathbf{x}_i to the true target, which, being corrupted by noise $\epsilon(\mathbf{x}_i)$, is measured as y_i . Gaussian Process Regression (GPR) is a non-parametric model that assumes

$$\mathbf{f} \equiv [f(\mathbf{x}_1), f(\mathbf{x}_2), \dots, f(\mathbf{x}_l)]^T \sim N(\mathbf{0}, K), \quad (1)$$

where K is the covariance matrix whose (i, j) th element is given by a kernel function $\mathcal{K}(\mathbf{x}_i, \mathbf{x}_j)$, and

$$(\mathbf{y}|\mathbf{f}) = ([y_1, y_2, \dots, y_l]^T | \mathbf{f}) \sim N(\mathbf{f}, \sigma^2 I), \quad (2)$$

which means the noise follows a zero-mean and independent joint Gaussian distribution. Note that (2) also implies \mathbf{y} 's conditional independence of $\{\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_l\}$ given \mathbf{f} . For a new instance \mathbf{x}^* , the goal is to estimate $P(f(\mathbf{x}^*)|\mathbf{x}^*, S)$. In the sequel we assume K to be invertible.

2 Derivation of The Predictive Distribution

For convenience, we denote $f(\mathbf{x}^*)$ as f^* . Following a standard Bayesian approach, we write

$$\begin{aligned} P(f^*|\mathbf{x}^*, S) &= \int P(f^*, \mathbf{f}|\mathbf{x}^*, S) d\mathbf{f} \\ &= \int P(f^*|\mathbf{f}, \mathbf{x}^*, S) P(\mathbf{f}|\mathbf{x}^*, S) d\mathbf{f}. \end{aligned} \quad (3)$$

We then derive $P(f^*|\mathbf{f}, \mathbf{x}^*, S)$ in Section 2.1, $P(\mathbf{f}|\mathbf{x}^*, S)$ in Section 2.2, and finally $P(f^*|\mathbf{x}^*, S)$ in Section 2.3.

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2.1 Derivation of $P(f^*|\mathbf{f}, \mathbf{x}^*, S)$

Define $\mathbf{k} \equiv [\mathcal{K}(\mathbf{x}^*, \mathbf{x}_1), \mathcal{K}(\mathbf{x}^*, \mathbf{x}_2), \dots, \mathcal{K}(\mathbf{x}^*, \mathbf{x}_l)]^T$. Then the joint distribution of $[\mathbf{f} \ f^*]^T$ is

$$\begin{bmatrix} \mathbf{f} \\ f^* \end{bmatrix} \sim N\left(\mathbf{0}, \begin{bmatrix} K & \mathbf{k} \\ \mathbf{k}^T & \mathcal{K}(\mathbf{x}^*, \mathbf{x}^*) \end{bmatrix}\right). \quad (4)$$

Since conditions on \mathbf{x}^* and S are embedded in the covariance matrix in (4), $P(f^*|\mathbf{f}, \mathbf{x}^*, S)$ is equivalent to

$$P(f^*|\mathbf{f}) = \frac{P(f^*, \mathbf{f})}{P(\mathbf{f})}. \quad (5)$$

Let

$$\begin{bmatrix} A & \mathbf{b} \\ \mathbf{b}^T & c \end{bmatrix} \equiv \begin{bmatrix} K & \mathbf{k} \\ \mathbf{k}^T & \mathcal{K}(\mathbf{x}^*, \mathbf{x}^*) \end{bmatrix}^{-1}, \quad (6)$$

then according to (5)

$$\begin{aligned} P(f^*|\mathbf{f}) &\propto \exp\left(-\frac{1}{2} \begin{bmatrix} \mathbf{f} \\ f^* \end{bmatrix}^T \begin{bmatrix} A & \mathbf{b} \\ \mathbf{b}^T & c \end{bmatrix} \begin{bmatrix} \mathbf{f} \\ f^* \end{bmatrix} + \frac{1}{2} \mathbf{f}^T K^{-1} \mathbf{f}\right) \\ &= \exp\left(-\frac{1}{2} \left(c(f^*)^2 + 2(\mathbf{b}^T \mathbf{f}) f^* + \mathbf{f}^T A \mathbf{f}\right) + \frac{1}{2} \mathbf{f}^T K^{-1} \mathbf{f}\right) \\ &\propto \exp\left(-\frac{1}{2c^{-1}} \left(f^* + \frac{\mathbf{b}^T \mathbf{f}}{c}\right)^2\right). \end{aligned} \quad (7)$$

From (6), we have

$$K\mathbf{b} + c\mathbf{k} = \mathbf{0} \text{ and } \mathbf{k}^T \mathbf{b} + c\mathcal{K}(\mathbf{x}^*, \mathbf{x}^*) = 1,$$

which in turn implies

$$\mathbf{b} = -\frac{K^{-1}\mathbf{k}}{\mathcal{K}(\mathbf{x}^*, \mathbf{x}^*) - \mathbf{k}^T K^{-1} \mathbf{k}}, \quad (8)$$

$$c = \frac{1}{\mathcal{K}(\mathbf{x}^*, \mathbf{x}^*) - \mathbf{k}^T K^{-1} \mathbf{k}}. \quad (9)$$

Plugging (8) and (9) into (7) yields

$$P(f^*|\mathbf{f}) \propto \exp\left(-\frac{(f^* - \mathbf{k}^T K^{-1} \mathbf{f})^2}{2(\mathcal{K}(\mathbf{x}^*, \mathbf{x}^*) - \mathbf{k}^T K^{-1} \mathbf{k})}\right),$$

that is,

$$(f^*|\mathbf{f}) \sim N\left(\mathbf{k}^T K^{-1} \mathbf{f}, \mathcal{K}(\mathbf{x}^*, \mathbf{x}^*) - \mathbf{k}^T K^{-1} \mathbf{k}\right). \quad (10)$$

2.2 Derivation of $P(\mathbf{f}|\mathbf{x}^*, S)$

Since \mathbf{f} does not depend on \mathbf{x}^* , it suffices to derive $P(\mathbf{f}|S)$. Again, we use standard Bayesian techniques to have

$$\begin{aligned} P(\mathbf{f}|S) &\propto P(S|\mathbf{f})P(\mathbf{f}) \\ &= P(\mathbf{y}, \mathbf{x}|\mathbf{f})P(\mathbf{f}) \\ &= P(\mathbf{y}|\mathbf{x}, \mathbf{f})P(\mathbf{x}|\mathbf{f})P(\mathbf{f}) \end{aligned} \tag{11}$$

$$\propto P(\mathbf{y}|\mathbf{f})P(\mathbf{f}). \tag{12}$$

From (11) to (12), we use the fact that $P(\mathbf{y}|\mathbf{x}, \mathbf{f}) = P(\mathbf{y}|\mathbf{f})$ and $P(\mathbf{x}|\mathbf{f}) = P(\mathbf{x})$, which is assumed to have a uniform distribution. According to (1), (2) and (12),

$$\begin{aligned} P(\mathbf{f}|S) &\propto \exp\left(-\frac{(\mathbf{y} - \mathbf{f})^T(\mathbf{y} - \mathbf{f})}{2\sigma^2} - \frac{\mathbf{f}^T K^{-1} \mathbf{f}}{2}\right) \\ &\propto \exp\left(-\frac{\mathbf{f}^T (K^{-1} + \sigma^{-2}I) \mathbf{f} - 2\sigma^{-2} \mathbf{y}^T \mathbf{f}}{2}\right) \\ &\propto \exp\left(-\frac{(\mathbf{f} - \mathbf{u}) \Sigma^{-1} (\mathbf{f} - \mathbf{u})}{2}\right), \end{aligned}$$

where

$$\begin{aligned} \Sigma &= (K^{-1} + \sigma^{-2}I)^{-1} \\ &= (K^{-1} + \sigma^{-2}K K^{-1})^{-1} \\ &= \left((I + \sigma^{-2}K)K^{-1}\right)^{-1} \\ &= \sigma^2 K (K + \sigma^2 I)^{-1} \end{aligned}$$

and

$$\mathbf{u} = \sigma^{-2} \Sigma \mathbf{y} = K (K + \sigma^2 I)^{-1} \mathbf{y}.$$

Therefore,

$$(\mathbf{f}|S) \sim N\left(K(K + \sigma^2 I)^{-1} \mathbf{y}, \sigma^2 K (K + \sigma^2 I)^{-1}\right). \tag{13}$$

2.3 Derivation of $P(f^*|\mathbf{x}^*, S)$

For the ease of presentation, we define

$$\begin{aligned} \mathbf{a} &\equiv K^{-1} \mathbf{k}, \\ \Delta &\equiv \mathcal{K}(\mathbf{x}^*, \mathbf{x}^*) - \mathbf{k}^T K^{-1} \mathbf{k}, \\ \mathbf{b} &\equiv K (K + \sigma^2 I)^{-1} \mathbf{y}, \\ \Sigma &\equiv \sigma^2 K (K + \sigma^2 I)^{-1}. \end{aligned}$$

Then according to (3), (10) and (13),

$$\begin{aligned}
& P(f^*|\mathbf{x}^*, S) \\
& \propto \int \exp\left(-\frac{(f^* - \mathbf{a}^T \mathbf{f})^2}{2\Delta} - \frac{(\mathbf{f} - \mathbf{b})^T \Sigma^{-1} (\mathbf{f} - \mathbf{b})}{2}\right) d\mathbf{f} \\
& = \int \exp\left(-\frac{(f^*)^2}{2\Delta} - \frac{1}{2}\left((\mathbf{f} - \mathbf{b})^T \Sigma^{-1} (\mathbf{f} - \mathbf{b}) - \frac{2f^* \mathbf{a}^T \mathbf{f}}{\Delta} + \mathbf{f}^T \frac{\mathbf{a} \mathbf{a}^T}{\Delta} \mathbf{f}\right)\right) d\mathbf{f} \\
& \propto \int \exp\left(-\frac{(f^*)^2}{2\Delta} - \frac{1}{2}\left(\mathbf{f}^T (\Sigma^{-1} + \frac{\mathbf{a} \mathbf{a}^T}{\Delta}) \mathbf{f} - 2(\Sigma^{-1} \mathbf{b} + \frac{f^* \mathbf{a}}{\Delta})^T \mathbf{f}\right)\right) d\mathbf{f} \\
& \propto \exp\left(-\frac{(f^*)^2}{2\Delta} + \frac{1}{2}\left(\frac{f^* \mathbf{a}}{\Delta} + \Sigma^{-1} \mathbf{b}\right)^T \left(\Sigma^{-1} + \frac{\mathbf{a} \mathbf{a}^T}{\Delta}\right)^{-1} \left(\frac{f^* \mathbf{a}}{\Delta} + \Sigma^{-1} \mathbf{b}\right)\right) \\
& \propto \exp\left(-\frac{1}{2}\left(\frac{1 - \mathbf{a}^T (\Delta \Sigma^{-1} + \mathbf{a} \mathbf{a}^T)^{-1} \mathbf{a}}{\Delta} (f^*)^2 - 2\mathbf{a}^T (\Delta \Sigma^{-1} + \mathbf{a} \mathbf{a}^T)^{-1} \Sigma^{-1} \mathbf{b} f^*\right)\right).
\end{aligned}$$

Therefore,

$$(f^*|\mathbf{x}^*, S) \sim N(\mu^*, (\sigma^*)^2)$$

where

$$\begin{aligned}
\mu^* &= \frac{\Delta \mathbf{a}^T (\Delta \Sigma^{-1} + \mathbf{a} \mathbf{a}^T)^{-1} \Sigma^{-1} \mathbf{b} f^*}{1 - \mathbf{a}^T (\Delta \Sigma^{-1} + \mathbf{a} \mathbf{a}^T)^{-1} \mathbf{a}}, \\
(\sigma^*)^2 &= \frac{\Delta}{1 - \mathbf{a}^T (\Delta \Sigma^{-1} + \mathbf{a} \mathbf{a}^T)^{-1} \mathbf{a}}.
\end{aligned}$$

Using the *Sherman-Morrison-Woodbury formula*:

$$(A + UV^T)^{-1} = A^{-1} - A^{-1}U(I + V^T A^{-1}U)^{-1}V^T A^{-1}, \quad (14)$$

we have

$$\begin{aligned}
(\Delta \Sigma^{-1} + \mathbf{a} \mathbf{a}^T)^{-1} &= \frac{\Sigma}{\Delta} - \frac{\Sigma}{\Delta} \mathbf{a} \left(1 + \frac{\mathbf{a}^T \Sigma \mathbf{a}}{\Delta}\right)^{-1} \mathbf{a}^T \frac{\Sigma}{\Delta} \\
&= \frac{1}{\Delta} \left(\Sigma - \frac{\Sigma \mathbf{a} \mathbf{a}^T \Sigma}{\Delta + \mathbf{a}^T \Sigma \mathbf{a}}\right).
\end{aligned}$$

Consequently,

$$\begin{aligned}
\mu^* &= \frac{\mathbf{a}^T \left(\Sigma - \frac{\Sigma \mathbf{a} \mathbf{a}^T \Sigma}{\Delta + \mathbf{a}^T \Sigma \mathbf{a}}\right) \Sigma^{-1} \mathbf{b} f^*}{1 - \frac{1}{\Delta} \mathbf{a}^T \left(\Sigma - \frac{\Sigma \mathbf{a} \mathbf{a}^T \Sigma}{\Delta + \mathbf{a}^T \Sigma \mathbf{a}}\right) \mathbf{a}} = \frac{1 - \frac{\mathbf{a}^T \Sigma \mathbf{a}}{\Delta + \mathbf{a}^T \Sigma \mathbf{a}}}{1 - \frac{\mathbf{a}^T \Sigma \mathbf{a}}{\Delta} + \frac{(\mathbf{a}^T \Sigma \mathbf{a})^2}{\Delta(\Delta + \mathbf{a}^T \Sigma \mathbf{a})}} \mathbf{a}^T \mathbf{b} f^* \\
&= \mathbf{a}^T \mathbf{b} f^* = \mathbf{k}^T (K + \sigma^2 I)^{-1} \mathbf{y},
\end{aligned}$$

and

$$\begin{aligned}
(\sigma^*)^2 &= \frac{\Delta}{1 - \frac{\mathbf{a}^T \Sigma \mathbf{a}}{\Delta} + \frac{(\mathbf{a}^T \Sigma \mathbf{a})^2}{\Delta(\Delta + \mathbf{a}^T \Sigma \mathbf{a})}} \\
&= \Delta + \mathbf{a}^T \Sigma \mathbf{a} \\
&= \Delta + \sigma^2 \mathbf{k}^T (\sigma^2 K + K K)^{-1} \mathbf{k} \tag{15}
\end{aligned}$$

$$\begin{aligned}
&= \Delta + \sigma^2 \mathbf{k}^T \left(\sigma^{-2} K^{-1} - \sigma^{-4} (K \sigma^{-2} + I)^{-1} \right) \mathbf{k} \tag{16} \\
&= \Delta + \mathbf{k}^T K^{-1} \mathbf{k} - \mathbf{k}^T (K + \sigma^2 I)^{-1} \mathbf{k} \\
&= \mathcal{K}(\mathbf{x}^*, \mathbf{x}^*) - \mathbf{k}^T (K + \sigma^2 I)^{-1} \mathbf{k}.
\end{aligned}$$

From (15) to (16), we use (14) with $A = \sigma^2 K$ and $U = V = K$. Finally, we are able to give the predictive distribution:

$$(f^* | \mathbf{x}^*, S) \sim N\left(\mathbf{k}^T (K + \sigma^2 I)^{-1} \mathbf{y}, \mathcal{K}(\mathbf{x}^*, \mathbf{x}^*) - \mathbf{k}^T (K + \sigma^2 I)^{-1} \mathbf{k}\right).$$